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Short Note

Zygoty Test Based on Growth Measures: A Bayesian Approach

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A statistical index is proposed in order to determine twin zygoty based on weight and height measures and a Bayesian test of hypotheses. Some Roman twin data are analyzed on the basis of the Euclidean distance index, which appears to be the best for an easy calculation and for its genetic interpretation based on the formulation of genetic distance.

Key words: Twin zygoty, Twin concordance, Bayesian approach, Genetic distance, Height, Weight

INTRODUCTION

Longitudinal studies of growth measures in twins indicate different patterns in monozygoty (MZ) versus dizygoty (DZ) twins (eg, see [3]). We should therefore like to propose a mathematical model of twin concordance allowing for a probabilistic zygoty determination based on twin longitudinal data. What we shall propose is a suitable distance function between cotwins through which, based on growth measures such as height and weight, MZ and DZ pairs can be discriminated with a high level of probability. We shall then consider growth measures of pairs of unknown zygoty in order to obtain a posterior probability of monozygoty versus dizygoty as a function of the distance index.

MATHEMATICAL MODEL

Many statistical indices of diversity are available as suitable distance functions between individuals of a twin pair based on growth measures, and many papers have been written about the mathematical properties of such functions (eg, [1], and for a review [2]). In the present paper we take into account only the so-called Eu-

clidean normalized distance, which allows good results for our problem in a very easy way. Let us define this function.

Consider two vectors $x \Leftrightarrow (x_1, x_2, \dots, x_n)$ and $y \Leftrightarrow (y_1, y_2, \dots, y_n)$ define:

$$d(x, y) = \left\{ \sum_{i=1}^n \left(\frac{x_i - y_i}{\frac{(x_i + y_i)}{2}} \right)^2 \right\}^{1/2}$$

The function $d(x, y)$ is the normalized Euclidean distance of the vectors x and y .

Let us now consider the following mathematical model of concordance for biometric traits of MZ and DZ pairs.

Ronald Fisher (1918) proposed that the genotypic expression of “continuous” hereditary traits is a normal random variable X_g with mean m_g and variance σ_g^2 , depending on sex and age. A random independent overlapping fluctuation X_e , due to environmental conditions, must be considered and results in a normal random variable with mean m_e and variance σ_e^2 . If we define $X(n)$ the random variable that represents height measure at age n , and $Y(n)$ the random variable that represents weight measure at age n , for any individual of a well defined population, we can write:

$$\begin{aligned} X(n) &= X_g(n) + X_e(n) \\ Y(n) &= Y_g(n) + Y_e(n) \end{aligned}$$

where

$$\begin{aligned} X(n) &\sim N (E(X_g(n)) + E(X_e(n)), \sigma^2(X_g(n)) + \sigma^2(X_e(n))) \\ Y(n) &\sim N (E(Y_g(n)) + E(Y_e(n)), \sigma^2(Y_g(n)) + \sigma^2(Y_e(n))) \end{aligned}$$

Let us now consider a twin pair and define $X^{(i)}(n)$ and $Y^{(i)}(n)$ ($i = 1, 2$) the height and weight measures of the i -th individual of the pair, at age n . We can write

	for MZ pairs	for DZ pairs
$X^{(1)}(n)$	$= X_g^{(1)}(n) + X_e^{(1)}(n)$	$X_g^{(1)}(n) + X_e^{(1)}(n)$
$Y^{(1)}(n)$	$= Y_g^{(1)}(n) + Y_e^{(1)}(n)$	$Y_g^{(1)}(n) + Y_e^{(1)}(n)$
$X^{(2)}(n)$	$= X_g^{(1)}(n) + X_e^{(2)}(n)$	$X_g^{(2)}(n) + X_e^{(2)}(n)$
$Y^{(2)}(n)$	$= Y_g^{(1)}(n) + Y_e^{(2)}(n)$	$Y_g^{(2)}(n) + Y_e^{(2)}(n)$

It follows immediately that, for all n :

$$E(d_n(DZ)) > E(d_n(MZ))$$

where

$$d_n(DZ) = \left\{ \left(\frac{X^{(1)}(n) - X^{(2)}(n)}{\frac{X^{(1)}(n) + X^{(2)}(n)}{2}} \right)^2 + \left(\frac{Y^{(1)}(n) - Y^{(2)}(n)}{\frac{Y^{(1)}(n) + Y^{(2)}(n)}{2}} \right)^2 \right\}^{1/2}$$

with the same expression for $d_n(\text{MZ})$.

Statistical Distribution of $d_n(\text{DZ})$ and $d_n(\text{MZ})$

Based on growth measures of 28 MZ and 28 DZ twin pairs obtained at $n = 0, 3, 6, 9,$ and 12 months of age, $d_n(\text{MZ})$ and $d_n(\text{DZ})$ values have been derived (Table 1). Statistical mean and variance have then been calculated so that:

$$D_n(\text{DZ}) = \frac{d_n(\text{DZ}) - E(d_n(\text{DZ}))}{\sigma(d_n(\text{DZ}))}$$

and

$$D_n(\text{MZ}) = \frac{d_n(\text{MZ}) - E(d_n(\text{MZ}))}{\sigma(d_n(\text{MZ}))}$$

have a probability density function t_{27} and then, for instance (using the table of Student's t),

$$\begin{aligned} P(d_n(\text{DZ}) \in (E(d_n(\text{DZ})) \pm 2.0526(d_n(\text{DZ}))) &= 0.99 = \\ &= P(d_n(\text{MZ}) \in (E(d_n(\text{MZ})) \pm 2.0526(d_n(\text{MZ}))). \end{aligned}$$

We could now apply standard hypothesis testing theory to any twin pair and get a probabilistic determination of zygosity based on the index d_n and an approximation of the likelihood functions given by the normal densities with means given by $\bar{d}_n(\text{DZ})$ and $\bar{d}_n(\text{MZ})$.

A Test From a Bayesian Point of View

Let us now consider a Bayesian statistical test for zygosity determination.

If we call $P(\text{MZ})$ the prior probability that any twin pair be MZ and $P(\text{DZ}) = 1 - P(\text{MZ})$ the prior probability that any twin pair be DZ¹, we can write:

$$P(\text{MZ}/d_n) = \frac{f(d_n/\text{MZ})P(\text{MZ})}{f(d_n/\text{MZ})P(\text{MZ}) + g(d_n/\text{DZ})P(\text{DZ})}$$

and

$$P(\text{DZ}/d_n) = \frac{g(d_n/\text{DZ})P(\text{DZ})}{f(d_n/\text{MZ})P(\text{MZ}) + g(d_n/\text{DZ})P(\text{DZ})}$$

where $f(d_n/\text{MZ})$ is the probability density function of the random variable $d_n(\text{MZ})$ and $g(d_n/\text{DZ})$ the probability density function of $d_n(\text{DZ})$, which have been approximated by normal densities with parameters estimated on the basis of our statistical data.

¹That can be statistically estimated. Based on the twin data available at the Mendel Institute, we obtained the following estimations: $P(\text{MZ}) = 0.29, P(\text{DZ}) = 0.71$.

TABLE 1. Statistical Distribution of $d_n(DZ)$ and $d_n(MZ)$ at $n = 0, 3, 6, 9, \text{ and } 12$ Months of Age

$d_n(MZ)$	$d_n(DZ)$	$d_3(MZ)$	$d_3(DZ)$	$d_6(MZ)$	$d_6(DZ)$	$d_9(MZ)$	$d_9(DZ)$	$d_{12}(MZ)$	$d_{12}(DZ)$
.280	.223	.186	.216	.167	.211	.194	.233	.120	.166
.077	.083	.129	.112	.054	.093	.106	.293	.050	.059
.044	.074	.035	.033	.046	.051	.013	.067	.041	.041
.049	.223	.014	.215	.112	.245	.046	.077	.013	.267
.255	.068	.146	.090	.050	.019	.104	.071	.017	.048
.250	.044	.103	.075	.024	.042	.034	.109	.023	.190
.043	.148	.034	.092	.017	.120	.032	.164	.022	.083
.074	.199	.000	.130	.030	.111	.274	.102	.013	.209
.087	.139	.018	.175	.030	.123	.083	.061	.024	.134
.148	.224	.049	.172	.091	.084	.074	.143	.083	.031
.174	.314	.139	.116	.396	.203	.021	.117	.087	.159
.174	.120	.420	.079	.078	.052	.012	.018	.044	.027
.109	.220	.075	.119	.008	.072	.062	.055	.102	.136
.068	.513	.041	.053	.142	.045	.022	.089	.024	.112
.024	.239	.018	.107	.015	.094	.013	.108	.034	.138
.043	.183	.091	.041	.046	.048	.105	.124	.146	.075
.122	.076	.084	.085	.109	.045	.179	.045	.171	.000
.188	.304	.077	.251	.029	.108	.042	.113	.041	.108
.056	.037	.181	.063	.051	.239	.054	.179	.040	.161
.326	.350	.054	.102	.158	.025	.021	.120	.022	.044
.164	.115	.138	.087	.033	.120	.012	.116	.072	.140
.293	.176	.044	.055	.038	.119	.044	.082	.019	.217
.051	.143	.038	.000	.050	.019	.035	.031	.052	.101
.067	.039	.039	.116	.016	.071	.014	.087	.025	.075
.066	.000	.010	.136	.017	.076	.038	.046	.029	.065
.137	.128	.062	.000	.029	.063	.036	.100	.052	.132
.073	.196	.078	.155	.084	.193	.027	.041	.048	.037
.313	.000	.020	.070	.922	.068	.019	.024	.030	.056
$E(d_0(MZ)) = 0.134$	$E(d_3(MZ)) = 0.083$	$E(d_6(MZ)) = 0.069$	$E(d_9(MZ)) = 0.061$	$E(d_{12}(MZ)) = 0.052$					
$\sigma^2(d_0(MZ)) = 0.008$	$\sigma^2(d_3(MZ)) = 0.007$	$\sigma^2(d_6(MZ)) = 0.006$	$\sigma^2(d_9(MZ)) = 0.004$	$\sigma^2(d_{12}(MZ)) = 0.002$					
$E(d_0(DZ)) = 0.164$	$E(d_3(DZ)) = 0.105$	$E(d_6(DZ)) = 0.099$	$E(d_9(DZ)) = 0.101$	$E(d_{12}(DZ)) = 0.108$					
$\sigma^2(d_0(DZ)) = 0.013$	$\sigma^2(d_3(DZ)) = 0.004$	$\sigma^2(d_6(DZ)) = 0.004$	$\sigma^2(d_9(DZ)) = 0.004$	$\sigma^2(d_{12}(DZ)) = 0.004$					

TABLE 2. Posterior Probabilities of Zygoty Based on Growth Measures at 0, 3, 6, 9, and 12 Months of Age

Function	X	P(MZ/X)	P(DZ/X)
d ₀	.05880	.36062	.63938
	.07060	.36341	.63659
	.09710	.36405	.63595
	.10210	.36329	.63671
	.10550	.36262	.63738
	.13330	.35237	.64763
	.14770	.34380	.65620
	.17060	.32587	.67413
	.22590	.26399	.73601
	.28830	.17673	.82327
d ₃	.05880	.27884	.72116
	.07060	.26149	.73851
	.10210	.23144	.76856
	.10550	.22947	.77053
	.13330	.22170	.77830
	.14770	.22334	.77666
	.17060	.23408	.76592
	.22590	.30860	.69140
	.28830	.50352	.49648
	d ₆	.05880	.28807
.07060		.26943	.73057
.10210		.23358	.76642
.10550		.23079	.76921
.13330		.21491	.78509
.14770		.21119	.78881
.17060		.21123	.78877
.22590		.24294	.75706
.28830		.34830	.65170
d ₉		.05880	.33774
	.07060	.31187	.68813
	.10210	.24855	.75145
	.10550	.24225	.75775
	.13330	.19492	.80508
	.14770	.17330	.82670
	.17060	.14290	.85710
	.22590	.08751	.91249
	.28830	.04887	.95113
	d ₁₂	.05880	.43591
.07060		.38688	.61312
.10210		.23650	.76350
.10550		.22036	.77964
.13330		.10705	.89295
.14770		.06652	.93348
.17060		.02724	.97276
.22590		.00171	.99829
.28830		.00003	.99997

The statistical test is based on such posterior distributions rather than only on likelihood functions. Posterior probabilities for twin pairs of unknown zygosity, based on functions $d_0, d_3, d_6, d_9, d_{12}$, are given in Table 2.

Much more could be said on the basis of the following comprehensive index:

$$d_{0N} = \left\{ \sum_{i=0}^N d_{fi}^2 \right\}^{1/2}$$

which takes into account more information from growth measures.

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