

Development of severe psychological distress among low-income individuals during the COVID-19 pandemic: longitudinal study

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Background

It has been indicated that the health impact of COVID-19 is potentially greater in individuals from lower socioeconomic status than in the overall population.

Aims

To examine how the spread of COVID-19 has altered the general public's mental health, and whether such changes differ in relation to individual income.

Method

An online longitudinal survey was conducted at three different time periods during the pandemic. We recruited 1993 people aged 20–70 years, living in the Tokyo metropolitan area in Japan. Participants' mental health was measured with the six-item version of the Kessler Psychological Distress Scale; the existence of severe psychological distress was ascertained through the cut-off data. Multiple logistic and mixed-model ordinal logistic regression analyses were performed, with income as the independent variable.

Results

Of the participants, 985 were male, with a mean age of 50.5 (± 15.8) years. Severe psychological distress percentages for each tested period were 9.3%, 11.2% and 10.7% for phases 1, 2 and 3, respectively. Between phases 1 and 2 or phases 2 and 3, the group that earned $<£15\,000$ had significantly higher

propensity to develop severe psychological distress than the group that earned $\geq£45\,000$ (odds ratio 2.09, 95% CI 0.95–4.56 between phases 1 and 2; odds ratio 3.00, 95% CI 1.01–9.58 between phases 2 and 3).

Conclusions

Although there has been significant deterioration in mental health among citizens during the COVID-19 pandemic, this was more significant among those with lower income. Therefore, mental health measures that focus on low socioeconomic groups may be necessary.

Keywords

Novel coronavirus; general population; inequality; socioeconomic status; mental health.

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COVID-19 spread from China to the rest of the world, and was declared a pandemic by the World Health Organization (WHO) in March 2020. As of 15 June 2020, the total number of people who had contracted COVID-19 was 7.5 million, globally, with over 300 000 deaths from the disease.¹ This pandemic negatively affects the mental health of the general public by instilling fear of infection, isolation and death; lowering access to mental healthcare; increasing the sense of isolation through physical distancing; and affecting economic activities (e.g. leading to a decline in wages and/or redundancy).^{2–4} The deterioration in individuals' mental health caused by COVID-19 has been identified as holding a risk for the development of secondary issues, such as increases in alcoholism, domestic violence and suicide.^{2,5,6} On 13 May 2020, the WHO announced its 'Policy Brief: COVID-19 and the Need for Action on Mental Health', which advocated tackling mental health measures parallel to conducting transmission control as part of COVID-19 countermeasures.⁷

General impact of COVID-19

Some have opined that the health impact of the COVID-19 pandemic has not been uniform across the board; rather, its greatest impact seems to be among those from lower socioeconomic groups.⁸ There are, inherently, 'social determinants of health inequality' that determine individuals' health in the low

socioeconomic group, which include poverty, unhygienic environments, unhealthy diet, few educational opportunities and low employment skills.⁹ The recent COVID-19 pandemic has also caused various economic crises that have resulted in decline in wages, redundancy and a general loss of employment across various industries.¹⁰ It should be noted that during the epidemics of the Spanish Flu (H1N1 influenza A virus) in the 1920s and severe acute respiratory syndrome in 2003, the number of suicide deaths increased as a result of the economic impact of the infectious disease epidemicity, which led to a deterioration in mental health.^{11,12} Similarly, it can be hypothesised that during the COVID-19 pandemic, the low socioeconomic group, with its increased financial vulnerability, is a high-risk group compared with higher socioeconomic groups. This hypothesis is based on how the former may be more likely to experience worsening of their mental health, which, in turn, could lead to a propensity toward more serious situations, such as suicide.

Impact on mental health

Considering the health impact on the lower socioeconomic group, it can also be hypothesised that deterioration in mental health would occur compared with higher socioeconomic groups. People with low income would be hesitant to seek medical care because of the related healthcare costs, even when they feel severe anxiety or depression. In addition, their occupations tend to be in restaurants

or grocery stores, where remote working is not allowed. Furthermore, a report indicates that such workplaces do not provide sufficient personal protective equipment, such as face masks.¹³ Therefore, it can be said that they have no choice but to work in such an environment, to make a living, even if there is fear of COVID-19 infection or death.

Although there have been some cross-sectional studies¹⁴ that have examined the correlation between socioeconomic factors and mental health during the COVID-19 pandemic, there are almost no longitudinal studies in this regard. Wang et al conducted a study at two time points, with a total of 1738 participants; however, the individuals examined at the two time points were different.¹⁵ As a result, the study could not discuss how mental health had changed among individuals. To examine whether the impact of the COVID-19 pandemic on mental health is greater among those of low socioeconomic standing compared with those in higher socioeconomic groups, it is necessary to conduct a longitudinal study that follows individuals to capture any changes in their mental health. However, currently, there is insufficient evidence to support this, and there have been indications that further research is needed.^{6,16} To that end, this study employed longitudinal data that followed members of the Japanese general public in the Tokyo metropolitan area, at three different time points. Through such measures, we clarify how the mental health of citizens has changed during the COVID-19 pandemic, and examine whether the degree of such changes differs depending on an individual's socioeconomic status.

Method

Study sample and data collection

This was an online longitudinal study conducted with members of the Japanese population. The details of this study are only briefly addressed here, since the subject recruitment method is described in more detail in our previous study.¹⁷ The 8156 individuals who were approached to be part of this study were those registered with My Voice Co., Ltd., an online survey company. The participants reside in the seven prefectures (Ibaraki, Tochigi, Gunma, Saitama, Chiba, Kanagawa and Tokyo) of the Tokyo metropolitan area. To obtain responses from a total of 2400 people, 200 respondents from each gender/age group (i.e. 20–29, 30–39, 40–49, 50–59, 60–69 and 70–79 years) were selected. First, the study's questionnaire was uploaded onto a secured online platform and the online survey company sent the questionnaire's address page for responses to its registered users. Then, the respondents who received the address could access the online questionnaire, and respond voluntarily. The responses were closed at the point where the set quotient had been met (i.e. 200 respondents per each gender/age group). It should be noted that the survey respondents registered on the site are paid points equivalent to 50 Japanese Yen (approximately 40 pence) for each completed survey.

Survey dates with respect to the spread of COVID-19 in Japan

This study was conducted at three different time points (phases 1–3) (Fig. 1). The baseline survey (phase 1) was conducted between 25 and 27 February 2020. These dates corresponded with the early phase of the COVID-19 outbreak in Japan. The infection cases in Japan, up to that point, were mainly people returning from outbreak areas (e.g. China) and those who had come into contact with such people. As such, there were not many cases whose route of infection could not be traced. The total number of patients infected with COVID-19 up until 25 February 2020, the day before the survey, was 157 patients in Japan, with 1 death (this death was of a

patient who had been infected on a cruise ship from China that had stopped at a Japanese port, and was not a case where COVID-19 was contracted within Japan).¹

A follow-up study (phase 2) of the respondents who took part in phase 1 was conducted between 1 and 6 April 2020. Between phases 1 and 2, the WHO had pronounced COVID-19 to be a pandemic (11 March 2020). Within Japan, the outbreak had also shifted from the initial phase (i.e. from when a sporadic outbreak was seen) to a community transmission phase (i.e. when the number of patients per day kept increasing exponentially). During this period, the Japanese government started taking infection control measures, such as requesting schools to close temporarily and for companies to work remotely. On 7 April 2020, the final research date, a state of emergency was declared within the Tokyo metropolitan area.^{18,19} The total number of infected patients in Japan at the start of phase 2 was 4477, with a total of 98 deaths.

Phase 3 was conducted between 12 and 17 May 2020. The COVID-19 outbreak in Japan between phases 2 and 3 saw the number of newly infected persons begin to decrease, and the pandemic was believed to have shifted to its 'post-peak' phase. A state of emergency was declared during this period, where Japanese citizens refrained from going outside and companies and schools widely continued to halt their operations. By 11 May 2020, the day before phase 3 research took place, there were 16 014 infected patients and 657 deaths.

Measurement

Assessment of severe psychological distress

In both the baseline and follow-up surveys, the six-item Kessler Psychological Distress Scale (K6) was used to measure severe psychological distress (SPD).²⁰ The K6 is broadly used in epidemiological studies²¹ because it measures psychological distress in the general population by using six simple items. Each item measures the extent of general, non-specific psychological distress, using a five-point response (0 'none of the time', 1 'a little of the time', 2 'some of the time', 3 'most of the time', and 4 'all of the time'); thus, the total scores ranged from 0 to 24. The K6 was translated into Japanese, and a previous study of 164 Japanese adults proved its internal consistency in relation to reliability (Cronbach's alpha, 0.849) and validity (100% sensitivity and 69.3% specificity for screening mood and anxiety disorder).²² This study used an established protocol to define a score of ≥ 13 as having SPD.²⁰

Assessment of annual income levels

In this study, individual annual income was set as an independent variable. Annual income was divided into four categories: <¥2 million (approximately £15 000), ¥2 million to <¥4 million (£15 000 to <£30 000), ¥4 million to <¥6 million (£30 000 to <£45 000) and \geq ¥6 million (\geq £45 000). Information on annual income was provided to the research team in phase 1, as the research company investigated and ascertained each participating individual's annual income for the 6 months before the start of this research.

Covariates

In the baseline survey, participants reported their gender, age, residential area (Northern Kanto area [Ibaraki, Tochigi, Gunma Prefectures], Saitama Prefecture, Chiba Prefecture, Kanagawa Prefecture, Tokyo Metropolis Area), working status (working, not working), marital status (single, divorced, separated, married), living arrangements (alone, with others but without children, with children aged ≥ 18 years, with others and children <18 years), smoking status (smokers, ex-smokers, non-smokers), alcohol consumption (never, seldom (one to four times per week), often (five

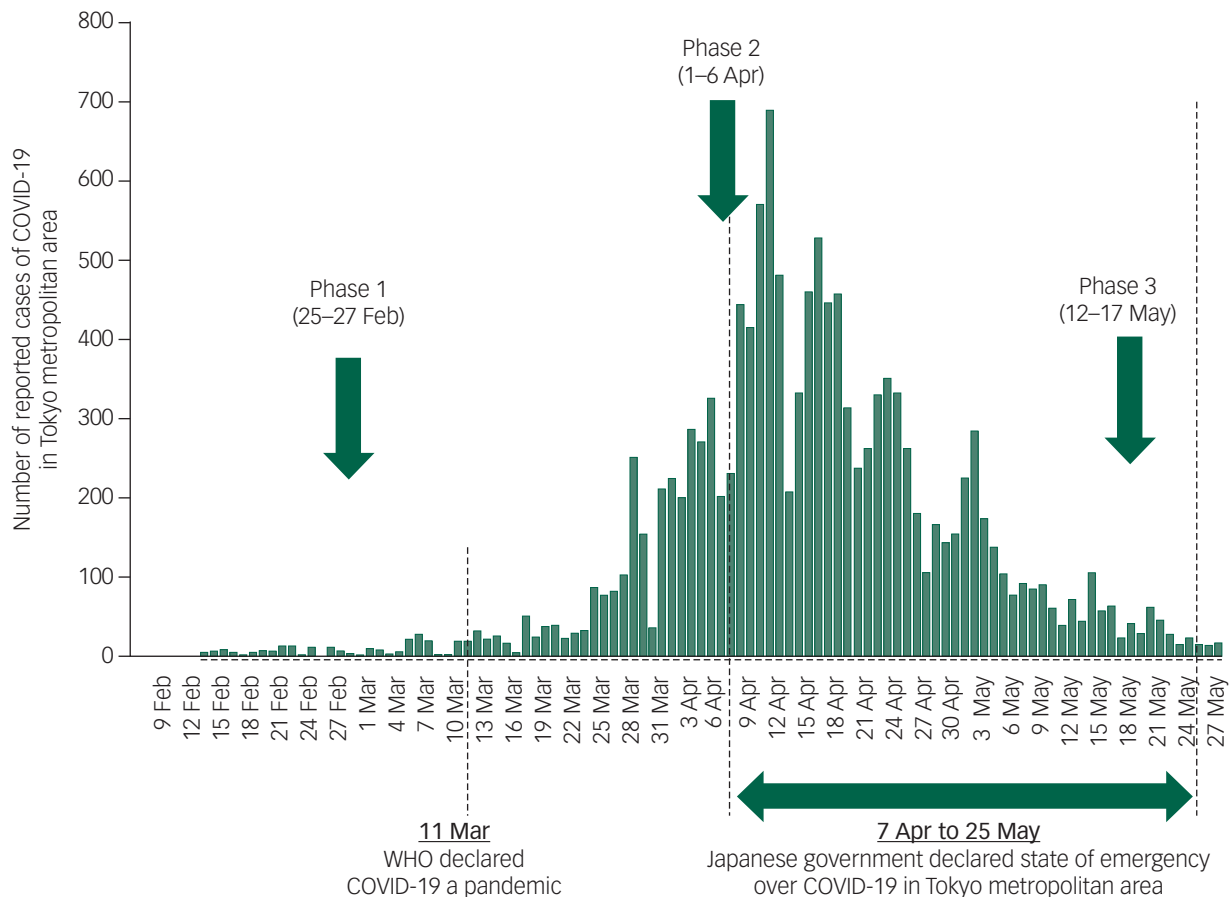


Fig. 1 Timeline of COVID-19 spread in Japan and study time points. WHO, World Health Organization.

to seven times per week), daily walking time (<30 mins, 30–59 mins, 60 mins), regular annual vaccination (yes, no) and past medical history (hypertension, diabetes, heart disease, stroke, respiratory disease, kidney disease, cancer). In addition, the research company provided categorised data of educational attainment (junior or high school graduate, junior college graduate, university graduate or above, other).

Statistical analysis

SPD percentage was calculated for both the overall and individual factors for each research period (i.e. from phase 1 to phase 3), and McNemar's test was conducted to investigate whether there was any significant change in the SPD percentages between the phases. The two-group comparison between the phases was examined with the Bonferroni (multiple comparison) method.

Thereafter, a multiple logistic regression model was used to examine the correlation between annual income and SPD, considering the covariates. First, the relationship between individual annual income and SPD was examined by making use of phase 1 data. Subsequently, a longitudinal analysis was performed, excluding respondents who had SPD in phase 1. The aim of the analysis was to determine whether new patients had developed SPD between phases 1 and 2. Finally, respondents with SPD at phase 1 or 2 were then excluded, to perform an analysis with SPD onset between phase 2 and phase 3 as the outcome. All covariates were fed into each model simultaneously. Up to four consecutive values were applied to each category of annual income, and the existence of linear tendency was examined in relation to annual income and SPD. An analysis that limited the subject to workers was also conducted as a sensitivity analysis.

Finally, to assess whether SPD status differed by income throughout this survey period, mixed-effects ordinal logistic regression analyses were performed by nesting each participant.²³ This analysis tested whether the trajectories of SPD proportion in the three surveys differed by income. In this analysis, fixed effects for all individual factors were estimated (model 1), and estimated by further adjusting K6 score at phase 1 to consider the mean difference of the K6 score at phase 1 (model 2). All analyses were performed with Stata for Windows (version 15.0).

Ethical Approval

This study was approved by the Ethics Committee of Tokyo Medical University, Tokyo, Japan (approval number T2019-0234). Informed consent was obtained from all respondents.

Results

Table 1 shows the sociodemographic characteristics of the participants and their SPD percentages during the baseline and follow-up surveys. Of the 1993 participants, 985 (49.5%) were men, the average age was 50.5 (s.d. 15.8) years and approximately 37.2% were workers.

The SPD percentage of each survey was 9.3% for phase 1, 11.2% for phase 2 and 10.7% for phase 3. McNemar's test indicated a significant deterioration in SPD percentages in participants overall, between phases 1 and 2. In terms of annual income, the percentages of SPD among those earning \geq £45 000 were 4.2% (phase 1), 6.9% (phase 2) and 5.9% (phase 3), indicating no significant difference in the SPD percentages between phases. Conversely, the percentages

Table 1 Proportion of severe psychological distress by individual factor

	<i>n</i>		<i>%</i>		Severe psychological distress (K6 ≥ 13)								
					Phase 1 (25–27 Feb)		Phase 2 (1–7 Apr)		Phase 3 (12–17 May)		McNemar's test		
					<i>n</i>	<i>%</i>	<i>n</i>	<i>%</i>	<i>n</i>	<i>%</i>	1 v. 2	1 v. 3	2 v. 3
Overall	1993				186	9.33%	224	11.24%	213	10.69%	0.005*	0.059	0.431
Gender													
Male	985	49.47%	95	9.64%	110	11.17%	103	10.46%	0.120	0.409	0.477		
Female	1008	50.63%	91	9.03%	114	11.31%	110	10.91%	0.016*	0.071	0.686		
Age, years													
20–29	268	13.46%	45	16.79%	54	20.15%	41	15.30%	0.297	0.537	0.069		
30–39	345	17.33%	60	17.39%	57	16.52%	68	19.71%	0.895	0.276	0.101		
40–49	352	17.68%	35	9.94%	46	13.07%	48	13.64%	0.048	0.063	0.758		
50–59	342	17.18%	24	7.02%	29	8.48%	28	8.19%	0.336	0.433	0.835		
60–69	355	17.83%	13	3.66%	23	6.48%	15	4.23%	0.499	0.655	0.088		
70–79	331	16.62%	9	2.72%	15	4.53%	13	3.93%	0.058	0.285	0.564		
Annual income, £													
<15 000	896	45.00%	101	11.27%	121	13.50%	130	14.51%	0.028+	0.005*	0.361		
15 000 to <30 000	504	25.31%	52	10.32%	58	11.51%	46	9.13%	0.453	0.446	0.128		
30 000 to <45 000	304	15.27%	21	6.91%	25	8.22%	20	6.58%	0.479	0.842	0.225		
≥45 000	289	14.52%	12	4.15%	20	6.92%	17	5.88%	0.074	0.166	0.491		
Residential area													
Northern Kanto (Ibaraki, Tochigi, Gunma Prefectures)	177	8.89%	19	10.73%	22	12.43%	19	10.73%	0.439	0.999	0.439		
Saitama Prefecture	321	16.12%	34	10.59%	41	12.77%	39	12.15%	0.317	0.423	0.732		
Chiba Prefecture	291	14.62%	24	8.25%	25	8.59%	32	11.00%	0.862	0.182	0.194		
Tokyo Metropolis	771	38.72%	67	8.69%	91	11.80%	76	9.86%	0.003*	0.249	0.067		
Kanagawa Prefecture	433	21.75%	42	9.70%	45	10.39%	47	10.85%	0.555	0.475	0.777		
Working status													
No	741	37.22%	61	8.23%	71	9.58%	80	10.80%	0.012*	0.015*	0.260		
Yes	1252	62.88%	125	9.98%	146	11.66%	133	10.62%	0.080	0.261	0.800		
Marital status													
Single, divorced, separated	834	41.89%	119	14.27%	134	16.07%	131	15.71%	0.127	0.261	0.770		
Married	1159	58.21%	67	5.78%	90	7.77%	82	7.08%	0.015*	0.116	0.399		
Living arrangement													
Living alone	384	19.29%	43	11.20%	50	13.02%	46	11.98%	0.336	0.680	0.516		
Living with others but without children	945	47.46%	96	10.16%	119	12.59%	120	12.70%	<0.001*	0.020+	0.923		
Living with children aged ≥18 years	335	16.83%	12	3.58%	19	5.67%	16	4.78%	0.127	0.371	0.467		
Living with children aged <18 years	329	16.52%	35	10.64%	36	10.94%	31	9.42%	0.879	0.493	0.384		
Education (years)													
Junior or high school graduate (≤12 years)	468	23.51%	54	11.54%	64	13.68%	56	11.97%	0.101	0.790	0.258		
Junior college graduate (13–15 years)	422	21.20%	32	7.58%	41	9.72%	51	12.09%	0.114	0.004*	0.105		
University graduate or above (≥16 years)	1078	54.14%	96	8.91%	117	10.85%	103	9.55%	0.050	0.495	0.170		
Other	25	1.26%	4	16.00%	2	8.00%	3	12.00%	0.317	0.999	0.564		
Smoking status													
Smoker	301	15.12%	29	9.63%	36	11.96%	28	9.30%	0.206	0.862	0.144		
Ex-smoker	296	14.87%	23	7.77%	28	9.46%	21	7.09%	0.273	0.683	0.162		
Non-smoker	1396	70.12%	134	9.60%	160	11.46%	164	11.75%	0.025+	0.014*	0.735		
Alcohol consumption													
None	843	42.34%	98	11.63%	101	11.98%	108	12.81%	0.838	0.297	0.473		
Seldom (1–4 days per week)	711	35.71%	59	8.30%	78	10.97%	76	10.69%	0.009*	0.047	0.803		
Often (5–7 days per week)	439	22.05%	29	6.61%	45	10.25%	29	6.61%	0.015*	0.999	0.008*		
Walking time (mins per day)													
<30	1005	50.48%	112	11.14%	133	13.23%	125	12.44%	0.039	0.229	0.424		
30–59	659	33.10%	47	7.13%	63	9.56%	58	8.80%	0.052	0.138	0.535		
≥60	329	16.52%	27	8.21%	28	8.51%	30	9.12%	0.602	0.602	0.715		
Regular vaccinations													
No	1115	56.00%	113	10.13%	128	11.48%	131	11.75%	0.119	0.098	0.776		
Yes	878	44.10%	73	8.31%	96	10.93%	82	9.34%	0.014*	0.335	0.127		
Comorbidities													
Hypertension	381	19.14%	24	6.30%	32	8.40%	28	7.35%	0.178	0.450	0.371		
Diabetes	118	5.93%	10	8.47%	12	10.17%	5	4.24%	0.480	0.096	0.020+		
Heart disease	58	2.91%	5	8.62%	9	15.52%	5	8.62%	0.059	0.999	0.046		
Stroke	18	0.90%	4	22.22%	4	22.22%	3	16.67%	1.000	0.317	0.317		
Respiratory disease	83	4.17%	14	16.87%	19	22.89%	16	19.28%	0.157	0.564	0.366		
Kidney disease	10	0.50%	1	10.00%	2	20.00%	1	10.00%	0.317	0.999	0.317		
Cancer	41	2.06%	2	4.88%	3	7.32%	3	7.32%	0.564	0.317	0.999		

The *P*-values with + and* indicate statistical significance levels of +*P*<0.1 (i.e. *P*<0.033) and **P*<0.05 (i.e. *P*<0.017) after Bonferroni correction, respectively. K6, six-item Kessler Psychological Distress Scale.

of SPD in the group earning <£15 000 were 11.3% (phase 1), 13.5% (phase 2) and 14.5% (phase 3), indicating a significant upward trend between phases 1 and 2, and a significant increase in SPD between phases 1 and 3.

The results of the logistic regression analysis are shown in Table 2 (also see Supplementary Table 1 available at <https://doi.org/10.1192/bjo.2021.5>). As an outcome of examining the relationship between SPD prevalence and annual income using phase 1 data, the odds ratio of SPD prevalence in the group with an annual income of £15 000 to <£30 000 compared with the group that reported an annual income of ≥£45 000 was 2.44 (95% CI 1.19–5.00), and the odds ratio in the group with an annual income of <£15 000 was 3.03 (95% CI 1.44–6.36). Among those without SPD at phase 1 (*n* = 1607), 121 (7.5%) had developed SPD at phase 2. Multivariate adjusted logistic regression analysis showed that those with an annual income of <£15 000 were significantly more likely to develop SPD than those with an annual income of ≥£45 000 (odds ratio 2.09, 95% CI 0.95–4.56). Then, after excluding those with SPD in phase 1 and/or phase 2 (*n* = 193), a significantly higher odds ratio was also observed at phase 3, among the group earning <£15 000 (odds ratio 3.00, 95% CI 1.01–9.58).

Finally, a mixed-model ordinal logistic regression analysis showing increased likelihood of developing SPD was observed among the lower income group (Fig. 2) (see also Supplementary Fig. 1 and Supplementary Table 2). In model 1, compared with those with higher income (≥£45 000 of annual personal income), a significantly high likelihood to develop SPD could be observed among those in the lower (£15 000 to £30 000: odds ratio 2.92, 95% CI 1.25–6.77) and the lowest income category (<£15 000: odds ratio 5.56, 95% CI 2.29–13.46). In model 2, estimated odds ratios were attenuated but the results remained significant in the lowest category (odds ratio 2.30, 95% CI 1.27–4.18). The results of the sensitivity analysis in which the participants were limited to workers was similar to those for all participants.

In addition, attrition analysis showed no significant differences in the study results, even with the younger participants who dropped out of the study and were lost to follow-up.

Discussion

Summary of findings

This study followed 1993 members of the Japanese public in the Tokyo metropolitan area at three different time points (phases 1–3), to examine whether there was an onset of SPD caused by the COVID-19 pandemic, depending on annual income level. Results indicated that SPD was more prevalent among those in the lower annual income group than for those in the higher annual income group. Furthermore, lower annual income groups were most likely to experience newly developed SPD (mental health deterioration) even between phases 1 and 2 and phases 1 and 3. These results suggest that the effect of the COVID-19 pandemic on the mental health of citizens shows more deterioration among those with low annual income.

Comparison with past findings

To the best of our knowledge, there has been only one study that investigated the changes in mental health among general people during the COVID-19 pandemic. Pierce et al used data from 53 351 people in the UK and reported that the average score of General Health Questionnaire in 2020 was significantly worse than the score in 2018–2019.²⁴ However, contrary to our findings, they found no clear pattern of variation in changes according to

Table 2 Adjusted odds ratios of prevalence and incidence of severe psychological distress by annual income: multivariable logistic regression results

Annual income in £10 000	Phase 1 (cross-sectional)				Phase 1 to phase 2 (longitudinal)				Phase 2 to phase 3 (longitudinal)										
	<i>n</i>	SPD (<i>K6</i> ≥ 13)	Prevalence	Odds ratio	95%CI	<i>P</i> -value	<i>n</i>	SPD (<i>K6</i> ≥ 13)	Incidence	Odds ratio	95% CI	<i>P</i> -value	<i>n</i>	SPD (<i>K6</i> ≥ 13)	Incidence	Odds ratio	95% CI	<i>P</i> -value	
<1.5	896	101	11.3%	3.03	(1.44–6.36)	0.003	795	57	7.2%	2.09	(0.95–4.56)	0.067	738	42	5.7%	3.11	(1.01–9.58)	0.048	
1.5 to <3.0	504	52	10.3%	2.44	(1.19–5.00)	0.015	452	33	7.3%	1.57	(0.75–3.31)	0.234	419	21	5.0%	2.47	(0.83–7.40)	0.106	
3.0 to <4.5	304	21	6.9%	1.41	(0.66–3.01)	0.377	283	17	6.0%	1.12	(0.52–2.42)	0.776	266	4	1.5%	0.60	(0.15–2.37)	0.467	
≥4.5	289	12	4.2%	1.00			277	14	5.1%	1.00			263	5	1.9%	1.00			
Trend <i>P</i> -value										<0.001						<0.001			

Odds ratios were adjusted by gender, age, residential area, marital status, living arrangement, education, smoking status, alcohol consumption, walking time, comorbidities (hypertension, diabetes, heart disease, stroke, respiratory disease, cancer) and regular vaccination. SPD, severe psychological distress; *K6*, six-item Kessler Psychological Distress Scale.

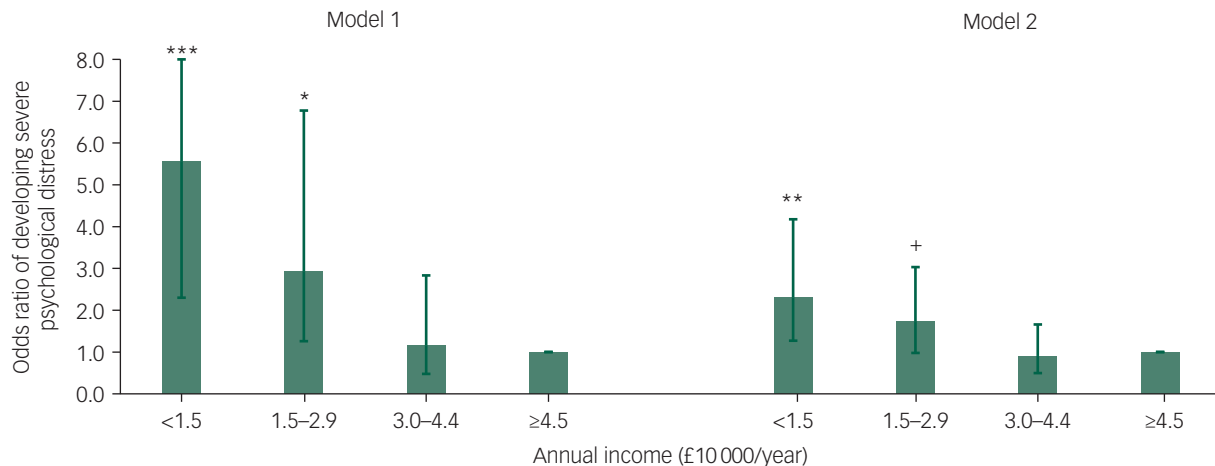


Fig. 2 Adjusted odds ratios of developing severe psychological distress in three phases: mixed-model ordinal logistic regression results. Model 1: odds ratios were adjusted by gender, age, residential area, marital status, living arrangement, education, smoking status, alcohol consumption, walking time, comorbidities (hypertension, diabetes, heart disease, stroke, respiratory disease, kidney disease, cancer) and regular vaccination. Model 2 was further adjusted by total Kessler Psychological Distress Scale score at baseline. *** $P < 0.001$, ** $P < 0.01$, * $P < 0.05$, + $P < 0.1$.

income. The impact of the COVID-19 pandemic on the lower income population would be different in each country, which may speculatively be because of the difference in the number of COVID-19 cases or lockdown policies.

Furthermore, this study also showed that those in the young age group (20–40 years) significantly developed SPD compared with older adults (60–70 years; the common retirement age is 65 years in Japan). Considering how previous studies have shown that mental illness is more likely to develop in young people,²⁵ younger people with low income may have an excessive risk for developing SPD during the COVID-19 pandemic.

Possible mechanism

In response to the COVID-19 pandemic, Japan announced the ‘Basic Policies for Novel Coronavirus Disease Control by the Government of Japan’ on 28 March 2020.¹⁹ These policies requested that Japanese citizens avoid nonessential outings, reduce social interactions and, if possible, work remotely. These requests were also extended to the operation of services where people gather. From this point onward, various store-based services (e.g. sports gyms, restaurants and concert venues) closed down temporarily in Japan. It should be noted that most of those who work in such service-based facilities tend to be part-time or provisional employees. Because of these closures, their lives may be under increased threat as a result of experiencing a great decline in their income or being made redundant.

In addition, low-income workers tend to be so-called ‘essential workers’ (e.g. workers in grocery stores, restaurants or factories), where remote work is not possible. It has been reported that Black Americans and Latinos with lower annual income are more likely to be such essential workers in the USA, and that several cases of workplace outbreaks of COVID-19 occurred as a result of inadequate personal protective equipment distribution.¹³ Another study found that the mortality rate of COVID-19 was higher among Black and Latino Americans.²⁶ Taken together, it can be speculated that essential workers with low income may be forced to work for a living with excessive fear of being infected with, or dying from COVID-19.

Approach toward the expansion of health inequality for the vulnerable

Before this pandemic, studies have indicated that income inequality causes depression through mechanisms on the individual level (e.g. feelings of withdrawal or shame), neighbourhood level (e.g. lower social capital) or national level (e.g. lack of investment in housing, education and public transportation, as well as pollution control, healthy food availability and accessibility of healthcare).²⁷ The present study also showed a clear association between income level and the proportion of SPD at phase 1 (11.27% and 4.15% among lowest and highest income category), implying that lower income citizens have suffered from income inequality, affecting their mental health, before this pandemic. In addition, this study showed that the onset of SPD was higher among low-income groups in phases 1–2 and phases 2–3, with both exhibiting linear tendency. Additionally, in phases 2–3, the group that earned an annual income of \geq £15 000 showed an SPD prevalence decrease of around 1–2% (from 6.9–11.5% to 5.9–9.1%), whereas the group that earned $<$ £15 000 showed an increase of about 1% (from 13.5% to 14.5%). These findings suggest that the health discrepancy between socioeconomic status may be widening in Japan during the COVID-19 pandemic. In the USA, the mortality rate of Black Americans is higher than for other ethnic groups, in relation to the country’s population ratio.²⁶ This rate is creating concerns regarding the widening of health discrepancies across the society as a result of the COVID-19 pandemic.²⁸ This study showed the gap in the proportion of SPD between lowest and highest income categories, which widened in just 3 months, implying that income inequality may be more severe in Japan. Lower income citizens may feel more withdrawn, lonely or shameful during the COVID-19 pandemic,^{25,29,30} which can lead to more unfavourable changes in mental health.

As previously described, the WHO has claimed that mental health measures should take ‘a whole of society approach’.⁷ Concurrently, perhaps an approach that provides more support to financially vulnerable people is necessary. For example, Japan, at first, proposed providing financial aid of ¥300 000 (approximately £2250) strictly to households facing financial difficulties, as an economic measure. However, this approach was criticised by

Japanese citizens for being unequal. The government, therefore, altered this initial aid, and executed a measure to provide ¥100 000 (approximately £750) to all citizens. Based on this study's results, a policy prioritising mental health measures for the financially vulnerable population, in which both resources and time are limited, may be effective from the viewpoint of better equalising health disparities during a pandemic.

Strengths and limitations

There are some limitations in our study that should be considered. First, since the participants were recruited from among those who had enrolled at a single online research company, the results may have been affected because of selection bias. Relatively little is known about the characteristics of people in online communities.³¹ Second, the study participants were recruited from the Tokyo metropolitan area, but not from all regions of Japan. In addition, the number of participants by age could not be set according to the proportion ratio in Japan. Thus, these results may not be directly applicable to the Japanese population. In addition, compared with middle-aged or older adults, younger adults were more likely to be lost to follow-up, which may cause selection bias. Third, no data on current or past history of medication for mental health were obtained for this study. Differential uptake of mental health interventions (e.g. medication or psychological therapies) could bias the results.

In conclusion, our study found that although the mental health of members of the Japanese public in the Tokyo metropolitan area did deteriorate as a result of the COVID-19 pandemic, the degree of deterioration was noted to be highest among those from lower income levels. Therefore, it may be that mental health countermeasures that focus specifically on low socioeconomic groups are necessary.

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Data availability

The data-set supporting these findings is not publicly available because of access restrictions imposed by the Tokyo Medical University Ethics Committee. Public data sharing is restricted to protect privacy and confidentiality.

Author contributions

M.M., I.N. and S.I. conceptualised the study. M.M., H.K. and S.I. were responsible for study methodology and data collection. H.K. conducted the formal analysis. H.K. and S.I. wrote the manuscript. M.M., I.N., R.S., Y.O., T.K. and H.W. reviewed and edited the manuscript.

Supplementary material

Supplementary material is available online at <https://doi.org/10.1192/bjo.2021.5>.

Declaration of interest

None.

References

- 1 World Health Organization (WHO). *Coronavirus Disease (COVID-2019) Situation Reports*. WHO, 2020 (<https://www.who.int/emergencies/diseases/novel-coronavirus-2019/situation-reports>).
- 2 Galea S, Merchant RM, Lurie N. The mental health consequences of COVID-19 and physical distancing: the need for prevention and early intervention. *JAMA Intern Med* 2020; **180**(6): 817–8.
- 3 Cowan K. *Survey Results: Understanding People's Concerns About the Mental Health Impacts of the COVID-19 Pandemic*. MQ: Transforming Mental Health and Academy of Medical Sciences, 2020 (<http://www.acmedsci.ac.uk/COVIDmentalhealthsurveys>).
- 4 Blustein DL, Duffy R, Ferreira JA, Cohen-Scali V, Cinamon RG, Allan BA. Unemployment in the time of COVID-19: a research agenda. *J Vocat Behav* 2020; **119**: 103436.
- 5 Bao Y, Sun Y, Meng S, Shi J, Lu L. 2019-nCoV epidemic: address mental health care to empower society. *Lancet* 2020; **395**: e37–8.
- 6 Gunnell D, Appleby L, Arensman E, Hawton K, John A, Kapur N, et al. Suicide risk and prevention during the COVID-19 pandemic. *Lancet Psychiatry* 2020; **7**: 468–71.
- 7 World Health Organization. *Policy Brief: COVID-19 and the Need for Action on Mental Health*. WHO, 2020 (https://www.un.org/sites/un2.un.org/files/un_policy_brief-covid_and_mental_health_final.pdf).
- 8 Dorn AV, Cooney RE, Sabin ML. COVID-19 exacerbating inequalities in the US. *Lancet* 2020; **395**: 1243–4.
- 9 Marmot M. Social determinants of health inequalities. *Lancet* 2005; **365**: 1099–104.
- 10 Nicola M, Alsaifi Z, Sohrabi C, Kerwan A, Al-Jabir A, Iosifidis C, et al. The socio-economic implications of the coronavirus and COVID-19 pandemic: a review. *Int J Surg* 2020; **78**: 185–93.
- 11 Yip PSF, Cheung YT, Chau PH, Law YW. The impact of epidemic outbreak: the case of severe acute respiratory syndrome (SARS) and suicide among older adults in Hong Kong. *Crisis* 2010; **31**: 86–92.
- 12 Wasserman IM. The impact of epidemic, war, prohibition and media on suicide: United States, 1910–1920. *Suicide Life-Threatening Behav* 1992; **22**: 240–54.
- 13 Lancet T. The plight of essential workers during the COVID-19 pandemic. *Lancet* 2020; **395**: 1587.
- 14 Rajkumar RP. COVID-19 and mental health: a review of the existing literature. *Asian J Psychiatr* 2020; **52**: 102066.
- 15 Wang C, Pan R, Wan X, Tan Y, Xu L, McIntyre RS, et al. A longitudinal study on the mental health of general population during the COVID-19 epidemic in China. *Brain Behav Immun* 2020; **87**: 40–8.
- 16 Hotopf M, Bullmore E, O'Connor RC, Holmes EA. The scope of mental health research in the COVID-19 pandemic and its aftermath. *Br J Psychiatry* 2020; **217**(4): 540–2.
- 17 Machida M, Nakamura I, Saito R, Nakaya T, Hanibuchi T, Takamiya T, et al. Adoption of personal protective measures by ordinary citizens during the COVID-19 outbreak in Japan. *Int J Infect Dis* 2020; **94**: 139–44.
- 18 Prime Minister of Japan and His Cabinet. *Declaration of a State of Emergency in Response to the Novel Coronavirus Disease (April 7)*. Cabinet Secretariat, 2020 (https://japan.kantei.go.jp/ongoingtopics/_00018.html).
- 19 Ministry of Health Labour and Welfare. *Basic Policies for Novel Coronavirus Disease Control by the Government of Japan*. Ministry of Health Labour and Welfare, 2020 (<https://www.mhlw.go.jp/content/10900000/000617686.pdf>).
- 20 Kessler RC, Barker PR, Colpe LJ, Epstein JF, Groerger JC, Hiripi E, et al. Screening for serious mental illness in the general population. *Arch Gen Psychiatry* 2003; **60**: 184–9.
- 21 Pratt LA. Serious psychological distress, as measured by the K6, and mortality. *Ann Epidemiol* 2009; **19**: 202–9.
- 22 Furukawa T, Kawakami N, Saitoh M, Ono Y, Nakane Y, Nakamura Y, et al. The performance of the Japanese version of the K6 and K10 in the World Mental Health Survey Japan. *Int J Methods Psychiatr Res* 2008; **17**: 152–8.
- 23 Hedeker D, Gibbons RD. A random-effects ordinal regression model for multi-level analysis. *Biometrics* 1994; **50**: 933–44.

- 24 Pierce M, Hope H, Ford T, Hatch S, Hotopf M, John A, et al. Mental health before and during the COVID-19 pandemic: a longitudinal probability sample survey of the UK population. *Lancet Psychiatry* 2020; **7**: 883–92.
- 25 Palgi Y, Shrira A, Ring L, Bodner E, Avidor S, Bergman Y, et al. The loneliness pandemic: loneliness and other concomitants of depression, anxiety and their comorbidity during the COVID-19 outbreak. *J Affect Disord* 2020; **275**: 109–11.
- 26 Yancy CW. COVID-19 and African Americans. *JAMA* 2020; **323**: 1891.
- 27 Patel V, Burns JK, Dhingra M, Tarver L, Kohrt BA, Lund C. Income inequality and depression: a systematic review and meta-analysis of the association and a scoping review of mechanisms. *World Psychiatry* 2018; **17**: 76–89.
- 28 Krouse HJ. COVID-19 and the widening gap in health inequity. *Otolaryngol Neck Surg* 2020; **163**(1): 65–6.
- 29 Grossman ES, Hoffman YSG, Palgi Y, Shrira A. COVID-19 related loneliness and sleep problems in older adults: worries and resilience as potential moderators. *Pers Individ Dif* 2021; **168**: 110371.
- 30 Shrira A, Hoffman Y, Bodner E, Palgi Y. COVID-19-related loneliness and psychiatric symptoms among older adults: the buffering role of subjective age. *Am J Geriatr Psychiatry* 2020; **28**(11): 1200–4.
- 31 Wright KB. Researching internet-based populations: advantages and disadvantages of online survey research, online questionnaire authoring software packages, and web survey services. *J Comput Commun* 2006; **10**: JCMC1034.

