

Dietary scores at midlife and healthy ageing in a French prospective cohort

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(Submitted 16 December 2015 – Final revision received 30 March 2016 – Accepted 27 April 2016 – First published online 15 June 2016)

Abstract

Although nutrition has been advocated as a major determinant of healthy ageing (HA), studies investigating the link between dietary quality and HA are scarce. We investigated the association between adherence to French food-based and nutrient-based guidelines at midlife, as assessed by three dietary scores, and HA. HA was assessed in 2007–2009, among 2329 participants of the Supplémentation en Vitamines et Minéraux Antioxydants study aged 45–60 years at baseline (1994–1995) and initially free of diabetes, CVD and cancer. HA was defined as not developing any major chronic disease, good physical and cognitive functioning, no limitations in instrumental activities of daily living, no depressive symptoms, no health-related limitations in social life, good overall self-perceived health and no function-limiting pain. Data from repeated 24-h dietary records provided at baseline permitted the computation of the modified French Programme National Nutrition Santé-Guideline Score (mPNNS-GS), the Probability of Adequate Nutrient Intake Dietary Score (PANDiet) and the Diet Quality Index-International (DQI-I). Associations of these scores with HA were assessed by logistic regression. In 2007–2009, 42% of men and 36% of women met our criteria of HA. After adjustment for potential confounders, higher scores of the mPNNS-GS (OR_{quartile 4 v. quartile 1} 1.44; 95% CI 1.10, 1.87; $P_{\text{trend}}=0.006$) and the PANDiet (1.28; 95% CI 1.00, 1.64; $P_{\text{trend}}=0.03$) were associated with higher odds of HA. We observed no association between DQI-I and HA. In conclusion, this study suggests a beneficial long-term role of high adherence to both food-based and nutrient-based French dietary guidelines for a HA process.

Key words: Nutritional recommendations: Dietary scores: Healthy ageing: Midlife exposures

Over the last few decades, new multidimensional concepts referred to as 'successful ageing' or 'healthy ageing' (HA) have emerged⁽¹⁾. These concepts aim to capture health during ageing as a whole, beyond specific medical conditions or body functions. A large part of these constructs is based on the model proposed by Rowe & Kahn⁽²⁾, which defines successful ageing as being at low risk of disease or disability, while maintaining high levels of cognitive and physical functioning, and an active engagement with life. Yet, a multitude of different models has been developed that differ in the choice of the included components⁽¹⁾ as well as in the indicators used to measure these components.

One of the modifiable environmental factors that have been advocated to have a decisive role for HA is diet⁽³⁾. In France, the official nutritional guidelines for the general public were developed in the context of the national public health nutrition programme (Programme National Nutrition Santé, PNNS)⁽⁴⁾. These guidelines include eight food-based items, in order to provide easily understandable public health messages. Moreover, the guidelines include a ninth item, which advises the population to have a regular physical activity, equivalent to at least 30 min of rapid walking per day. Nutrient-specific reference values such as the French *Apports Nutritionnels Conseillés* are less comprehensible for the general population, but equally important for

Abbreviations: DQI-I, Diet Quality Index-International; HA, healthy ageing; mPNNS-GS, modified Programme National Nutrition Santé-Guideline Score; PANDiet, Probability of Adequate Nutrient Intake Dietary Score; PNNS-GS, Programme National Nutrition Santé-Guideline Score; SU.VI.MAX, Supplémentation en Vitamines et Minéraux Antioxydants.

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prevention research. Although food-based recommendations can also account for non-nutrient components and food matrix effects, they have the disadvantage that the attainment of nutrient reference values is not guaranteed by construction. Thus, food-based and nutrient-based reference systems are complementary approaches whose comparison is of interest.

In epidemiological research, adherence to nutritional recommendations is evaluated by *a priori*-defined dietary scores⁽⁵⁾. Moreover, *a posteriori* methods are applied to characterise empirically derived overall dietary patterns. To the best of our knowledge, only four cohort studies^(6–9) have, to this day, evaluated measures of the overall diet with respect to multidimensional concepts of HA; three studies have identified positive and/or negative roles of specific empirically derived *a posteriori* dietary patterns^(6,7) and a positive role of adherence to the Mediterranean diet⁽⁸⁾. Moreover, although higher scores on the Alternative Healthy Eating Index-2010 (AHEI-2010) were positively related to HA in one study⁽⁸⁾, scores on the original AHEI (published in 2002) were unrelated to HA in another investigation⁽⁶⁾. A further study has reported that higher adherence to the Australian dietary guidelines was associated with a higher probability for HA⁽⁹⁾.

Hence, only one study has directly examined adherence to national nutrition recommendations (i.e. the Australian dietary guidelines), and no study has yet investigated the role of both food-based national dietary guidelines and nutrient reference values with respect to multidimensional concepts of HA.

We thus aimed to provide specific data on the pertinence of both food-based dietary guidelines and nutrient reference values in a French context, thus increasing the knowledge on the role of these important elements of public health nutrition strategies for a holistic prevention of age-related health decline. Notably, two different scores reflect adherence to French nutritional recommendations and reference values: the PNNS-Guideline Score (PNNS-GS), reflecting the official food-based guidelines⁽⁴⁾, and the Probability of Adequate Nutrient Intake Dietary Score (PANDiet), measuring adequacy to current nutrient reference values⁽¹⁰⁾. Moreover, an international index, the Diet Quality Index-International (DQI-I), comprising both nutrient- and food-group items, has been developed to facilitate between-country comparisons⁽¹¹⁾. This study examined the association between the above-mentioned dietary scores estimated at midlife and HA evaluated 13 years later in a large French cohort.

Methods

Study design

Initially, the 'Supplémentation en Vitamines et Minéraux Antioxydants' study (SU.VI.MAX, 1994–2002) was a French, randomised, double-blind, placebo-controlled, primary, prevention trial with a planned follow-up of 8 years⁽¹²⁾. In brief, after a national recruitment campaign with a call for volunteers living in France (women aged 35–60 years, men aged 45–60 years), 12 741 subjects returned a completed baseline questionnaire, met the eligibility criteria (no disease likely to hinder active participation or threatened 5-year survival; acceptance of participation constraints; no previous regular supplementation with

the tested antioxidants), were present at the inclusion visit, and included into the final study sample. The trial's objective was to investigate a potential effect of antioxidant supplementation at nutritional doses on the incidence of cancers, CVD and mortality^(12,13).

The SU.VI.MAX participants were invited, on a voluntary basis, to participate in an additional observational follow-up, the SU.VI.MAX 2 study (2007–2009), 5 years after the end of the trial. This follow-up study included 6860 subjects who completed clinical and neuropsychological examinations and a number of questionnaires⁽¹⁴⁾.

The SU.VI.MAX and SU.VI.MAX 2 studies were conducted according to the guidelines laid down in the Declaration of Helsinki, and all procedures involving human subjects were approved by the Ethics Committee for Studies with Human participants of Paris-Cochin Hospital (CCPPRB nos 706 and 2364, respectively) and the Commission National Informatique et Liberté (CNIL nos 334 641 and 907 094, respectively). Written informed consent was obtained from all subjects.

Baseline dietary data (1994–1996)

During the SU.VI.MAX study, all participants were invited to complete a 24-h dietary record every 2 months, using computerised questionnaires. An instruction manual including validated photographs of >250 foods was used to assist participants. Subjects could choose from seven possible portion sizes⁽¹⁵⁾. As previously stated⁽¹⁶⁾, dietary records were considered as invalid if energy intake was <418.4 kJ/d (<100 kcal/d) or >25 104 kJ/d (>6000 kcal/d). In addition, men reporting <3347 kJ/d (<800 kcal/d) and women reporting <2092 kJ/d (<500 kcal/d) across $\geq 1/3$ of records were excluded to account for energy under-reporting. Information on alcohol and seafood consumption was obtained by baseline questionnaires, as these food groups tend to be consumed less frequently than others. Alcohol consumption was estimated using a short, validated, semi-quantitative dietary questionnaire⁽¹⁷⁾. The computation of food and nutrient intakes was based on all eligible 24-h records collected during the first 2 years following inclusion. On average, 10.2 dietary records (interquartile range: 8–13) were available per participant. The construction of the different dietary scores is presented in Fig. 1 and detailed in the online Supplementary Material S1.

Other baseline (1994–1996) variables

Data on sex, date of birth, education (primary, secondary, university level), occupational category (homemakers, manual workers, intermediate professions, managerial staff/intellectual profession), living arrangement (living alone, living in a couple), smoking status (never smoked, former or current smoker), physical activity (irregular, <1 h of walking/d, ≥ 1 h of walking/d) and subjective memory complaints (yes/no) were collected using self-administered questionnaires. BMI (kg/m²) was calculated using anthropometric measurements performed by trained personnel. Systolic and diastolic blood pressures were measured three times using a standard mercury sphygmomanometer after lying down for 10 min, and the mean values were calculated.



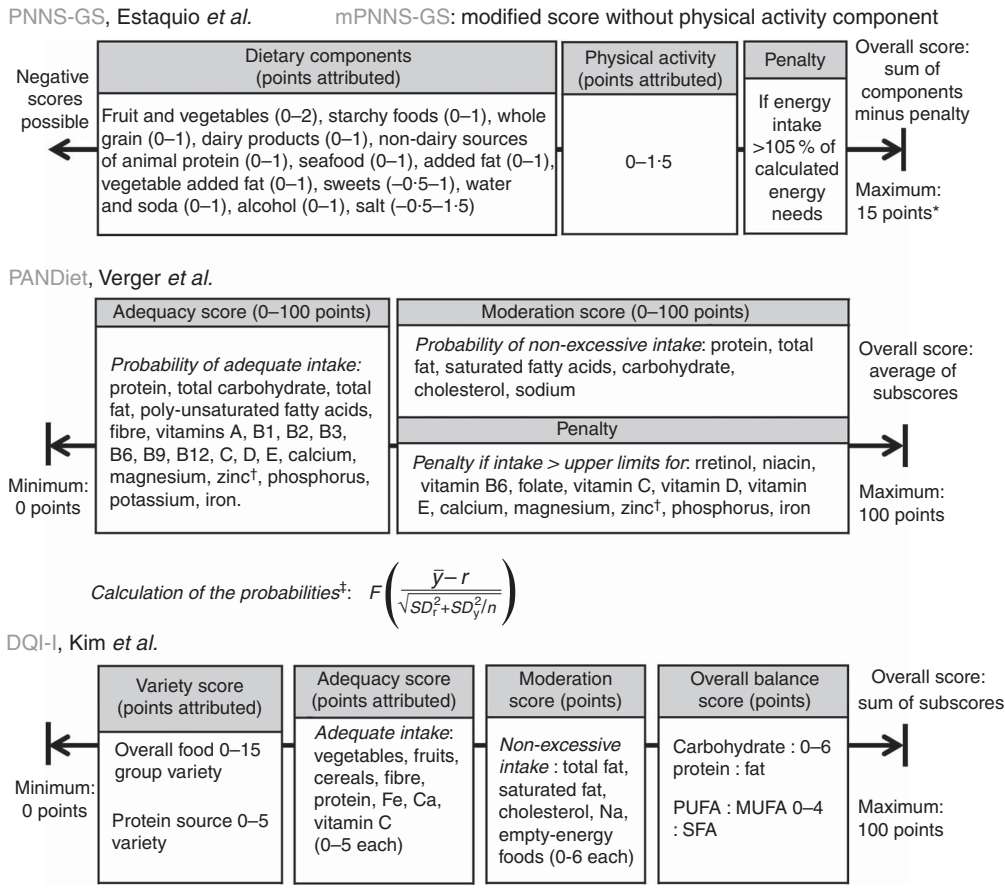


Fig. 1. Characteristics and computation of the investigated dietary scores. * Maximum of the modified Programme National Nutrition Santé-Guideline Score (mPNNS-GS): 13.5 points. † We did not have any information on Zn consumption, and thus did not include this variable in the calculation of the score. ‡ Probability calculation formula (PANDiet): *F*, function 'probnorm' in SAS software package; \bar{y} , mean intake; *r*, nutrient reference value; SD_t^2 , interindividual variability; SD_d^2 , day-to-day variability of intake; *n*, number of dietary record days. Probability values range from 0 to 1, and are multiplied by 100 in order to obtain subscores of 0–100 points. PNNS-GS, Programme National Nutrition Santé-Guideline Score; PANDiet, Probability of Adequate Nutrient Intake Dietary Score; DQI-I, Dietary Quality Index-International.

Data on treatment and storage (at –80°C) of fasting blood samples as well as the measurement of serum concentrations of vitamin C, β-carotene, α-tocopherol, Zn, Se⁽¹⁸⁾ and glucose⁽¹²⁾ have been described previously⁽¹²⁾.

Events of cancer and CVD during follow-up (1994–2009)

Data collection concerning events of cancer and CVD during follow-up has been extensively described^(12,19). In brief, an independent expert committee validated such events after review of relevant medical records, relying on the 10th International World Health Organization Classification of Diseases⁽²⁰⁾.

Definition of 'healthy ageing'

This study's definition of HA was largely based on the concept proposed by Rowe & Kahn⁽²⁾. Hence, HA was defined as follows: (a) the absence of incident major chronic disease (cancer, CVD or diabetes) during follow-up, limitations in instrumental activities of daily living, function-limiting pain, depressive symptomatology and health-related limitations in social life; and (b) the presence of good physical and cognitive functioning and

good overall self-perceived health (see Table 1). We defined HA as a binary variable (meeting all of the above criteria or not). An extensive description of the construction of our HA concept has been published previously⁽³⁰⁾.

Study sample selection

We selected those participants of the SU.VI.MAX study aged 45–60 years at inclusion into the SU.VI.MAX study (*n* 9867), who were free of diabetes, CVD or cancer at inclusion (*n* 9180), and had available data for computation of the dietary scores (*n* 4434). After exclusion of subjects with incomplete information on HA status, a final study sample of 2329 individuals (1246 men and 1083 women) was obtained.

The explanation for the large number of individuals with missing data on HA status is that HA-related information was only collected among those SU.VI.MAX participants who decided to accept the invitation to complete the SU.VI.MAX 2 follow-up point, which was completely voluntary. Death was only a minor reason for non-participation (of the above-mentioned 4434 individuals who corresponded to our inclusion criteria concerning age, prevalent diseases and dietary data, only fifty had died during follow-up).

Table 1. Criteria used to define 'healthy ageing', SUPplémentation en Vitamines et Minéraux AntioXydants (SU.VI.MAX) and SU.VI.MAX 2 studies, France, 1994–2009*

Criteria†	Definition	Additional information on the instrument used
Good physical functioning	SPPB $\geq 11/12$	Physical test battery administered by trained physicians (including repeated chair stands, balance testing, and gait speed testing)
Good cognitive functioning	MMSE $\geq 27/30$ and RI-48 $\geq 19/48$ and DK-TMT ≥ 5.5	Cognitive test battery administered by trained physicians (evaluating overall cognitive functioning, verbal episodic memory and executive function, respectively)
No limitations in IADL	<1 limitation	Self-administered questionnaire (including, among others, questions on the ability to travel, go shopping and do housekeeping)
No depressive symptoms	CES-D <16/60	Self-administered questionnaire developed for the evaluation of depressive symptomatology in the general population, in the context of epidemiological studies
No health-related limitations in social life	SF-36 responses: 1–2 to item 6 and 3–5 to item 10	SF-36: very widely used, self-administered questionnaire designed to measure vitality, physical functioning, bodily pain, general health perceptions, physical role functioning, emotional role functioning, social role functioning and mental health
Good overall self-perceived health	SF-36 responses: 1–3 to item 1	SF-36: see above explanations
No function-limiting pain	SF-36 responses: 1–3 to item 7 and 1–2 to item 8	SF-36: see above explanations
No incident major chronic disease	No incident cancer (i.e. cancer of any kind, except for basal cell carcinoma), or CVD‡ during follow-up No incident diabetes during follow-up	Validation of events by independent expert committee No fasting blood glucose value ≥ 1.26 g/l, anti-diabetic medication use or self-reported diabetes at the end of follow-up

SPPB, Short Physical Performance Battery⁽²¹⁾; MMSE, Mini Mental State Examination^(22,23); RI-48, rappel indicé-48 items⁽²⁴⁾; DK-TMT, Delis–Kaplan version of the Trail-making test⁽²⁵⁾; IADL, instrumental activities of daily living^(26,27); CES-D, Center for Epidemiologic Studies Depression Scale⁽²⁸⁾; SF-36, Medical Outcome Study Short Form-36⁽²⁹⁾.

* This table is based on another article in which the SU.VI.MAX definition of 'healthy ageing' was originally published⁽³⁰⁾.

† All criteria were assessed at follow-up (2007–2009), except for events of major chronic disease, which were assessed over the follow-up (1994–2009). The test batteries were administered in visit centres within hospitals near the participants' homes, and the questionnaires were completed by the participants at home, and then verified by technicians.

All subjects were free of major chronic disease at inclusion.

‡ CVD was defined as codes I20–I25, I63, I65, I66, I70, I71 and I74 from the 10th International World Health Organization Classification of Diseases.

Descriptive statistics

Participant characteristics were compared across quartiles of the dietary scores, using linear contrast tests and Cochran–Mantel–Haenszel tests. In addition, Mann–Whitney *U* and χ^2 tests were used to compare participants included in our analyses with those excluded because of missing data on HA status, considering the above-mentioned sample of 4434 participants with complete dietary data as the 'source population'. Finally, we investigated the interrelations between dietary scores (using Spearman's correlations) and between components of our HA definition (using κ coefficients).

Main statistical analyses

Logistic regression was performed to estimate the association between the modified Programme National Nutrition Santé-Guideline Score (mPNNS-GS), the PANDiet and the DQI-I, modelled as quartiles, and subsequent HA. Tests for (log-)linear trend were performed by modelling the quartiles of dietary scores as ordinal variables. Model 1 was adjusted for baseline (1994–1996) age and sex, and model 2a was adjusted for age, sex, follow-up time, supplementation group, occupation, living arrangement, smoking status, educational level, follow-up time, energy intake, number of 24-h records, physical activity (except for the PNNS-GS) and alcohol intake (except for the PNNS-GS and mPNNS-GS).

As BMI may be a potential mediator of the relationship between diet and HA, we created a supplemental model, model 2b, which was further adjusted for baseline BMI. As our

principal objective was to examine the overall association between diet and HA (thus including the part of this association that may be mediated by anthropometric indicators), we considered model 2a as our main model. Finally, we conducted a sensitivity analysis using versions of the mPNNS-GS from which we took out, one by one, the different components of the score. For example, one modified score was the same score as the mPNNS-GS except that the component on salt consumption was taken out, and another corresponded to the mPNNS-GS without the component on fruit and vegetable consumption. These 'reduced' scores were modelled as standardised continuous variables, and we adjusted each analysis for the component that was taken out from the score.

As there was no interaction between sex and any of the scores on HA (all $P > 0.1$), we analysed data from men and women simultaneously.

To partly correct for the selection bias related to exclusion of participants with missing data on HA status, all analyses were carried out using inverse probability weighting⁽³¹⁾. The probability of inclusion into our analysis for each individual of the 'source population' ($n = 4434$, cf. paragraph 'Descriptive statistics') was calculated using logistic regression (as a function of baseline variables). The inverse of the probability to be included (multiplied by the sampling proportion $n_{\text{included}}/n_{\text{total}}$) was then used as weights for our analyses.

Additional statistical analyses

To investigate the cumulative effect of dietary quality and physical activity, we also investigated the association between



unmodified PNNs-GS score and HA. Furthermore, in addition to models in which the scores were modelled as quartiles, we created models including the dietary scores as standardised continuous variables. In addition, as 'age' is a key notion of our study's outcome, we conducted analyses stratified by age at follow-up (<65 *v.* ≥65 years).

All analyses were conducted with SAS (version 9.3; SAS Institute, Inc.). Missing values for covariables (*n* 50 for smoking status, *n* 33 for living arrangement and *n* 26 for occupational category) were dealt with by multiple imputation. Details on how exactly multiple imputation was carried out are given in the online Supplementary Material S2.

Results

The present analysis included 2329 participants. Our study sample's mean age at follow-up was 65.3 (SD 4.5) years, and the mean follow-up time was 13.5 (SD 0.4) years. Our criteria of HA

were met by 42% of men and 36% of women. The distribution of age categories at the SU.VI.MAX 2 follow-up visit according to HA status was as follows: HA=no: 11% <60 years, 37% 60–64 years, 30% 65–69 years, 22% ≥70 years; HA=yes: 12% <60 years, 45% 60–64 years, 26% 65–69 years, 18% ≥70 years.

A total of 2329 participants were included in the present analysis. Compared with them, participants who were excluded because of missing data on HA status (*n* 2105) were younger, less educated, more often smokers, consumed less fruits and vegetables, had lower scores on the PANDiet and the DQI-I, a higher mean BMI and a higher mean fasting blood glucose value (online Supplementary Table S3).

Tables 2 and 3 and online Supplementary Tables S4–S7 present participant characteristics according to quartiles of the dietary scores. Most of the investigated variables (socio-demographic, lifestyle, health and nutritional data) differed according to quartiles of the different dietary scores. For

Table 2. Baseline general participant characteristics according to quartiles (Q) of the modified Programme National Nutrition Santé-Guideline Score (Numbers and percentages; medians and 1st, 3rd quartiles)

Baseline characteristics	<i>n</i>	Q1*		Q2*		Q3*		Q4*		<i>P</i> †
		<i>n</i>	%	<i>n</i>	%	<i>n</i>	%	<i>n</i>	%	
Age (years)	2329									0.009
Median		64.43		64.82		64.75		65.32		
1st, 3rd quartiles		61.33, 68.29		61.63, 69.08		61.75, 69.17		61.79, 69.69		
Sex	2329									<0.001
Men		402	66.6	306	57.2	302	47.9	236	42.2	
Women		202	33.4	229	42.8	329	52.1	323	57.8	
Educational level (%)	2329									0.005
Primary education only		143	23.7	119	22.2	135	21.4	93	16.6	
Secondary education		247	40.9	210	39.3	231	36.6	243	43.5	
University level		214	35.4	206	38.5	265	42.0	223	39.9	
Occupational status (%)	2303									<0.001
Homemaker		23	3.9	41	7.8	45	7.2	64	11.6	
Manual worker		46	7.7	28	5.3	42	6.7	18	3.3	
Employees		335	56.1	285	53.9	336	53.8	304	55.0	
Managerial staff‡		193	32.3	175	33.1	201	32.2	167	30.2	
Living arrangement (%)	2296									0.03
Living alone		63	10.6	67	12.7	82	13.1	83	15.0	
Living in a couple		529	89.4	460	87.3	542	86.9	470	85.0	
Smoking habits (%)	2279									<0.001
Never smoker		259	43.6	264	51.0	334	54.0	298	54.3	
Former smoker		246	41.4	200	38.6	231	37.4	211	38.4	
Current smoker		89	15.0	54	10.4	53	8.6	40	7.3	
Physical activity level (%)	2329									<0.001
Irregular or none		369	61.1	301	56.3	327	51.8	272	48.7	
<1 h/d		122	20.2	121	22.6	161	25.5	134	24.0	
≥1 h/d		113	18.7	113	21.1	143	22.7	153	27.4	
BMI (kg/m ²)	2329									0.04
Median		24.16		24.11		23.84		23.74		
1st, 3rd quartiles		21.97, 26.48		22.22, 26.20		21.91, 25.88		21.93, 26.04		
Systolic blood pressure (mmHg)	1996									0.04
Median		125.0		120.0		125.0		120.0		
1st, 3rd quartiles		120.0, 135.0		115.0, 130.0		115.0, 130.0		115.0, 135.0		
Diastolic blood pressure (mmHg)	1996									0.003
Median		80.0		80.0		80.0		80.0		
1st, 3rd quartiles		75.0, 90.0		70.0, 85.0		75.0, 85.0		70.0, 85.0		
Fasting blood glucose (g/l)	2318									<0.001
Median		5.72		5.66		5.61		5.61		
1st, 3rd quartiles		5.33, 6.11		5.33, 6.05		5.27, 6.00		5.22, 5.88		

* Quartile cut-off values (maximal values per quartile): Q1, 6.05; Q2, 7.25; Q3, 8.50.

† Linear contrast tests (continuous variables) or Cochran–Mantel–Haenszel tests (categorical variables). Occupational status: χ^2 -test, as the variable values were not ordinal.

‡ Or intellectual profession.

Table 3. Baseline nutrition-related participant characteristics according to quartiles (Q) of the modified Programme National Nutrition Santé-Guideline Score (mPNNS-GS) (Medians and 1st, 3rd quartiles)

Baseline characteristics	n	Q1*		Q2*		Q3*		Q4*		P†
		Median	1st, 3rd quartiles	Median	1st, 3rd quartiles	Median	1st, 3rd quartiles	Median	1st, 3rd quartiles	
Dietary scores										
mPNNS-GS (points)	2329	5.25	4.58, 5.75	6.75	6.50, 7.00	7.80	7.55, 8.08	9.30	8.80, 9.80	<0.001
PNNS-GS (points)	2329	5.74	4.79, 6.25	7.08	6.73, 7.80	8.30	7.80, 9.05	10.00	9.30, 10.75	<0.001
PANDiet (points)	2329	60.92	57.70, 65.24	61.92	58.49, 65.84	63.53	59.59, 67.40	67.04	62.03, 71.39	<0.001
DQI-I (points)	2329	52.24	48.44, 56.49	54.71	51.07, 58.37	56.55	53.05, 59.58	60.30	56.67, 64.01	<0.001
Alcohol consumption (g/d)	2329	20.99	5.91, 36.07	18.10	0.00, 28.27	6.43	0.00, 20.99	5.91	0.00, 20.99	<0.001
Fruit/vegetable intake (g/d)	2329	337.0	260.5, 445.0	368.2	274.8, 491.3	408.4	322.6, 509.5	473.5	361.3, 599.7	<0.001
Total energy intake (kJ/d)	9744	10 770	9079, 12 410	9401	7452, 10 891	8598	7109, 10 134	7912	6652, 9376	<0.001
Total energy intake (kcal/d)	2329	2574	2170, 2966	2247	1781, 2603	2055	1699, 2422	1891	1590, 2241	
%Total fat	2329	37.69	34.35, 40.96	37.69	34.93, 40.75	37.85	34.83, 41.07	37.11	33.43, 40.41	0.15
%SFA	2329	16.00	14.31, 17.62	15.63	14.14, 17.44	15.48	13.85, 17.10	14.67	12.76, 16.41	<0.001
%MUFA	2329	13.98	12.52, 15.32	14.08	12.92, 15.59	14.29	12.93, 15.82	14.17	12.55, 15.57	0.21
%PUFA	2329	5.06	4.43, 5.89	5.31	4.56, 6.40	5.64	4.87, 6.69	5.86	5.02, 6.74	<0.001
%Carbohydrates	2329	38.86	33.94, 43.64	39.27	34.73, 43.15	39.95	36.15, 44.08	40.91	37.09, 45.16	<0.001
%Added sugars	2329	7.77	5.46, 10.68	7.34	5.29, 9.60	7.22	5.24, 9.37	6.87	4.98, 8.54	<0.001
%Protein	2329	15.87	14.55, 17.46	16.08	14.83, 17.72	16.39	15.00, 18.01	16.98	15.44, 18.87	<0.001
%Animal protein	2329	11.57	9.95, 13.34	11.81	10.29, 13.52	11.83	10.32, 13.72	12.38	10.48, 14.30	<0.001
%Plant protein	2329	4.27	3.80, 4.80	4.32	3.88, 4.88	4.51	4.03, 5.10	4.67	4.17, 5.24	<0.001
Dietary fibre intake (g/4184 kJ (1000 kcal))	2329	7.77	6.79, 8.94	8.44	7.31, 9.88	9.32	8.16, 10.71	10.67	9.22, 12.63	<0.001
Na intake (mg/d)	2329	4.07	3.40, 4.86	3.59	2.88, 4.40	3.40	2.64, 4.14	2.97	2.42, 3.58	<0.001
Vitamin C (µmol/l)‡	1804	51.73	36.91, 63.54	54.48	41.24, 67.05	55.17	44.18, 67.52	59.94	49.99, 69.54	<0.001
Vitamin E (µmol/l)‡	1991	30.80	26.10, 35.30	31.20	26.80, 36.14	31.00	26.46, 36.63	31.73	26.59, 36.40	0.10
β-Carotene (µmol/l)‡	1991	0.44	0.28, 0.67	0.48	0.32, 0.76	0.52	0.34, 0.83	0.59	0.39, 0.89	<0.001
Se (µmol/l)‡	2260	1.10	0.99, 1.22	1.10	0.98, 1.24	1.10	0.99, 1.22	1.10	0.99, 1.23	0.54
Zn (µmol/l)‡	2264	13.10	11.90, 14.30	13.35	12.00, 14.60	13.10	11.90, 14.20	13.20	12.00, 14.40	0.83

Dietary scores and healthy ageing

PNNS-GS, Programme National Nutrition Santé-Guideline Score; PANDiet, Probability of Adequate Nutrient Intake Dietary Score; DQI-I, Dietary Quality Index-International.

* Quartile cut-offs (maximal values per quartile): Q1, 6.05; Q2, 7.25; Q3, 8.50.

† Linear contrast tests.

‡ Blood serum concentrations.

Table 4. Association between quartiles (Q) of dietary scores and healthy ageing (Odds ratios and 95% confidence intervals; *n* 2329)*

	Q1		Q2		Q3		Q4		<i>P</i> _{for trend} †
	OR	95% CI	OR	95% CI	OR	95% CI	OR	95% CI	
Principal analyses									
mPNNS-GS									
Model 1*	1	–	1.18	0.92, 1.50	1.36	1.08, 1.72	1.60	1.25, 2.04	<0.001
Model 2a†	1	–	1.14	0.88, 1.46	1.26	0.98, 1.61	1.44	1.10, 1.87	0.006
Model 2b‡	1	–	1.15	0.89, 1.48	1.27	0.99, 1.63	1.45	1.11, 1.89	0.005
PANDiet									
Model 1*	1	–	1.02	0.81, 1.30	1.15	0.91, 1.46	1.35	1.07, 1.72	0.008
Model 2a†	1	–	1.01	0.79, 1.28	1.15	0.90, 1.47	1.28	1.00, 1.64	0.03
Model 2b‡	1	–	0.99	0.77, 1.26	1.10	0.86, 1.41	1.22	0.95, 1.56	0.08
DQI-I									
Model 1*	1	–	1.36	1.07, 1.73	1.50	1.18, 1.91	1.39	1.09, 1.77	0.005
Model 2a†	1	–	1.32	1.03, 1.68	1.38	1.08, 1.77	1.28	0.99, 1.64	0.05
Model 2b‡	1	–	1.32	1.03, 1.68	1.39	1.08, 1.77	1.25	0.97, 1.61	0.07
Supplemental analyses									
PNNS-GS									
Model 1*	1	–	1.24	0.98, 1.57	1.41	1.11, 1.80	1.70	1.33, 2.18	<0.001
Model 2a†	1	–	1.21	0.95, 1.55	1.40	1.08, 1.80	1.64	1.26, 2.13	<0.001
Model 2b‡	1	–	1.20	0.94, 1.54	1.38	1.07, 1.79	1.61	1.24, 2.10	<0.001

mPNNS-GS, modified Programme National Nutrition Santé-Guideline Score; PANDiet, Probability of Adequate Nutrient Intake Dietary Score; DQI-I, Dietary Quality Index-International; PNNS-GS, Programme National Nutrition Santé-Guideline Score.

* Model 1 is adjusted for age and sex.

† Model 2a is adjusted for age, sex, supplementation group, occupation, living arrangement, smoking status, educational level, follow-up time, energy intake, number of 24-h records, physical activity (except for the PNNS-GS) and alcohol intake (except for the PNNS-GS and the mPNNS-GS).

‡ Model 2b is adjusted for the same variables as model 2a, and additionally for baseline BMI. We consider model 2a as our main model.

instance, higher scores on the mPNNS-GS were associated with higher age, a higher probability to be female, to have a higher educational level and occupational status, a healthier lifestyle, more favourable health parameters, higher serum vitamin C and β -carotene and a nutritional profile that was closer to the French official nutrition guidelines. Spearman's coefficients for the correlation of the mPNNS-GS with the PANDiet and the DQI-I were 0.33 and 0.45, respectively (data not shown). κ Coefficients illustrating the interrelations between the different components of our HA definition are displayed in the online Supplementary Table S8.

Table 4 presents the association between quartiles of dietary scores and HA, providing OR with 95% CI and *P* values for linear trend. In the fully adjusted model, higher scores on the mPNNS-GS (OR_{quartile 4 v quartile 1} 1.44; 95% CI 1.10, 1.87; *P*_{trend} = 0.006) and the PANDiet (1.28; 95% CI 1.00, 1.64; *P*_{trend} = 0.03), but not on the DQI-I, were related to a higher probability of HA. In fact, the highest odds for HA were not observed for quartile 4 of the DQI-I but for quartile 3 of the DQI-I – indicating the absence of a (log-)linear relationship. Our supplemental analysis concerning the PNNS-GS showed that high values on this index were linked to an even higher increase in odds of HA than high scores on the modified score mPNNS-GS.

Our models in which we additionally included baseline BMI revealed that after adjusting for BMI essentially the same results were observed for the mPNNS-GS, the observed association was slightly weaker for the DQI-I and the PNNS-GS and no longer statistically significant for the PANDiet.

Fig. 2 shows the association between standardised dietary scores (mean = 0, sd = 1) and HA. Positive associations between the scores and HA were only observed for the mPNNS-GS and

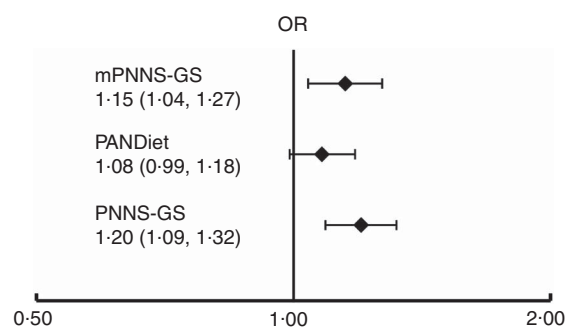


Fig. 2. Association between continuous standardised dietary scores and healthy ageing (*n* 2329). OR were calculated via a multivariable logistic regression model, adjusted for age, sex, supplementation group, occupation, living arrangement, smoking status, educational level, follow-up time, energy intake, number of 24-h records, physical activity (except for the Programme National Nutrition Santé-Guideline Score (PNNS-GS)) and alcohol intake (except for the PNNS-GS and the modified Programme National Nutrition Santé-Guideline Score (mPNNS-GS)). PANDiet, Probability of Adequate Nutrient Intake Dietary Score.

the PNNS-GS, but not for the PANDiet. We do not present the association between continuous DQI-I and HA as the analysis by quartiles had indicated the absence of a (log-)linear relationship.

Online Supplementary Table S9 presents analyses stratified by age at follow-up. A positive relationship between the mPNNS-GS and the probability for HA was present among both younger and older participants. Significant results for the other dietary scores were not observed in either age group. Online Supplementary Table S10 presents our analysis in which we took out the different components of the mPNNS-GS. For some of the 'reduced' scores, a slightly weakened (scores without

seafood, without salt or without the component on meat, poultry, seafood and eggs) or a slightly strengthened (score without the component on vegetable added fats) association was observed. However, all observed OR remained quite similar to the OR observed for the mPNNS-GS.

Discussion

Summary of findings

In this large cohort of French adults, higher adherence to the French official nutrition guidelines, as measured by the mPNNS-GS, and higher adequacy regarding French nutrient-based recommendations, as measured by the PANDiet, were prospectively related to higher odds of HA.

The magnitude of these associations was stronger for the mPNNS-GS than for the PANDiet, which was only related to HA when modelled as quartiles, but not when modelled as a continuous variable. The relationship between DQI-I and HA was not linear – as the highest odds of HA were found for the third quartile. We did not attempt to model a non-linear relationship between DQI-I and HA, as nutritional indices should by definition yield the most favourable public health outcomes for the highest scores. One potential explanation for the shape of this association may be that the highest scores for the ‘total fat’ item of the ‘moderation’ subscore of the DQI-I are attributed to subjects covering <20% of their total energy intakes by fat. Recommendations including such drastically low fat intakes have been controversially discussed in the literature^(32,33). A potential reason for the generally better performance of the mPNNS-GS in terms of predicting HA, with respect to the other tested scores, may be that only the mPNNS-GS penalises energy over-consumption. However, using an unpenalised version of the mPNNS-GS yielded very similar results as our main analyses (data not shown). The relatively low correlations between the mPNNS-GS and the PANDiet and the DQI-I indicate that the overall dietary quality is reflected in a different manner by this score than by the other scores. The main difference between the scores is that the mPNNS-GS reflects food-based dietary guidelines, whereas the PANDiet reflects nutrient-based reference recommendations, and the DQI-I reflects both food-based and nutrient-based recommendations. Moreover, only the mPNNS-GS includes a component on alcohol consumption – however, our sensitivity analyses showed that a version of the mPNNS-GS in which the alcohol component was excluded yielded essentially the same results as the complete mPNNS-GS score. This was also true for the other components of the mPNNS-GS score, which did not, in an isolated manner, appear to drive the observed association of mPNNS-GS with HA. Further investigations are needed to determine to which degree adherence to food-based dietary guidelines genuinely shows a stronger association with HA than the adequacy to nutrient reference values, and to which degree such findings are influenced by technical aspects such as different types of scoring systems.

Our supplemental results concerning the association between PNNS-GS and HA suggest that the highest health benefits can be obtained by a combination of adherence to nutritional recommendations and a high physical activity level.

Moreover, in our age-stratified supplemental analyses, a positive relationship between mPNNS-GS and HA was present among both younger and older participants. The fact that we observed no significant results for the PANDiet in age-stratified analyses is probably related to limited statistical power.

Our analyses in which we additionally adjusted for baseline BMI indicate that, although the observed associations for the mPNNS-GS, the DQI-I and the PNNS-GS were probably not strongly mediated by this anthropometric indicator, there may have been a substantial mediation effect for the PANDiet. However, these results should be cautiously interpreted as we did not carry out a formal mediation analysis.

Comparison with the literature

To the best of our knowledge, only three studies have investigated the association between *a priori* dietary scores and HA. In a cross-sectional study based on the Nurses’ Health Study⁽⁸⁾, higher scores on the food- and nutrient-based AHEI-2010⁽³⁴⁾ and on the Alternate Mediterranean diet score were associated with a higher probability of HA. In a further prospective study, there was a positive association of the adherence to the Australian food-based dietary guidelines with odds for HA⁽⁹⁾. In contrast, in a prospective study based on the Whitehall II cohort⁽⁶⁾, lower adherence to the original AHEI was unrelated to HA.

In the study published by Akbaraly *et al.*⁽⁶⁾, the relationship between *a posteriori* dietary patterns and HA was also investigated, and an inverse association between a *Western-type* dietary pattern and the probability of HA was observed. This is in line with a study based on the Melbourne Collaborative Cohort⁽⁷⁾, which observed an inverse association between a pattern characterised by high consumption of meat and fatty foods and HA, in addition to a positive association between a high fruit dietary pattern and HA.

In a previous investigation of data from the SU.VI.MAX study⁽³⁵⁾, no relationship between a *Western-type* pattern and HA was observed. On the other hand, there was a positive prospective association between a *healthy-type* dietary pattern and HA among participants with below-median energy intakes, underlying the importance of both high dietary quality and regulated energy intake.

One aspect that should be taken into consideration when comparing the present study with the above-mentioned other studies are substantial differences in the applied HA concepts. The criteria applied by Akbaraly *et al.* (surviving to ≥ 60 years, absence of chronic conditions and mental health problems, presenting above sex- and age-specific median performance in cardiometabolic, respiratory, musculoskeletal and cognitive functioning tests) and Samieri *et al.* (survival to ≥ 70 years, absence of major chronic disease and of major impairments in cognitive, physical and mental functioning) appear as particularly severe, as only 4 and 11 % of participants were identified as ‘ideal’ or ‘healthy’ agers, respectively. Hodge *et al.* (HA definition: surviving to at least age 70 years, good mental health, absence of major chronic disease and of major limitations in physical functioning) and Gopinath *et al.* (HA definition: absence of disability, depressive symptoms, cognitive impairment, respiratory symptoms and chronic diseases) identified 19

and 16% of participants as 'successful' agers, respectively. In our study, even more individuals (39%) were classed as 'healthy' agers. This is probably related to the fact that, in contrast to the above-cited studies, survival to a specific age was not part of our HA definition. However, differences in HA proportions across studies are not merely a function of age but also of the chosen indicators and cut-off values, as is illustrated by the fact that the mean age at HA assessment was about 67 years in the study by Akbaraly *et al.* (showing very low HA proportions) and about 76 years in the study by Hodge *et al.* (showing markedly higher HA proportions).

In contrast to the limited literature on the link between overall diet and HA, the number of publications on the association between holistic representations of the diet and specific diseases has much increased during the past few decades^(3,5,36–39). A review article concluded that observational studies suggested a modest favourable role of 'healthful' dietary patterns concerning all-cause mortality and CHD – but did not consistently suggest a role concerning cancer⁽⁵⁾. Moreover, a recent publication issued from a conference on 'Nutrition and healthy ageing' concluded that the available studies globally suggested an important role of 'dietary patterns rich in fruit, vegetables, fish, whole grains and starchy low-fat staple foods' for HA-related outcomes such as life expectancy, cardiometabolic health and cognitive health⁽³⁾.

Overall, scientific literature on the link between overall diet and health status including physical and cognitive functioning among elderly persons is scant, and does not yet permit an estimation of the quantity of a potential protective effect of a high-quality overall diet.

Limitations and strengths

Some limitations of our study should be noted. First, published definitions of HA are quite heterogeneous⁽¹⁾. However, the HA model used in the present study was largely based on the 'classical' concept proposed by Rowe & Kahn⁽²⁾. Second, in our study, HA status was not available at baseline (1994–1996). Yet, subjects in our study were only aged 45–60 years at inclusion into the SU.VI.MAX trial and free of major chronic disease. Accordingly, our working hypothesis that individuals were initially healthy is quite plausible. Third, the external validity of our results may be limited as the SU.VI.MAX participants are a sample of selected volunteers. Notably, individuals with a particularly poor overall diet may be under-represented in a nutrition-related study. Although potential selection bias was addressed by inverse probability weighting, this may have been insufficient to fully counterbalance the observed differences between excluded and included participants concerning important lifestyle and health characteristics. Finally, given our observational study design, we cannot exclude potential residual confounding. Important strengths of our analysis are the prospective study design, the use of a large set of HA indicators and the availability of accurate nutritional data, given the availability of, on average, ten dietary records per subject.

In conclusion, this study suggests a beneficial role of high adherence to the food-based and nutrient-based French nutritional recommendations for a HA process, including the

avoidance of chronic disease, good physical and cognitive functioning, as well as good self-perceived, mental and social health. The indicator that was most strongly related to a higher probability of HA in our study was the PNNS-GS, which reflects both the food-based items and the physical activity item of the French official nutrition guidelines. This suggests a high pertinence of the French official nutrition guidelines for the prevention of age-related health decline, and more generally a high importance of both high dietary quality and adequate physical activity level for health ageing. Further prospective observational and intervention studies are needed to confirm these findings.

Acknowledgements

The authors thank Younes Esseddik, Gwenaël Monot, Paul Flanzky, Mohand Ait Oufella, Yasmina Chelghoum and Than Duong Van (computer scientists), Rachida Mehroug (logistic assistant) and Nathalie Arnault, Véronique Gourlet, Fabien Szabo, Laurent Bourhis and Stephen Besseau (statisticians) for their technical contribution to the SU.VI.MAX study. The authors also thank Stéphane Raffard, who was responsible for standardisation of the cognitive evaluation, and Frédérique Ferrat, who coordinated the logistic aspects of the neuro-psychological evaluation.

This study was funded by the French National Research Agency (no. ANR-05-PNRA-010), the French Ministry of Health, Médéric, Sodexo, Ipsen, Mutuelle générale de l'Éducation nationale (MGEN) and Pierre Fabre. Mederic and MGEN are French health insurance organisations, which are complementary to the National Health Insurance System. Ipsen and Pierre Fabre are private pharmaceutical companies. Sodexo, a food catering company, supported the study by organising meetings between researchers and study participants. The funding bodies did not have any involvement in the design/conduct of the research, in data analysis/interpretation or in writing/approval of the manuscript. K. E. A. was supported by a doctoral fellowship from the Ecole Doctorale Galilée, University of Paris 13, Sorbonne Paris Cité. None of the funders had any role in study design, data collection and analysis, decision to publish or writing of the manuscript.

S. H., P. G. and E. K.-G. designed the study (development of the concept, design and protocol of the SU.VI.MAX/ SU.VI.MAX 2 studies and coordinating of data collection); K. E. A. performed the statistical analysis and wrote the article; E. K.-G. provided methodological guidance; K. E. A., V. A. A., G. M. C., E. O. V., C. J., S. H., P. G. and E. K.-G. were involved in interpreting the results and editing the manuscript for important intellectual content; E. K.-G. and K. E. A. had primary responsibility for the final content. All the authors read and approved the final manuscript.

E. O. V. was supported by Danone Research and a grant from the Association nationale de la recherche et de la technologie (CIFRE 474/2010) from October 2010 to November 2013. Danone Research had no role in study design, data collection and analysis, decision to publish or preparation of the manuscript. None of the other authors declare any conflicts of interest.

Supplementary material

For supplementary material/s referred to in this article, please visit <http://dx.doi.org/doi:10.1017/S0007114516002233>

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