Coexisting forms of malnutrition among under-5 children in Bangladesh: results from 2012-13 and 2019 Multiple Indicator Cluster Surveys

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Short title: Coexisting forms of malnutrition in U-5 children



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Indicator Cluster Surveys (MICS), where all personally identifiable information has been

removed to ensure anonymity. As the study used secondary data, no ethical clearance is

required. However, prior to data collection, ethical clearance was obtained from the relevant

authority by the implementing body. Written informed consents were also taken prior to the

commencement of the interview from the participants.

Abstract

Objective: Underweight, wasting, and stunting are crucial malnutrition indicators responsible

for morbidities among children. Data regarding coexisting forms of malnutrition (CFM) is

scarce. We aimed to investigate the prevalence and associated factors of CFM across two

survey years among under-5 Bangladeshi children.

Design: Cross-sectional study.

Setting: Data were acquired from two rounds of Multiple Indicator Cluster Survey (MICS),

Bangladesh conducted in 2012-13 and 2019.

Subjects: The analysis included 43,946 (2012-13: 20,885; 2019: 23,061) under-5 children.

Results: Binomial proportion test, slope index of inequality, and multinomial logistic

regression models were used for analysis. The prevalence of CFM was 27.45% and 18.56%

in 2012-13 and 2019, respectively. A significant decrease in the prevalence of CFM was seen

across the surveys (p-value<0.001). Children from urban residence [(2012-13:aOR=0.70,

95% CI:0.64,0.77); (2019:aOR=0.71, 95% CI:0.65,0.78], higher maternal education [(2012-

13:aOR=0.28, 95% CI:0.24,0.32); (2019:aOR=0.28, 95% CI:0.24,0.32], larger size at birth

[(2012-13:aOR=0.62, 95% CI:0.52,0.73); (2019:aOR: 0.60, 95% CI:0.50,0.73], richest

wealth quintile [(2012-13:aOR=0.25, 95% CI:0.22,0.28); (2019:aOR: 0.30, 95%

CI:0.27,0.34)] had lower odds of suffering from CFM compared to their counterparts.

Children from poorer quintiles were more influenced by CFM than richer quintiles (Coef.:-

0.175, 95% CI:-0.192,-0.157, p-value<0.001). Higher percentage of CFM was observed

among rich families in 2019 compared to 2012-13 (24.50% and 20.15%, respectively; p-

value<0.001)

Conclusion: The findings of this study should help the researchers and policymakers to

understand CFM more clearly and plan prospective studies to explore CFM outcomes.

Targeted interventional approaches are needed among parents of rural communities to control

the burden of CFM.

Keywords: Pediatric malnutrition, MICS, Under-5 children, multinomial logistic regression

Introduction

Malnutrition is a prevailing and alarming condition that affects millions of under-5 children all over the world, mostly in low-and middle-income countries (LMICs) ⁽¹⁾. Every one in five children is at risk of suffering from malnutrition in developing countries like Bangladesh ⁽¹⁾. Childhood malnutrition, including undernutrition, is attributable to compromised growth, metabolic and immune dysfunction, higher risk of infection, altered development of brain and cognitive function, and various abnormalities ^(1, 2). According to the World Health Organization (WHO) child growth standards, wasting, stunting, and underweight are three of the most critical indicators of malnutrition that can affect a child's health outcomes and general wellbeing ⁽³⁾. While wasting and underweight condition is considered to be the result of acute malnutrition, stunting represents chronic malnutrition, mainly causing irreversible damage to the child ^(4, 5). A wasted child becomes too thin for one's height measured by weight-for-height z score, whereas underweight includes children whose weight is low compared to their age ⁽³⁾. Stunting refers to the condition where a child's height is less in contrast to their age and is measured by height-for-age z score ⁽³⁾.

According to the WHO factsheet, among the five million under-5 children deaths in the world, about 45% or 2.4 million deaths can be linked to malnutrition, most of which occurred in LMICs ⁽⁶⁾. The recent "Levels and trends in child malnutrition" report in 2023 by WHO, UNICEF, and World Bank states that around 148 million children are still stunted, whereas 45 million under-5 children are suffering from wasting ⁽⁷⁾. A staggering number of these children are from Southern Asia and mostly are from resource-poor countries ⁽⁷⁾. The Bangladesh Demographic and Health Survey (BDHS) 2022 reports that the prevalence of stunting is currently 24%, which declined from 31% (BDHS 2017-18), but the proportion still contains a large number of children. Relating to this fact, the prevalence of wasting increased to 11% from 8%, and the percentage of underweight remained constant to 22% ⁽⁸⁾.

The presence of multiple nutritional dysfunctions is referred to as coexisting forms of malnutrition (CFM) ⁽⁹⁾. Under-5 children with CFM bear a higher risk of mortality compared to standalone forms of malnutrition ⁽¹⁰⁾. It also becomes challenging and difficult to control when multiple forms of malnutrition coexist in the body, resulting in increased health risks compared to a single form of malnutrition ⁽¹¹⁾. *McDonald et al.* published an article in 2013 where they showed that more than half of the malnourished children are suffering from CFM worldwide, and even one unit increase in their anthropometric deficits can increase the risk of death among under-5 children ⁽¹²⁾. Although children who are suffering from a single form of

malnutrition possess a two times higher risk of death compared to non-malnourished children, the risk expands tenfold in children suffering from coexisting forms of undernutrition ⁽¹²⁾. In Bangladesh, children who are suffering from wasting, stunting, and underweight conditions combined have nineteen times higher odds of mortality compared to healthy children ⁽¹²⁾.

It is evident from previous studies conducted on demographic and health surveys that socio-demographic and socio-economic factors are crucial determinants of child malnutrition (13-16). The prevalence of undernutrition can be attributed to multiple factors such as the mother's age and height, parent's education, residence, wealth index, and other characteristics (17, 18). Although an LMIC like Bangladesh is burdened with a major public health concern like childhood malnutrition, results depicting associated factors of CFM among under-5 children are limited. The primary objective of our analysis was to investigate the prevalence across survey years, and important factors that are associated with coexisting forms of malnutrition among under-5 children, where the CFM was defined as the presence of any two or all three nutritional deficiencies (wasting, stunting, and underweight) in under-5 children.

Methods

The data used in the current study were obtained from two separate rounds of the Multiple Indicator Cluster Survey (MICS) conducted in Bangladesh. The two survey rounds were 2012-13 (Survey 01) and 2019 (Survey 02). MICS is an important tool for gathering standardized, scientifically sound, and nationally representative information. These surveys provide valuable insights on socio-economic standings and various health-related indicators. The survey dataset and findings have been made publicly available on UNICEF's MICS website at (https://mics.unicef.org/surveys). The current study accessed the children and women datasets from each of Bangladesh's MICS rounds of 2012-13 and 2019. Both of the surveys used a probability-based sampling method and a two-stage sample selection procedure. Census enumeration areas were identified systematically using the probability proportional to size method. This was followed by the selection of 20 households from each enumeration area (55,120 and 64,400 households for 2012-13 and 2019, respectively). More comprehensive information on the MICS methodology can be found elsewhere (19,20).

Children aged <60 months were included in the present study from the publicly accessible datasets and merged with the women 'wm' dataset following MICS guidelines for data merging. Cases with "missing values" for age and biologically implausible or inappropriate

anthropometric values were excluded (<-5 standard deviation (SD) or > 5 SD for WHZ, <-5 SD or > 5 SD for HAZ, <-6 SD or > 5 SD for WAZ) (21). Finally, 20,885 children from 2012-13 and 23,061 from 2019 survey rounds were included in the study. The procedure used to select children for this investigation is described in detail in Figure 01.

Operational definition of the dependent variables

According to the WHO, malnutrition among children can be defined as the excess, deficiency, or imbalance in the intake of nutrients and or energy ⁽¹⁾. We defined the nutritional indicators as per the WHO child growth standards (2006) in this analysis ⁽³⁾.

Stunting: Stunting was defined as height/length-for-age z score <-2 standard deviations (SD)

Wasting: Wasting was defined as weight-for-height/length z score <-2 SDs

Underweight: Underweight was defined as weight-for-age z score <-2 SDs

A composite variable "Type of malnutrition" was created by combining stunting, wasting, and underweight, which had three categories:

Absent: If the participant did not suffer from any of the three defined forms of malnutrition Single form of malnutrition (SFM): If the participant was suffering from any one of the three defined forms of malnutrition

Coexisting forms of malnutrition (CFM): If the participant was suffering from more than one of the three defined forms of malnutrition simultaneously

Independent variables

The current study considered the MICS dataset and selected independent variables on an apriori basis. The variables included are the child's age (in months), sex (male/female), residence (urban/rural), ever breastfed (yes/no), childhood illness (yes/no), mothers age in years, maternal education, size at birth (average/larger/smaller), maternal media exposure (not exposed/exposed), wealth index (poorest/poor/middle/rich/richest). For 2012-13 survey, the variable religion (Islam/Hinduism/Buddhism/Christianity/Others) and for 2019 survey, ethnicity (Bengali/Others) and maternal life satisfaction (satisfied/neutral/dissatisfied) were additionally added. A dummy variable (age group in months) was created for trend test which had five categories: "0-12 months", "13-24 months", "25-36 months", "37-48 months" and "49-59 months" for both survey years.

Mother's education

The variable "mother's education" was categorized differently in the two survey years. For survey year 2012-13, the categories were none, primary incomplete, secondary incomplete, secondary complete, or higher. However, for survey year 2019, it was pre-primary or none, primary, secondary, higher secondary, or more.

Wealth index

The study categorized participants into different socio-economic groups based on a wealth index, which is used to place participants into one of five wealth categories: poorest, poor, middle, rich, and richest, by taking inventory of observable household items and amenities ^(22, 23). This method provides an opportunity to gauge relative economic standing without directly measuring income or expenditures ^(24, 25). A dummy variable was also created having two categories, "poor-to-middle" (combining poorest, poor, and middle quintiles) and "rich" (combining richest and rich quintiles).

Maternal life satisfaction

Maternal life satisfaction was measured using a single item: "First, taking all things together, would you say you are very happy, somewhat happy, neither happy nor unhappy, somewhat unhappy or very unhappy?" Out of five possible responses, "very happy" and "somewhat happy" was categorized as "satisfied", "neither happy nor unhappy" as "neutral", and "somewhat unhappy or very unhappy" as "dissatisfied".

Size at birth

Child's size at birth was measured using a single item: "When (name of the child) was born, was (he/she) very large, larger than average, average, smaller than average, or very small?" The current study categorized "very large" and "larger than average" as "large", "average" as "average", and "smaller than average" and "very small" as "small".

Maternal media exposure

"Maternal media exposure" was created using the summation of three variables: "frequency of reading newspaper or magazine", "frequency of listening to the radio", "frequency of watching TV". This variable has two categories as per previous literature: Not exposed (0 score) and exposed (>0 score) (26).

Statistical analysis

In the current study, descriptive statistics were presented as frequency and percentage for categorical variables and median and IQR for continuous variables. Kolmogorov-Smirnov test was conducted on continuous variables, and all were found to be significantly deviated

(p-value <0.05) from normal distribution. The χ^2 goodness-of-fit test was used to assess whether the wealth index of participants with CFM across two survey years were significantly different. The Binomial proportion test was applied to determine the statistical significance of the difference between the prevalence of CFM across two survey years. Kruskal-Wallis test, χ^2 test of association, and Fisher's exact test were used to evaluate the association between independent and outcome variables. The Cochran-Armitage test for trend was conducted to evaluate changes in trend of CFM across participant's age. We used bivariate screening method to select variables for the multinomial logistic regression models. A p-value of <0.05 was considered as significant in the bi-variate screening method where only participants with non-missing values were considered. Separate multinomial logistic regression models were developed for each of the survey year (2012-13 and 2019). We adjusted for age and sex to calculate the adjusted odds ratio for each predictor variable, as such missing values for one variable did not have any effect on the adjusted odds ratio for other variables. The adjusted odds ratios (after adjusting for age and sex) and 95% confidence intervals (CIs) were reported. The slope index of inequality (SII) was used to compare and quantify the association between the wealth index and CFM across both surveys. It quantifies absolute inequality by measuring the difference in an indicator between the most advantaged and most disadvantaged groups, while accounting for all intermediate subgroups. (27) To calculate the Slope Index of Inequality (SII), the entire population is ranked from the most disadvantaged subgroup (rank 0) to the most advantaged subgroup (rank 1), with rankings weighted by each subgroup's proportional population size. The indicator of interest is then regressed against these midpoint values using an appropriate model. Predicted values of the indicator are calculated for the extremes (rank 0 and rank 1), and the difference between these values gives the SII. An SII value of zero indicates no inequality. For our study, a negative coeficiant indicates that CFM is concentrated among the poorest quintile than richest quintile. (27) A p-value of <0.05 was considered statistically significant ($\alpha = 0.05$). All descriptive and inferential analysis was done using STATA 17.

Results

We have included 43,946 children from the two MICS surveys where 20,885 were selected from the year 2012-13 and 23,061 belonged to the year 2019. In 2012-13, 52.35% (n=10,934) children did not have any form of malnutrition, 20.20% (n=4,219) children had SFM, and 27.45% (n=5,732) children had CFM. For 2019, this proportion was 63.57% (n=14,661),

17.86% (n=4,119), and 18.57% (n=4,281), respectively. Table 1 describes the socio-demographic characteristics among under-5 children and their associations with types of malnutrition across two survey years.

Socio-demographic characteristics

During survey year 2012-13, the median (IQR) age was higher among CFM group [34 (20,46) months] compared to the non-malnourished group where the association was statistically significant [28 (12,45); p-value<0.001]. Several factors were found to be significantly associated with different types of malnutrition, including residence, maternal age, maternal education, size at birth, maternal media exposure, and wealth index. Most of the children (84.07%) were from rural areas, which was also the case among each type of malnutrition. Additionally, around 60% of mothers in all groups fell into the age categories of 20-24 or 25-29, with over 33.79% not completing secondary education. In the CFM group specifically, 31.47% of mothers had no formal education, and 50.46% was exposed to media. Regarding the wealth index, 29.97% belonged to the poorest group, which was true for each of the malnutrition groups.

In survey year 2019, several factors were found to be significantly associated with malnutrition types, including the age of child, residence, size at birth, breastfeeding history, childhood illness, maternal age, maternal education, ethnicity, caregiver of other children, maternal life satisfaction, maternal media exposure, and wealth index. Among children with CFM, the median (IQR) age was 32 (18,45) months, with 27.99% being smaller in size at birth. Furthermore, 85.07% of these children lived in rural areas, and 33.71% belonged to the poorest wealth index category. Overall, 95.62% of children were breastfed at some point before the survey, while among CFM cases, it was 94.64%. More than half of the mothers in each type of malnourished children belonged to the 20-24 years or 25-29 years age group.

Variations in CFM prevalence across surveys

Figure 2 shows a decrease in the prevalence of CFM from 27.45% (95% CI: 26.86%-28.05%; n=5,732) in the 2012-13 survey to 18.56% (95% CI: 18.07%-19.07%; n=4,281) in 2019, which was statistically significant (p-value < 0.001).

Figure 3 shows a statistically significant rise in trend of CFM from 3.5% among participants aged 0-12 months to 5.8% among 49-59 months in survey year 2012-13. However, the maximum proportion was 6.5% in age group 37-48 months followed by 6.2% in 25-36 months. A similar rising trend was also observed in survey year 2019 where proportion of

CFM raised from 2.7% to 3.7% with a maximum of 4.2% in age groups 25-36 months and 37-48 months.

Figure 4 represents the distribution of wealth index among participants who were suffering from CFM across both the survey years. Higher percentage of CFM was observed among rich families in 2019 compared to 2012-13 (24.50% and 20.15%, respectively) which was significantly different (p-value<0.001).

Slope index of inequality (SII) analysis

The slope index of inequality (SII) presented in figure 05. The negative values of the coef. for survey years 2012-13 and 2019 (Coef.: -0.239, 95% CI: -0.259 to -0.218; Coef.: -0.175, 95% CI: -0.192 to -0.157 respectively) indicate that, belonging to richer quintile families was significantly associated with lesser likelihood of under-5 children suffering from CFM.

Multinomial logistic regression analysis

Multinomial logistic regression models were constructed after adjusting for age and sex of the child across both the survey years in Table 02. In survey year 2012-13, the model showed that place of residence, maternal age in years, maternal education, size at birth, maternal media exposure, and wealth index had significant effect on the type of malnutrition. The children living in urban areas had lower odds of CFM (aOR=0.70, 95% CI: 0.64, 0.77, pvalue<0.001) and SFM (aOR=0.78, 95% CI: 0.71, 0.86, p-value<0.001) when compared to children without malnutrition that are living in rural areas. Mothers age group (in years) up to 35-39 were found to have lower odds of CFM, which was statistically significant. Considering the size of the child at birth, if the child was of larger size, the odds for CFM (aOR=0.62, 95% CI: 0.52, 0.73, p-value<0.001) and SFM (aOR=0.73, 95% CI: 0.62, 0.85, pvalue<0.001) decreased, but if it was smaller, the odds of CFM (aOR=2.32, 95% CI: 2.06, 2.61, p-value<0.001) and SFM (aOR=1.42, 95% CI:1.24, 1.62, p-value<0.001) increased. Moreover, when mothers completed secondary level education, the odds of CFM decreased (aOR=0.28, 95% CI: 0.24, 0.32, p-value<0.001) along with SFM (aOR=0.45, 95% CI: 0.40, 0.52, p-value<0.001). Also, when considering maternal media exposure, if the mother was exposed to media, the odds of CFM decreased (aOR=0.61, 95% CI: 0.56, 0.66, pvalue<0.001) compared to those with no media exposure. The results from Table 2 also showed that when the wealth index increased to be on the richer side, the odds of CFM and SFM decreased compared to those with no malnutrition and the poorest group, which was

statistically significant. In Table 2, when considering survey year 2019, we observed that the findings of predictor variables such as residence, mother's age, maternal education, child's size at birth, maternal media exposure, wealth index were similar to the survey year 2012-13. Also, if the caregiver had another child that increased the odds of both malnutrition categories (CFM: aOR=1.25, 95% CI: 1.15-1.35, p-value<0.001; SFM: aOR=1.17, 95% CI: 1.08-1.27, p-value < 0.001), while childhood illness increased the odds of CFM (aOR=1.14, 95% CI: 1.06-1.23, p-value < 0.001). However, if the mother is satisfied with her life, then it decreases the odds of both (CFM: aOR=0.63, 95% CI: 0.57-0.7, p-value<0.001; SFM: aOR=0.79, 95% CI: 0.71-0.89, p-value<0.001).

Discussion

This study presents the prevalence of CFM in Bangladesh, its prevalence across both survey years (2012-13, 2019), and related factors affecting CFM. In both the survey years, we could see that the child's residence, mother's age, maternal education, size at the time of birth, mother's exposure to media, and wealth index affected CFM similarly. Although CFM presented a decreasing prevalence from year 2012-13 to 2019, the overall prevalence of this nutritional disorder is still high in the country. Further analysis revealed that children from poorer families suffered from CFM more compared to the richer families in both survey years.

Bangladesh, a South Asian country, is burdened with a high prevalence of wasting, underweight, and stunting among under-5 children ⁽⁸⁾. For example, in case of severe wasting, more than 0.3 million under-5 children reside in the country, which is one of the major positions among other Asian countries ⁽²⁸⁾. So, it is no surprise that the prevalence of CFM is also high in the country. In our study, we found that CFM exhibits a downward prevalence across the survey years. This is mostly due to the fact that the Bangladesh government has taken the burden of malnutrition among children seriously. The National Nutrition Service (NNS), Institute of Public Health Nutrition (IPHN), non-profit research organizations like icddr,b are working together to achieve the Health, Population and Nutrition Sector Program (HPNSP) target to reduce the prevalence of malnutrition in the country ⁽⁸⁾. The prevalence of wasting, which was precariously high in 2011 at 16%, descended to 11% in 2022 ⁽⁸⁾. Also, the prevalence of stunting and underweight dropped to 24% and 22% in 2022 from the year 2011, where the percentages were 41% and 36%, respectively ⁽⁸⁾. The government has also taken multiple steps countrywide and integrated guideline-based management protocols in the

hospitals to combat malnutrition in children ⁽²⁸⁻³⁰⁾. For example, a recent study found that the physicians in government hospitals had adequate knowledge and a positive attitude while treating malnourished children ⁽²⁸⁾. All of these factors might have influenced the decreasing prevalence of coexisting forms of malnutrition in the country.

Community-level socio-demographic and economic factors play a vital role in the prevalence of malnutrition (31-33). Characteristics like socio-economic condition, parents' educational status, place of residence, childhood illness, and several other factors can affect the nutritional status of a child in developing countries (32-34). A study conducted using multiple BDHS surveys (1996-1997, 1999-2000, 2004, 2007, and 2011) examined the trend of childhood malnutrition and the impact maternal education has on it. The authors reported a consistently significant association where lower maternal education favours childhood malnutrition (35). Another study conducted in Bangladesh also supports this finding (36). A 2023 study conducted in Bangladesh found that parental education and occupation had a profound impact on undernutrition among under-5 children in both urban and rural areas of Bangladesh. The findings indicated that children whose fathers worked low-paying jobs, for example, in the agricultural field, were more likely to experience wasting, stunting, or underweight conditions (36). Another Bangladeshi survey also reported that children from the poorest quintiles had three times higher odds of getting affected by wasting (33). A household's low economic status directly affects its purchasing power, which in turn influences the availability of nutritious food, resulting in food insecurity (37). Households that face social and economic challenges can be associated with decreased physical growth, mostly due to facts like poor food consumption, chronic illnesses, unhygienic lifestyle, insufficient access to pure water, etc. (31). In both our multinomial logistic regression analysis and slope index of inequality, we have observed that CFM was more likely to be associated with children emerging from poorer families compared to children from richer families. These findings can be justified by the above discussion. Additionally, another study conducted among Bangladeshi under-5 children using a nationally representative data reported that CFM is favored by families that belong to lower wealth index quintiles (38). An interesting finding from our analysis showed that over the years CFM is increasing among richer families, though the percentage is still low. A significant rise of CFM was observed in year 2019 compared to the year 2012-13. A recent paper from Bangladesh showed that stunting was present among children whose mothers were engaged in different income

generating activities ⁽³⁹⁾. It can be hypothesized that as both father and mother are participating in earning activities in rich families the children are left at home to unskilled caregivers who has little knowledge about nutrition. The daycare system is also underdeveloped in a LMIC country like Bangladesh. These factors might be affecting the rise of malnutrition among rich families which needs further evaluation.

When it comes to the topic of childhood malnutrition, the urban-rural divergence has remained persistent over the years (40). This pertains to the difference in socio-economic status, poor accessibility to healthcare facilities, medicine, education, and various other factors (33). Multiple studies have reported that rural children are more prone to nutritional deficiency and tend to become malnourished, which corresponds to our findings (36, 40). Our study also indicated that maternal age and education are important predictors of CFM, and if the mother was older and received more education, the odds of her children suffering from CFM decreased. A Bangladeshi study found that children from illiterate fathers or mothers were more likely (approximately 20 to 30%) to remain underweight or stunted compared to parents who received higher education (33). This was coherent with earlier studies conducted in Bangladesh, which found significant impact of parenteral education on both SFM and CFM ^(35, 36, 38). Mothers who are older and educated possess a clear perception and greater understanding regarding the nutrition and health of the children, healthcare-seeking behavior, improved care of the child, washing, hygiene, and sanitation practices. So, it is understandable that these factors are affecting CFM among children. Also, children born with low birth weight are more likely to suffer from different forms of malnutrition, which is consistent with the findings from our study (41). Previous works of literature found that mothers who are exposed to media have better awareness regarding healthcare practices and can significantly decrease malnutrition through household actions (26, 42). This corresponds to our study results as well. Children whose mothers were happy or satisfied with their lives had lower odds of suffering from CFM, according to our study. An article presented that mothers who were unmarried or cohabiting and had low family incomes were more likely to have malnourished children (43). Similarly, poverty-stricken families have lower life satisfaction, which can affect the growth of their children. Another important finding of our study was that children who suffered from illnesses were more vulnerable to CFM. Diseases related to undernutrition can progress to acute or chronic conditions. This mostly occurs due to energy imbalance, high catabolism rate, decreased appetite, and increased energy consumption (34, 44).

So, repeated illness can progress to undernutrition in children. A similar finding was reported by Sumon *et al.*, who found a significant association between CFM and childhood illness ⁽³⁸⁾.

As a nationally representative sampling method was used in both surveys, the current study can be considered to have painted the actual picture of the burden of CFM. Although a few studies have been conducted before focusing on the existence of multiple nutritional disorders the definition of different types of malnutrition in two different time periods is unique to the current study. However, it also had some limitations. Due to the cross-sectional nature of the study design, establishing a causal relationship was not possible. Key variables like food history, maternal malnutrition, trend etc., were not considered due to unavailability, which would have been interesting to explore.

Conclusion

The current study indicates that malnutrition is still a colossal public health concern in an LMIC like Bangladesh. The study findings evaluated the prevalence, trends, and several determinants of coexisting forms of malnutrition among children in the country. This article should capture the attention of researchers, policymakers, and stakeholders regarding the coexisting forms of malnutrition, a topic that is considered minor in the field of nutrition research. The authors propose further prospective studies to investigate the outcomes of CFM in community as well as hospital settings. In light of the findings, the authors recommend enhancing maternal education programs and providing financial assistance to vulnerable families that can improve child nutrition. Scaling up initiatives like the NNS and public awareness campaigns by leveraging media to promote knowledge of nutrition, hygiene, and healthcare-seeking behaviors can reduce the burden of SFM and CFM. Collaboration between multiple sectors can prove to be a beneficial way forward in achieving sustainable reductions in childhood malnutrition.

References

- 1. Black RE, Allen LH, Bhutta ZA *et al.* (2008) Maternal and child undernutrition: global and regional exposures and health consequences. *The lancet* 371, 243-260.
- 2. Black RE, Victora CG, Walker SP *et al.* (2013) Maternal and child undernutrition and overweight in low-income and middle-income countries. *The lancet* 382, 427-451.
- 3. Organization WH (2006) WHO child growth standards: length/height-for-age, weight-for-age, weight-for-length, weight-for-height and body mass index-for-age: methods and development. Geneva, Switzerland: World Health Organization.
- 4. Soliman A, De Sanctis V, Alaaraj N *et al.* (2021) Early and long-term consequences of nutritional stunting: from childhood to adulthood. *Acta Bio Medica: Atenei Parmensis* 92.
- 5. Islam MR, Nuzhat S, Alam J *et al.* (2023) Inpatient Morbidity and Mortality of Severely Underweight Children Was Comparable to That of Severely Wasted Children in Dhaka, Bangladesh. *The American journal of tropical medicine and hygiene* 109, 698-703.
- 6. Organization WH (2024) Child mortality (under 5 years). https://www.who.int/news-room/fact-sheets/detail/levels-and-trends-in-child-under-5-mortality-in-2020 (accessed 15 March 2024 2024)
- 7. Organization WH (2023) Levels and trends in child malnutrition: UNICEF/WHO/World Bank Group Joint Child Malnutrition Estimates, Key findings of the 2023 edition.
- 8. (NIPORT) NIoPRaT (2023) Bangladesh demographic and health survey BDHS 2022: key indicators. Dhaka, Bangladesh, and Rockville, Maryland, USA.
- 9. Khaliq A, Wraith D, Nambiar S *et al.* (2022) A review of the prevalence, trends, and determinants of coexisting forms of malnutrition in neonates, infants, and children. *BMC Public Health* 22, 879.
- 10. Garenne M, Myatt M, Khara T *et al.* (2019) Concurrent wasting and stunting among under-five children in Niakhar, Senegal. *Maternal & child nutrition* 15, e12736.
- 11. Biswas T, Townsend N, Magalhaes RS *et al.* (2019) Current progress and future directions in the double burden of malnutrition among women in South and Southeast Asian countries. *Current developments in nutrition* 3, nzz026.

- 12. McDonald CM, Olofin I, Flaxman S *et al.* (2013) The effect of multiple anthropometric deficits on child mortality: meta-analysis of individual data in 10 prospective studies from developing countries. *The American journal of clinical nutrition* 97, 896-901.
- 13. Abdulla F, El-Raouf MA, Rahman A *et al.* (2023) Prevalence and determinants of wasting among under-5 Egyptian children: Application of quantile regression. *Food Science & Nutrition* 11, 1073-1083.
- 14. Hossain MM, Abdulla F Rahman A (2023) Prevalence and risk factors of underweight among under-5 children in Bangladesh: Evidence from a countrywide cross-sectional study. *PLoS One* 18, e0284797.
- 15. Panda BK, Mohanty SK, Nayak I *et al.* (2020) Malnutrition and poverty in India: does the use of public distribution system matter? *BMC nutrition* 6, 1-14.
- 16. Rahman MA, Halder HR, Rahman MS *et al.* (2021) Poverty and childhood malnutrition: Evidence-based on a nationally representative survey of Bangladesh. *PLoS One* 16, e0256235.
- 17. Félix-Beltrán L, Macinko J Kuhn R (2021) Maternal height and double-burden of malnutrition households in Mexico: stunted children with overweight or obese mothers. *Public health nutrition* 24, 106-116.
- 18. Rahman A & Hossain MM (2022) Quantile regression approach to estimating prevalence and determinants of child malnutrition. *Journal of Public Health*, 1-17.
- 19. Bangladesh Bureau of Statistics UB (2014) *Bangladesh multiple indicator cluster survey 2012–2013, ProgotirPathey: Final reports.* Dhaka, Bangladesh: BBS.
- 20. Bangladesh Bureau of Statistics UB (2019) *ProgotirPathey Bangladesh, multiple indicator cluster survey 2019: survey finding report.* Dhaka, Bangladesh: BBS.
- 21. Corsi DJ, Perkins JM Subramanian S (2017) Child anthropometry data quality from Demographic and Health Surveys, Multiple Indicator Cluster Surveys, and National Nutrition Surveys in the West Central Africa region: are we comparing apples and oranges? *Global health action* 10, 1328185.
- 22. Al Fidah MF & Efa SS (2023) Skin-to-skin contact and early initiation of breast feeding in Bangladesh: a cross-sectional study using MICS6, Bangladesh (2019) data. *BMJ Paediatrics Open* 7.

- 23. Al Fidah MF, Nabin AA Efa SS (2024) Factors associated with acute respiratory infection and healthcare-seeking behaviour among primary caregivers in Bangladesh: a study based on MICS 2019. *BMJ Public Health* 2.
- 24. Al Fidah MF, Islam MR, Rahman MM *et al.* (2025) Community-level expenses for under-5 diarrhoeal cases prior to reporting at Dhaka Hospital: A cross-sectional analysis in Bangladesh. *Clinical Epidemiology and Global Health* 31, 101846.
- 25. Al Fidah MF, Haider T, Roy D *et al.* (2024) Temporal trends of diarrheal diseases and associated treatment responses among children aged under five: Insight from the multiple Indicator Cluster surveys from 2006 to 2019. *Pediatrics & Neonatology*.
- 26. Fatema K & Lariscy JT (2020) Mass media exposure and maternal healthcare utilization in South Asia. *SSM-Population Health* 11, 100614.
- 27. Schlotheuber A & Hosseinpoor AR (2022) Summary measures of health inequality: a review of existing measures and their application. *International journal of environmental research and public health* 19, 3697.
- 28. Islam MR, Fahim SM, Rasul MG *et al.* (2022) Health care providers' knowledge, attitude, and practice regarding facility-based management of children with severe acute malnutrition in Bangladesh. *Food and Nutrition Bulletin* 43, 465-478.
- 29. (IPHN) IoPHN (2017) National Guidelines for the Facility-based Management of Children with Severe Acute Malnutrition in Bangladesh. Dhaka, Bangladesh: Ministry of Health and Family Welfare, Bangladesh.
- 30. (IPHN) IoPHN (2011) National Guidelines for the Community-based Management of Children with Severe Acute Malnutrition in Bangladesh. Dhaka, Bangladesh: Ministry of Health and Family Welfare, Bangladesh.
- 31. Ahmad D, Afzal M Imtiaz A (2020) Effect of socioeconomic factors on malnutrition among children in Pakistan. *Future Business Journal* 6, 1-11.
- 32. Talukder A (2017) Factors associated with malnutrition among under-five children: illustration using Bangladesh demographic and health survey, 2014 data. *Children* 4, 88.
- 33. Das S & Gulshan J (2017) Different forms of malnutrition among under five children in Bangladesh: a cross sectional study on prevalence and determinants. *Bmc Nutrition* 3, 1-12.

- 34. Prasadajudio M, Devaera Y, Noormanto N *et al.* (2023) Disease-related malnutrition in pediatric patients with chronic disease: a developing country perspective. *Current Developments in Nutrition* 7, 100021.
- 35. Hasan MT, Soares Magalhaes RJ, Williams GM *et al.* (2016) The role of maternal education in the 15-year trajectory of malnutrition in children under 5 years of age in B angladesh. *Maternal & child nutrition* 12, 929-939.
- 36. Nahar MZ & Zahangir MS (2024) The role of parental education and occupation on undernutrition among children under five in Bangladesh: A rural-urban comparison. *Plos one* 19, e0307257.
- 37. Asim M & Nawaz Y (2018) Child malnutrition in Pakistan: evidence from literature. *Children* 5, 60.
- 38. Sumon IH, Hossain M, Ar Salan S *et al.* (2023) Determinants of coexisting forms of undernutrition among under-five children: Evidence from the Bangladesh demographic and health surveys. *Food Science & Nutrition* 11, 5258-5269.
- 39. Haque MA, Choudhury N, Wahid BZ *et al.* (2023) A predictive modelling approach to illustrate factors correlating with stunting among children aged 12–23 months: a cluster randomised pre-post study. *BMJ open* 13, e067961.
- 40. Srinivasan CS, Zanello G Shankar B (2013) Rural-urban disparities in child nutrition in Bangladesh and Nepal. *BMC public health* 13, 1-15.
- 41. Jana A, Dey D Ghosh R (2023) Contribution of low birth weight to childhood undernutrition in India: evidence from the national family health survey 2019–2021. BMC Public Health 23, 1336.
- 42. Huo S, Wang K, Liu Z *et al.* (2022) Influence of maternal exposure to mass media on growth stunting among children under five: mediation analysis through the water, sanitation, and hygiene program. *JMIR Public Health and Surveillance* 8, e33394.
- 43. Tette EM, Sifah EK, Nartey ET *et al.* (2016) Maternal profiles and social determinants of malnutrition and the MDGs: What have we learnt? *BMC public health* 16, 1-11.
- 44. Mehta NM, Corkins MR, Lyman B *et al.* (2013) Defining pediatric malnutrition: a paradigm shift toward etiology-related definitions. *Journal of Parenteral and Enteral Nutrition* 37, 460-481.

Table 1: Association between types of undernutrition and socio-demographic variables across survey years (2012-13, 2019)

Survey Year 2012-13							
Characteristics	Total	Types of under					
		(n=20885)					
		Absent	SFM	CFM	p-value		
		(n=10934)	(n=4219)	(n=5732)			
Age median	31 (15-45)	28 (12-45)	30 (16-44)	34 (20-46)	<0.001		
(IQR)*							
Sex (female) n	10161 (48.65)	5374 (49.15)	2033 (48.19)	2754 (48.05)	0.318		
(%)							
Residence (rural) n (%)	17557 (84.07)	8979 (82.12)	3607 (85.49)	4971 (86.72)	<0.001		
, , , ,	20462 (09.04)	10726 (09.22)	4122 (09.02)	5504 (07.69)	0.054		
Ever breastfed (yes) n (%)	20463 (98.04)	10736 (98.23)	4133 (98.03)	5594 (97.68)	0.054		
Childhood	6098 (29.2)	3201 (29.28)	1205 (28.56)	1692 (29.52)	0.564		
illness (yes) n	0000 (20:2)	3201 (23:20)	1200 (20.00)	1092 (29.02)	0.001		
(%)**							
Mothers age in ye	ears				<0.001		
15-19	1230 (6.03)	643 (6.02)	255 (6.19)	332 (5.94)	_		
20-24	5932 (29.10)	3218 (30.14)	1162 (28.19)	1552 (27.77)	-		
25-29	6622 (32.48)	3487 (32.66)	1343 (32.58)	1792 (32.07)	_		
30-34	3599 (17.65)	1849 (17.32)	745 (18.07)	1005 (17.98)			
35-39	1994 (9.78)	1022 (9.57)	399 (9.68)	573 (10.25)			
40-44	715 (3.51)	329 (3.08)	152 (3.69)	234 (4.19)	_		
	` ′		\ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \	, ,,			

295 (1.45)	129 (1.21)	66 (1.60)	100 (1.79)				
Maternal education							
5252 (25.15)	2335 (21.36)	1113 (26.38)	1804 (31.47)	_			
3030 (14.51)	1379 (12.61)	661 (15.67)	990 (17.27)				
3219 (15.41)	1531 (14.00)	713 (16.90)	975 (17.01)				
7056 (33.79)	4044 (36.99)	1382 (32.76)	1630 (28.44)	=			
2328 (11.15)	1645 (15.04)	350 (8.30)	333 (5.81)				
				0.010			
				0.069			
18491 (88.54)	9681 (88.54)	3734 (88.50)	5076 (88.56)				
1670 (8.00)	903 (8.26)	316 (7.49)	451 (7.87)				
492 (2.36)	237 (2.17)	113 (2.68)	142 (2.48)				
111 (0.53)	61 (0.56)	28 (0.66)	22 (0.38)				
121 (0.58)	52 (0.48)	28 (0.66)	41 (0.72)	_			
181)			1	<0.001			
5801 (63.18)	3360 (65.59)	1151 (64.12)	1290 (57.00)	_			
1297 (14.13)	870 (16.98)	218 (12.14)	209 (9.24)				
2083 (22.69)	893 (17.43)	426 (23.73)	764 (33.76)	-			
xposure (n=139	95) ***	1		<0.001			
5875 (41.98)	2945 (37.67)	1243 (44.83)	1687 (49.54)	_			
	5252 (25.15) 3030 (14.51) 3219 (15.41) 7056 (33.79) 2328 (11.15) 18491 (88.54) 1670 (8.00) 492 (2.36) 111 (0.53) 121 (0.58) 181) 5801 (63.18) 1297 (14.13) 2083 (22.69) xposure (n=139)	5252 (25.15) 2335 (21.36) 3030 (14.51) 1379 (12.61) 3219 (15.41) 1531 (14.00) 7056 (33.79) 4044 (36.99) 2328 (11.15) 1645 (15.04) 18491 (88.54) 9681 (88.54) 1670 (8.00) 903 (8.26) 492 (2.36) 237 (2.17) 111 (0.53) 61 (0.56) 121 (0.58) 52 (0.48) 181) 5801 (63.18) 3360 (65.59) 1297 (14.13) 870 (16.98) 2083 (22.69) 893 (17.43) xposure (n=13995) ***	1113 (26.38) 1113 (26.38) 3030 (14.51) 1379 (12.61) 661 (15.67) 3219 (15.41) 1531 (14.00) 713 (16.90) 7056 (33.79) 4044 (36.99) 1382 (32.76) 2328 (11.15) 1645 (15.04) 350 (8.30) 1670 (8.00) 903 (8.26) 316 (7.49) 492 (2.36) 237 (2.17) 113 (2.68) 111 (0.53) 61 (0.56) 28 (0.66) 121 (0.58) 52 (0.48) 28 (0.66) 181) 5801 (63.18) 3360 (65.59) 1151 (64.12) 1297 (14.13) 870 (16.98) 218 (12.14) 2083 (22.69) 893 (17.43) 426 (23.73) xposure (n=13995) ***				

Exposed	8120 (58.02)	4872 (62.33	5) 1530 (55.1	17) 1718 (50	.46)
Wealth index					<0.001
Poorest	6260 (29.97)	2721 (24.89	1360 (32.2	24) 2179 (38	.01)
Poor	4509 (21.59)	2168 (19.83	952 (22.56	5) 1389 (24	.23)
Middle	3771 (18.06)	2004 (18.33	758 (17.97	7) 1009 (17	.60)
Rich	3415 (16.35)	2010 (18.38	,	<u> </u>	,
				<u> </u>	,
Richest	2930 (14.03)	2031 (18.58	498 (11.80	0) 401 (7.00))
Survey Year 201	9				•
Characteristics	Total	Types of und	ernutrition (n=	=23,061)	
		Absent	SFM	CFM	p-value
		(n=14661)	(n=4119)	(n=4281)	
Age median	30 (15-44)	29 (13-45)	29 (16-43)	32 (18-45)	<0.001
(IQR)*					
Sex (female) n	11133	7107 (48.48)	1937 (47.03)	2089 (48.80)	0.194
(%)	(48.28)				
Residence	18773	11749	3382 (82.11)	3642 (85.07)	<0.001
(rural) n (%)	(81.41)	(80.14)			
Ever breastfed	13016	8347 (95.90)	2407 (95.59)	2262 (94.64)	0.029
(yes) n (%)	(95.62)				
(n=13,612)					
Childhood	8148 (35.33)	5147 (35.11)	1405 (34.11)	1596 (37.28)	0.006
illness (yes) n					
(%)**					
Maternal age (n=	=22759)				<0.001
15-19	1690 (7.43)	1088 (7.51)	314 (7.74)	288 (6.83)	

20-24	6811 (29.93)	4353 (30.06)	1224 (30.16)	1234 (29.25)	
25-29	6637 (29.16)	4310 (29.76)	1147 (28.27)	1180 (27.97)	
30-34	4668 (20.51)	2957 (20.42)	816 (20.11)	895 (21.21)	
35-39	2183 (9.59)	1323 (9.14)	416 (10.25)	444 (10.52)	
40-44	575 (2.53)	339 (2.34)	100 (2.46)	136 (3.22)	
45-49	195 (0.86)	112 (0.77)	41 (1.01)	42 (1.00)	
Maternal educat	ion				<0.001
Pre-primary or	2591 (11.24)	1341 (9.15)	544 (13.21)	706 (16.49)	
none					
Primary	5555 (24.09)	3172 (21.64)	1107 (26.88)	1276 (29.81)	
Secondary	11334	7483 (51.04)	1932 (46.90)	1919 (44.83)	
	(49.15)				
Higher	3581 (15.53)	2665 (18.18)	536 (13.01)	380 (8.88)	
secondary or more					
	22712	1.12.15	2004 (0.5.5=)	1222 (22.17)	0.004
Ethnicity (9()	22543	14347	3994 (96.67)	4202 (98.15)	<0.001
(Bengali) n (%)	(97.75)	(97.86)			
Caregiver of	12535	7566 (66.07)	2345 (69.13)	2624 (70.82)	<0.001
another child	(67.58)				
(Yes) n (%)					
(n=18549)					
Wealth index					<0.001
Poorest	5750 (24.93)	3094 (21.10)	1213 (29.45)	1443 (33.71)	
Poor	4831 (20.95)	2895 (19.75)	918 (22.29)	1018 (23.78)	
Middle	4345 (18.84)	2838 (19.36)	736 (17.87)	771 (18.01)	

4303 (18.66)	2985 (20.36)	670 (16.27)	648 (15.14)			
3832 (16.62)	2849 (19.43)	582 (14.13)	401 (9.37)			
Maternal life satisfaction (n = 22752)						
19751	12787	3478 (85.75)	3486 (82.67)			
(86.81)	(88.31)					
2291 (10.07)	1287 (8.89)	442 (10.90)	562 (13.33)			
710 (3.12)	405 (2.80)	136 (3.35)	169 (4.01)			
Size at birth (n= 10920)						
7698 (70.49)	5098 (71.94)	1384 (70.54)	1216 (64.96)			
1240 (11.36)	922 (13.01)	186 (9.48)	132 (7.05)			
1982 (18.15)	1066 (15.04)	392 (19.98)	524 (27.99)			
Maternal media exposure (n = 22749)						
8471 (37.24)	4924 (34.01)	1679 (41.41)	1868 (44.29)			
14278	9552 (65.99)	2376 (58.59)	2350 (55.71)			
(62.76)						
	3832 (16.62) isfaction (n = 2 19751 (86.81) 2291 (10.07) 710 (3.12) 10920) 7698 (70.49) 1240 (11.36) 1982 (18.15) exposure (n = 2 8471 (37.24) 14278	3832 (16.62) 2849 (19.43) isfaction (n = 22752) 19751 12787 (86.81) (88.31) 2291 (10.07) 1287 (8.89) 710 (3.12) 405 (2.80) 10920) 7698 (70.49) 5098 (71.94) 1240 (11.36) 922 (13.01) 1982 (18.15) 1066 (15.04) exposure (n = 22749) 8471 (37.24) 4924 (34.01) 14278 9552 (65.99)	3832 (16.62) 2849 (19.43) 582 (14.13) isfaction (n = 22752) 19751	3832 (16.62) 2849 (19.43) 582 (14.13) 401 (9.37) isfaction (n = 22752) 19751 12787 3478 (85.75) 3486 (82.67) (86.81) (88.31) 442 (10.90) 562 (13.33) 710 (3.12) 405 (2.80) 136 (3.35) 169 (4.01) 10920) 7698 (70.49) 5098 (71.94) 1384 (70.54) 1216 (64.96) 1240 (11.36) 922 (13.01) 186 (9.48) 132 (7.05) 1982 (18.15) 1066 (15.04) 392 (19.98) 524 (27.99) exposure (n = 22749) 8471 (37.24) 4924 (34.01) 1679 (41.41) 1868 (44.29) 14278 9552 (65.99) 2376 (58.59) 2350 (55.71)		

SFM: Single form of malnutrition

CFM: Coexisting form of malnutrition

***: Fisher's exact test

^{*:} Kruskal-wallis test

^{**:} Childhood illness: diarrhoea and cough (for 2012-13); diarrhoea, fever, and cough (for 2019)

Table 2: Age and sex adjusted multinomial logistic regression models across two survey years (2012-13, 2019) presenting the association between predictor variables and types of malnutrition

Survey year	2012-13				
Predictor va	riables	SFM		CFM	
		aOR (95% CI)	p-value	aOR (95% CI)	p-value
Residence	Rural	Ref			
	Urban	0.78 (0.71-0.86)	<0.001	0.70 (0.64-0.77)	<0.001
Mother's	15-19	Ref			
age in years	20-24	0.87 (0.74-1.02)	0.083	0.79 (0.68-0.92)	0.002
	25-29	0.91 (0.77-1.07)	0.248	0.80 (0.69-0.93)	0.003
	30-34	0.95 (0.80-1.12)	0.519	0.83 (0.71-0.97)	0.019
	35-39	0.91 (0.75-1.10)	0.323	0.84 (0.70-0.99)	0.040
	40-44	1.07 (0.84-1.36)	0.608	1.03 (0.83-1.28)	0.795
	45-49	1.17 (0.84-1.64)	0.353	1.10 (0.82-1.48)	0.544
Maternal	None	Ref			
education	Primary incomplete	1.01 (0.90-1.14)	0.845	0.95 (0.86-1.05)	0.323
	Primary complete	0.99 (0.88-1.11)	0.828	0.86 (0.77-0.95)	0.003
	Secondary incomplete	0.73 (0.66-0.80)	<0.001	0.55 (0.50-0.59)	<0.001
	Secondary complete or higher	0.45 (0.40-0.52)	<0.001	0.28 (0.24-0.32)	<0.001
Size at	Average	Ref		l	<u> </u>

birth	Larger	0.73 (0.62-0.85)	<0.001	0.62 (0.52-0.73)	<0.001
(n=9181)	Smaller	1.42 (1.24-1.62)	<0.001	2.32 (2.06-2.61)	<0.001
Maternal	Not exposed	Ref			
media	-				
exposure	Exposed	0.74 (0.68-0.81)	<0.001	0.61 (0.56-0.66)	<0.001
(n=13995)					
XX7 - 141-	Doomost	Dof			
Wealth	Poorest	Ref			
index	Poor	0.88 (0.80-0.97)	0.014	0.81 (0.74-0.88)	<0.001
	Middle	0.76 (0.68-0.85)	<0.001	0.64 (0.56-0.79)	<0.001
	Rich	0.65 (0.58-0.73)	<0.001	0.47 (0.43-0.52)	<0.001
	Kicii	0.03 (0.36-0.73)	<0.001	0.47 (0.43-0.32)	\0.001
	Richest	0.49 (0.44-0.55)	<0.001	0.25 (0.22-0.28)	<0.001
Survey year	2019			<u> </u>	
Predictor va	riables	SFM		CFM	
Predictor va	riables	SFM aOR (95% CI)	p-value	CFM aOR (95% CI)	p-value
Predictor variable Residence	riables Rural		p-value		p-value
		aOR (95% CI)	p-value 0.005		p-value <0.001
	Rural	aOR (95% CI)		aOR (95% CI)	_
Residence	Rural Urban No	aOR (95% CI) Ref 0.88 (0.80-0.96) Ref	0.005	aOR (95% CI) 0.71 (0.65-0.78)	<0.001
Residence Ever	Rural Urban	aOR (95% CI) Ref 0.88 (0.80-0.96)		aOR (95% CI)	_
Residence Ever breastfed	Rural Urban No	aOR (95% CI) Ref 0.88 (0.80-0.96) Ref	0.005	aOR (95% CI) 0.71 (0.65-0.78)	<0.001
Residence Ever breastfed (n=13612)	Rural Urban No Yes	aOR (95% CI) Ref 0.88 (0.80-0.96) Ref 1.07 (0.86-1.33)	0.005	aOR (95% CI) 0.71 (0.65-0.78)	<0.001
Ever breastfed (n=13612) Childhood	Rural Urban No Yes No	aOR (95% CI) Ref 0.88 (0.80-0.96) Ref 1.07 (0.86-1.33)	0.005	aOR (95% CI) 0.71 (0.65-0.78) 0.96 (0.78-1.19)	< 0.001 0.721
Residence Ever breastfed (n=13612) Childhood illness	Rural Urban No Yes No Yes 15-19	aOR (95% CI) Ref 0.88 (0.80-0.96) Ref 1.07 (0.86-1.33) Ref 0.96 (0.89-1.03)	0.005	aOR (95% CI) 0.71 (0.65-0.78) 0.96 (0.78-1.19) 1.14 (1.06-1.23)	<0.001 0.721 <0.001
Ever breastfed (n=13612) Childhood illness Mother's	Rural Urban No Yes No Yes 15-19 20-24	aOR (95% CI) Ref 0.88 (0.80-0.96) Ref 1.07 (0.86-1.33) Ref 0.96 (0.89-1.03)	0.005 0.556 0.242	aOR (95% CI) 0.71 (0.65-0.78) 0.96 (0.78-1.19)	< 0.001 0.721
Ever breastfed (n=13612) Childhood illness Mother's age in years	Rural Urban No Yes No Yes 15-19	aOR (95% CI) Ref 0.88 (0.80-0.96) Ref 1.07 (0.86-1.33) Ref 0.96 (0.89-1.03)	0.005	aOR (95% CI) 0.71 (0.65-0.78) 0.96 (0.78-1.19) 1.14 (1.06-1.23)	<0.001 0.721 <0.001

35-39 40-44 45-49 Maternal Pre-prin	1	.09 (0.92-1.29) .02 (0.79-1.32) .26 (0.86-1.84)	0.347 0.899 0.239	1.11 (0.93-1.32) 1.29 (1.10-1.64) 1.19 (0.82-1.75)	0.244 0.041
45-49	1	.26 (0.86-1.84)		, ,	
			0.239	1.19 (0.82-1.75)	
Maternal Pre-prir	nary or F	Ref			0.362
education none					
Primary	C	0.86 (0.76-0.97)	0.014	0.78 (0.70-0.87)	<0.001
Seconda	ary C	0.64 (0.57-0.71)	<0.001	0.50 (0.45-0.55)	<0.001
Higher or more		0.50 (0.43-0.57)	<0.001	0.28 (0.24-0.32)	<0.001
Ethnicity Others	F	Ref			
Bengali	. 0	0.70 (0.57-0.86)	<0.001	1.19 (0.93-1.52)	0.177
Caregiver No	F	Ref			
of another Yes child (n=18549)	1	.17 (1.08-1.27)	<0.001	1.25 (1.15-1.35)	<0.001
Size at Average	e F	Ref			
hirth			.0.001	0.60 (0.50, 0.72)	.0.001
(n=10920) Larger		0.74 (0.63-0.88)	<0.001	0.60 (0.50-0.73)	<0.001
Smaller	1	.18 (1.21-1.57)	<0.001	2.11 (1.87-2.39)	<0.001
Maternal Not exp	oosed F	Ref			
media Expose	d C	0.73 (0.68-0.78)	<0.001	0.65 (0.60-0.69)	<0.001
exposure (n=22749)					
Wealth Poorest	F	Ref			
index Poor	C	0.81 (0.73-0.89)	<0.001	0.75 (0.69-0.83)	<0.001

	Middle	0.66 (0.60-0.74)	<0.001	0.58 (0.53-0.65)	<0.001
	Rich	0.57 (0.51-0.64)	<0.001	0.47 (0.42-0.52)	<0.001
	Richest	0.52 (0.47-0.58)	<0.001	0.30 (0.27-0.34)	<0.001
Maternal	Neutral	Ref	<u> </u>		<u> </u>
life					
inc	Satisfied	0.79 (0.71-0.89)	< 0.001	0.63 (0.57-0.70)	< 0.001
satisfaction					
(n = 22752)	Dissatisfied	0.98 (0.78-1.22)	0.853	0.94 (0.77-1.16)	0.581

Reference level: Type of malnutrition 'absent'

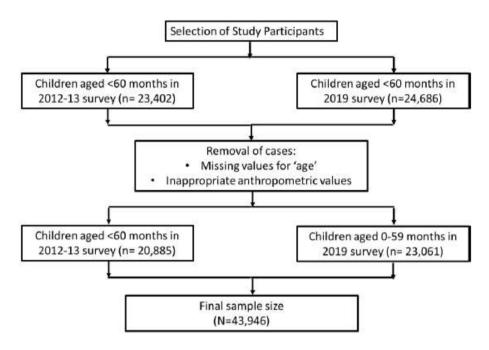


Figure 1: Selection of study participants using combined Bangladesh MICS data from 20012-13 and 2019

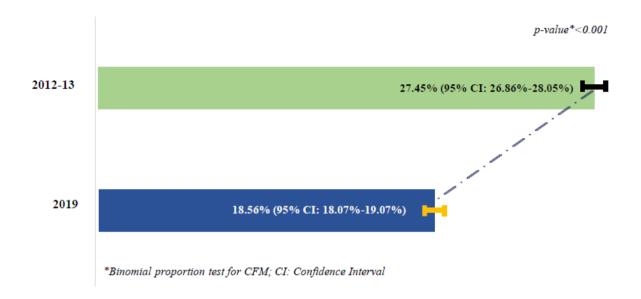
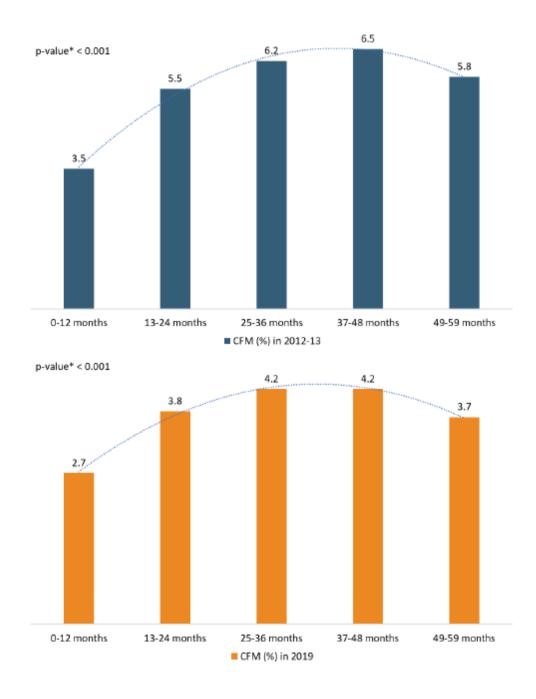
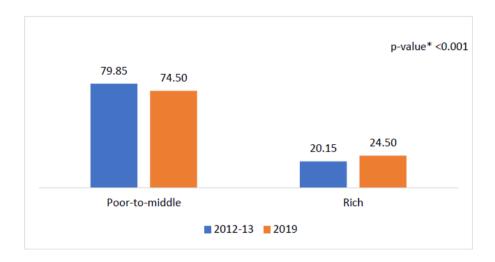


Figure 2: Prevalence of CFM across survey years (2012-13, 2019)



^{*:} p-value obtained using Cochran-Armitage trend test

Figure 3: Trend of CFM across participant's age groups in survey years (2012-13, 2019)

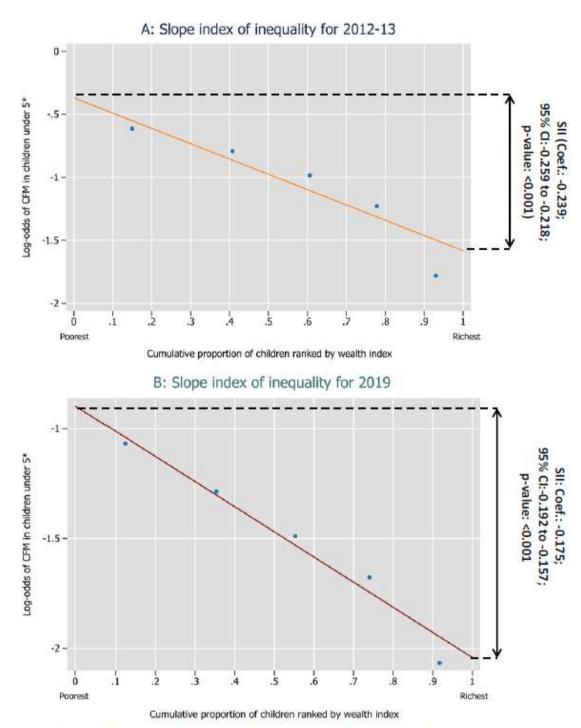


^{*:}p-value obtained using χ^2 goodness-of-fit test

Rich: Participants from rich and richest quintile families

Poor-to-middle: Participants from poorest, poor, and middle quintile families

Figure 4: Proportion of wealthier quintile families among participants with CFM



^{*:} Represents the log-odds (likelihood) of children under 5 having CFM (-2 indicates a lesser likelihood of CFM among richer quintiles);

Figure 5: Slope index of inequality (SII) for the prevalence of coexisting forms of malnutrition according to wealth quintiles across survey years; A: For 2012-13; B: For 2019

SII: Slope index of inequality; Coefficient: Derived using logistic regression model using two variables (coexisting forms of malnutrition and wealth index); CI: Confidence Interval