

THE INTEGRAL OF THE SUPREMUM PROCESS OF BROWNIAN MOTION

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Abstract

In this paper we study the integral of the supremum process of standard Brownian motion. We present an explicit formula for the moments of the integral (or area) $\mathcal{A}(T)$ covered by the process in the time interval $[0, T]$. The Laplace transform of $\mathcal{A}(T)$ follows as a consequence. The main proof involves a double Laplace transform of $\mathcal{A}(T)$ and is based on excursion theory and local time for Brownian motion.

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1. Introduction

Let $B(t)$, $t \geq 0$, be a standard Brownian motion. Consider the following associated processes: the supremum process $S(t) = \max_{0 \leq s \leq t} B(s)$ and the local time $L(t)$, which can be regarded as a measure of the time $B(t)$ spends at 0 in the interval $[0, t]$; see [10, Chapter VI] for details. It is well known that these two processes, although pathwise quite different, have the same distribution [10, Chapter VI.2],

$$\{S(t)\}_{t \geq 0} \stackrel{D}{=} \{L(t)\}_{t \geq 0},$$

where ' $\stackrel{D}{=}$ ' denotes equality in distribution.

The purpose of this paper is to study the distribution of the area under $S(t)$ or, equivalently, $L(t)$ over a given time interval $[0, T]$. That is, the integral

$$\mathcal{A}(T) := \int_0^T S(t) dt \stackrel{D}{=} \int_0^T L(t) dt. \quad (1.1)$$

For ease of notation, let $\mathcal{A} := \mathcal{A}(1)$.

The area (1.1) appeared as a random parameter when analysing displacements for linear probing hashing. The Laplace transform of \mathcal{A} , which is presented in Corollary 2.1, provided the means to prove one of the main theorems in [9].

Note that the usual Brownian scaling

$$\{B(Tt)\}_{t \geq 0} \stackrel{D}{=} \{T^{1/2}B(t)\}_{t \geq 0} \quad \text{for any } T > 0,$$

implies the corresponding scaling for the supremum process,

$$\{S(Tt)\}_{t \geq 0} \stackrel{D}{=} \{T^{1/2}S(t)\}_{t \geq 0}.$$

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Thus, for $T > 0$,

$$\mathcal{A}(T) = T \int_0^1 S(Tt) dt \stackrel{D}{=} T^{3/2} \mathcal{A}, \tag{1.2}$$

and it is enough to study \mathcal{A} .

2. Results

Let $\psi(s) := E(e^{-s\mathcal{A}})$ denote the Laplace transform of \mathcal{A} . An essential part of this paper is devoted to proving the following formula for the Laplace transform of a variation of ψ , or in other words, a *double* Laplace transform of \mathcal{A} . Such formulae have already been derived for the integral of $|B(t)|$ and other similar integrals of processes related to Brownian motion; see [8] and the survey [3].

Theorem 2.1. *Let ψ be the Laplace transform of \mathcal{A} . For all $\alpha, \lambda > 0$,*

$$\int_0^\infty \psi(\alpha s^{3/2}) e^{-\lambda s} ds = \int_0^\infty \left(1 + \frac{3\alpha s}{2\sqrt{2\lambda}}\right)^{-2/3} e^{-\lambda s} ds.$$

Remark 2.1. One of the parameters α and λ in Theorem 2.1 can be eliminated (by setting it equal to 1, for instance) without loss of generality. In fact, for any $\beta > 0$, the formula is preserved by the substitutions $\lambda \mapsto \beta\lambda$, $\alpha \mapsto \beta^{3/2}\alpha$, and $s \mapsto \beta^{-1}s$.

The proof is given in Section 5. It is based on the excursion theory for Brownian motion and is inspired by similar arguments for other Brownian areas; see [8].

Theorem 2.2. *The n th moment of \mathcal{A} is*

$$E(\mathcal{A}^n) = \frac{n! \Gamma(n + 2/3)}{\Gamma(2/3)\Gamma(3n/2 + 1)} \left(\frac{3\sqrt{2}}{4}\right)^n, \quad n \in \mathbb{N}.$$

Proof. Set $\lambda = 1$ in Theorem 2.1 and denote the left- and right-hand sides by

$$I(\alpha) := \int_0^\infty \psi(\alpha s^{3/2}) e^{-s} ds$$

and

$$J(\alpha) := \int_0^\infty \left(1 + \frac{3\alpha s}{2\sqrt{2}}\right)^{-2/3} e^{-s} ds.$$

The integrand of $I(\alpha)$ and all its derivatives with respect to α are dominated by functions of the form $s^K e^{-s}$, uniformly in $\alpha > 0$. Differentiation of $I(\alpha)$ is therefore allowed indefinitely due to dominated convergence. The same argument applies to $J(\alpha)$.

Also, the dominated convergence theorem shows that integration (with respect to s) can be interchanged with taking the limit $\alpha \rightarrow 0+$. Thus,

$$\begin{aligned} \lim_{\alpha \rightarrow 0+} \frac{d^n I(\alpha)}{d\alpha^n} &= \lim_{\alpha \rightarrow 0+} \int_0^\infty \frac{d^n}{d\alpha^n} \psi(\alpha s^{3/2}) e^{-s} ds \\ &= \int_0^\infty \lim_{\alpha \rightarrow 0+} (-s^{3/2})^n E(\mathcal{A}^n \exp\{-\alpha s^{3/2} \mathcal{A}\}) e^{-s} ds \\ &= (-1)^n E(\mathcal{A}^n) \int_0^\infty s^{3n/2} e^{-s} ds \\ &= (-1)^n \Gamma\left(\frac{3n}{2} + 1\right) E(\mathcal{A}^n) \end{aligned}$$

TABLE 1: The first four moments of \mathcal{A} .

n	$E(\mathcal{A}^n)$
1	$4/3\sqrt{2\pi}$
2	$\frac{5}{12}$
3	$64/63\sqrt{2\pi}$
4	$\frac{11}{24}$

and

$$\begin{aligned} \lim_{\alpha \rightarrow 0^+} \frac{d^n J(\alpha)}{d\alpha^n} &= \lim_{\alpha \rightarrow 0^+} \int_0^\infty \frac{d^n}{d\alpha^n} \left(1 + \frac{3\alpha s}{2\sqrt{2}}\right)^{-2/3} e^{-s} ds \\ &= \int_0^\infty \lim_{\alpha \rightarrow 0^+} \frac{\Gamma(n + 2/3)}{\Gamma(2/3)} \left(\frac{-3s}{2\sqrt{2}}\right)^n \left(1 + \frac{3\alpha s}{2\sqrt{2}}\right)^{-n-2/3} e^{-s} ds \\ &= \frac{\Gamma(n + 2/3)}{\Gamma(2/3)} \left(\frac{-3}{2\sqrt{2}}\right)^n \int_0^\infty s^n e^{-s} ds \\ &= \frac{\Gamma(n + 2/3)}{\Gamma(2/3)} \left(\frac{-3\sqrt{2}}{4}\right)^n n!. \end{aligned}$$

The fact that $I(\alpha) = J(\alpha)$ completes the proof.

The first four moments of \mathcal{A} are listed in Table 1. Furthermore, Stirling’s formula provides the asymptotic relation

$$E(\mathcal{A}^n) \sim \frac{2\sqrt{3\pi}}{3\Gamma(2/3)} n^{1/6} \left(\frac{n}{3e}\right)^{n/2}, \quad n \rightarrow \infty. \tag{2.1}$$

Corollary 2.1. *The Laplace transform of \mathcal{A} is*

$$\psi(s) = \frac{1}{\Gamma(2/3)} \sum_{n=0}^\infty \frac{\Gamma(n + 2/3)}{\Gamma(3n/2 + 1)} \left(\frac{-3\sqrt{2}s}{4}\right)^n.$$

Proof. The corollary follows from the identity

$$\psi(s) = \sum_{n=0}^\infty \frac{(-s)^n}{n!} E(\mathcal{A}^n).$$

Note that the sum converges absolutely for every complex s .

The graph of $\psi(s)$ is shown in Figure 1.

Remark 2.2. The Laplace transform of \mathcal{A} can also be expressed in terms of generalised hypergeometric functions:

$$\psi(s) = {}_1F_1\left(\frac{5}{6}; \frac{4}{6}; \frac{s^2}{6}\right) - \frac{4s}{3\sqrt{2\pi}} {}_2F_2\left(\frac{6}{6}, \frac{8}{6}; \frac{7}{6}, \frac{9}{6}; \frac{s^2}{6}\right).$$

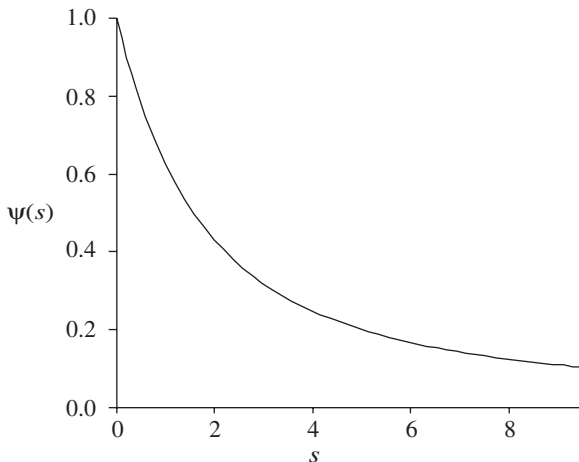


FIGURE 1: The Laplace transform of \mathcal{A} .

3. Tail asymptotics

Tauberian theorems by Davies [1] and Kasahara [7] (see [4, Theorem 4.5] for a convenient version) show that the moment asymptotics (2.1) imply the estimate $\ln P(\mathcal{A} > x) \sim -3x^2/2$ for the tail of the distribution function. Thus, the following corollary is obtained.

Corollary 3.1. \mathcal{A} has the tail estimate

$$P(\mathcal{A} > x) = \exp\left\{-\frac{3x^2}{2} + o(x^2)\right\}, \quad x \rightarrow \infty.$$

(This result can also be proved by large deviation theory; cf. similar results in [2].)

It seems difficult to obtain more precise tail asymptotics from the moment asymptotics, but it is natural to make a conjecture.

Conjecture 3.1. \mathcal{A} has a density function $f_{\mathcal{A}}(x)$ satisfying

$$f_{\mathcal{A}}(x) \sim \frac{2 \cdot 3^{1/6}}{\Gamma(2/3)} x^{1/3} \exp\left\{-\frac{3x^2}{2}\right\}, \quad x \rightarrow \infty.$$

In fact, if \mathcal{A} has a density with $f_{\mathcal{A}}(x) \sim ax^b \exp\{-cx^d\}$ for some constants $a, b, c,$ and $d,$ then it is the only possible choice that yields the moment asymptotics (2.1); cf. [5].

Conjecture 3.1 may be compared with similar results for several Brownian areas in [5]; see also [3]. Note that in these results for Brownian areas, the exponent of x is always an integer (0, 1, or 2). It is therefore of little surprise that here the exponent seems to be $\frac{1}{3}$, corresponding to the power $n^{1/6}$ in (2.1).

4. Preliminaries on point processes

Let \mathfrak{S} be a measurable space. (In this paper, \mathfrak{S} is either an interval of the real line or the product of two such intervals.) Although a point process Ξ will be regarded as a random set $\{\xi_i\} \subset \mathfrak{S}$, it is technically convenient to formally define it as an integer-valued random

measure $\sum_i \delta_{\xi_i}$. Hence, $\Xi(A)$ denotes the number of points ξ_i that belong to a (measurable) subset $A \subseteq \mathfrak{S}$. Also, $x \in \Xi$ is equivalent to $\Xi(\{x\}) > 0$. For further details, see, e.g. [6].

A Poisson process with intensity $d\mu$, where $d\mu$ is a measure on \mathfrak{S} , is a point process Ξ such that $\Xi(A)$ has a Poisson distribution with mean $\mu(A)$ for every measurable $A \subseteq \mathfrak{S}$ and $\Xi(A_1), \dots, \Xi(A_k)$ are independent for every family A_1, \dots, A_k of disjoint measurable sets. Lemma 4.1, below, is a standard formula for Laplace functionals; see, e.g. [6, Lemma 12.2(i)].

Lemma 4.1. *If Ξ is a Poisson process with intensity $d\mu$ on a set \mathfrak{S} and $f: \mathfrak{S} \rightarrow [0, \infty)$ is a measurable function, then*

$$E\left(\exp\left\{-\sum_{\xi \in \Xi} f(\xi)\right\}\right) = \exp\left\{-\int_{\mathfrak{S}} (1 - e^{-f(x)}) d\mu(x)\right\}.$$

Lemma 4.2, below, on the other hand, is more of a digression. The result follows from a standard gamma integral by integration by parts. (The result can also be written as $2\Gamma(\frac{1}{2})\lambda^{1/2}$.)

Lemma 4.2. *If $\lambda > 0$ then*

$$\int_0^\infty (1 - e^{-\lambda x})x^{-3/2} dx = 2\sqrt{\pi\lambda}.$$

5. Proof of Theorem 2.1

The set $\{t: B(t) = 0\}$ is almost surely (a.s.) closed and unbounded, so its complement $\{t: B(t) \neq 0\}$ is an infinite union of finite open intervals, denoted by $I_\nu = (g_\nu, d_\nu)$, $\nu = 1, 2, \dots$, in some order. (The intervals cannot be ordered by appearance, since there is a.s. an infinite number of them in, say, $[0,1]$. Fortunately, the order does not matter.) The restrictions of $B(t)$ to these intervals are called the *excursions* of $B(t)$. Let \hat{e}_ν be the excursion during I_ν .

The local time $L(t)$ is constant during each excursion. Let τ_ν be the local time during \hat{e}_ν and let $\ell_\nu := d_\nu - g_\nu$ be the length of \hat{e}_ν . It is well known (see [10, Chapter XII]) that the collection of pairs $\{(\tau_\nu, \ell_\nu)\}_{\nu=1}^\infty$ forms a Poisson process in $[0, \infty) \times (0, \infty)$ with intensity

$$d\Lambda = (2\pi\ell^3)^{-1/2} d\tau d\ell.$$

Note also that, a.s., if the excursion \hat{e}_{ν_1} comes before \hat{e}_{ν_2} then $\tau_{\nu_1} < \tau_{\nu_2}$.

Next, consider a Poisson process $\{T_i\}_{i=1}^\infty$ on $[0, \infty)$ with intensity λdt , independent of $\{B(t)\}$. Assume that the points are ordered with $0 < T_1 < T_2 < \dots$. Then $T_1, T_2 - T_1, \dots$ are i.i.d. $\text{Exp}(\lambda)$ random variables with density function $\lambda e^{-\lambda t}$. Furthermore, T_1 is independent of $\{B(t)\}$ and, thus, of $\{\mathcal{A}(T)\}$. It follows from (1.2) that $\mathcal{A}(T_1) \stackrel{D}{=} T_1^{3/2} \mathcal{A}$ and, consequently,

$$E(\exp\{-\alpha \mathcal{A}(T_1)\}) = E(\exp\{-\alpha T_1^{3/2} \mathcal{A}\}) = E(\psi(\alpha T_1^{3/2})) = \lambda \int_0^\infty e^{-\lambda s} \psi(\alpha s^{3/2}) ds. \tag{5.1}$$

The times T_i are called *marks*, and an excursion is called *marked* if it contains at least one of the marks T_i . The marks $\{T_i\}$ are placed by first constructing $\{B(t)\}$ and then adding marks according to independent Poisson processes with intensities λdt in each excursion. Thus, given the excursions $\{\hat{e}_\nu\}$, each excursion \hat{e}_ν is marked with probability $1 - \exp\{-\lambda \ell_\nu\}$, independently of the other excursions. The Poisson process $\Xi := \{(\tau_\nu, \ell_\nu)\}$ defined by the excursions can be written as the union $\Xi' \cup \Xi''$, where

$$\Xi' := \{(\tau_\nu, \ell_\nu): \hat{e}_\nu \text{ is unmarked}\}, \quad \Xi'' := \{(\tau_\nu, \ell_\nu): \hat{e}_\nu \text{ is marked}\}.$$

By the general independence properties of Poisson processes, Ξ' and Ξ'' are *independent* Poisson processes with intensities

$$d\Lambda' := e^{-\lambda\ell} d\Lambda = (2\pi)^{-1/2} \ell^{-3/2} e^{-\lambda\ell} d\tau d\ell \tag{5.2}$$

and

$$d\Lambda'' = (1 - e^{-\lambda\ell}) d\Lambda = (2\pi)^{-1/2} \ell^{-3/2} (1 - e^{-\lambda\ell}) d\tau d\ell,$$

respectively. In particular, if the lengths are ignored, the local times of the marked excursions form a Poisson process $\tilde{\Xi}$ on $(0, \infty)$ with intensity

$$\int_{\ell=0}^{\infty} (1 - e^{-\lambda\ell}) d\Lambda = \tilde{\lambda} d\tau,$$

where, using Lemma 4.2,

$$\tilde{\lambda} = \int_0^{\infty} (2\pi)^{-1/2} \ell^{-3/2} (1 - e^{-\lambda\ell}) d\ell = \sqrt{2\lambda}. \tag{5.3}$$

Owing to the fact that $B(T_1) \neq 0$ a.s., there exists a unique excursion \hat{e}_{ν^*} that contains the first mark T_1 , i.e. $T_1 \in I_{\nu^*}$. Let $\zeta := L(T_1) = \tau_{\nu^*}$ be the local time at T_1 (and, thus, during \hat{e}_{ν^*}). Since \hat{e}_{ν^*} is the first marked excursion, its local time ζ is the first of the points in the Poisson process $\tilde{\Xi}$ and, hence,

$$\zeta \sim \text{Exp}(\sqrt{2\lambda}). \tag{5.4}$$

The restriction of $B(t)$ to the interval $[0, T_1]$ consists of all excursions \hat{e}_{ν} with local time $\tau_{\nu} < \tau_{\nu^*} = \zeta$ and the part of \hat{e}_{ν^*} on (g_{ν^*}, T_1) , plus the set

$$[0, T_1] \setminus \bigcup_{\nu} I_{\nu} = \{t \leq T_1 : B(t) = 0\},$$

which, a.s., has measure 0 and, thus, may be ignored. Consequently, since $L(t) = \tau_{\nu}$ on I_{ν} ,

$$\begin{aligned} \mathcal{A}(T_1) &:= \int_0^{T_1} L(t) dt \\ &= \sum_{\nu: \tau_{\nu} < \tau_{\nu^*}} \int_{I_{\nu}} L(t) dt + \int_{g_{\nu^*}}^{T_1} L(t) dt \\ &= \sum_{\nu: \tau_{\nu} < \zeta} \tau_{\nu} \ell_{\nu} + \zeta(T_1 - g_{\nu^*}) \\ &=: \mathcal{A}' + \mathcal{A}''. \end{aligned}$$

The sum defined as $\mathcal{A}' = \sum_{\nu: \tau_{\nu} < \zeta} \tau_{\nu} \ell_{\nu}$ only contains terms for unmarked excursions \hat{e}_{ν} . Thus,

$$\mathcal{A}' = \sum_{(\tau_{\nu}, \ell_{\nu}) \in \Xi': \tau_{\nu} < \zeta} \tau_{\nu} \ell_{\nu}.$$

Recall that ζ is determined by Ξ'' (as the smallest τ with $(\tau, \ell) \in \Xi''$ for some ℓ) and that Ξ' and Ξ'' are independent. Hence, Ξ' and ζ are independent. It follows from Lemma 4.1, with $\Theta = (0, \zeta) \times (0, \infty)$ and $f((\tau, \ell)) = \alpha\tau\ell$, that

$$E(e^{-\alpha\mathcal{A}'} \mid \zeta) = \exp\left\{-\int_{\tau=0}^{\zeta} \int_{\ell=0}^{\infty} (1 - e^{-\alpha\tau\ell}) d\Lambda'(\tau, \ell)\right\}.$$

By (5.2) and Lemma 4.2,

$$\begin{aligned} \int_{\tau=0}^{\zeta} \int_{\ell=0}^{\infty} (1 - e^{-\alpha\tau\ell}) d\Delta'(\tau, \ell) &= \int_{\tau=0}^{\zeta} \int_{\ell=0}^{\infty} (1 - e^{-\alpha\tau\ell})(2\pi)^{-1/2} \ell^{-3/2} e^{-\lambda\ell} d\ell d\tau \\ &= (2\pi)^{-1/2} \int_{\tau=0}^{\zeta} \int_{\ell=0}^{\infty} (e^{-\lambda\ell} - e^{-(\lambda+\alpha\tau)\ell}) \ell^{-3/2} d\ell d\tau \\ &= \int_{\tau=0}^{\zeta} \sqrt{2}(\sqrt{\lambda + \alpha\tau} - \sqrt{\lambda}) d\tau \\ &= \frac{2\sqrt{2}}{3\alpha} ((\lambda + \alpha\zeta)^{3/2} - \lambda^{3/2}) - \sqrt{2\lambda}\zeta, \end{aligned}$$

and it follows that

$$E(e^{-\alpha\mathcal{A}'} \mid \zeta) = \exp\left\{ \sqrt{2\lambda}\zeta - \frac{2\sqrt{2}}{3\alpha} ((\lambda + \alpha\zeta)^{3/2} - \lambda^{3/2}) \right\}. \tag{5.5}$$

Now consider $\mathcal{A}'' = \zeta(T_1 - g_{v^*})$. Note that $T_1 - g_{v^*}$ is the location (relative to the left endpoint of the excursion) of the first mark in the first marked excursion. Since Ξ is a Poisson process with intensity independent of τ , the location $T_1 - g_{v^*}$ is independent of the local time ζ of the first marked excursion. Furthermore, the joint distribution of $(\ell_{v^*}, T_1 - g_{v^*})$ has density

$$(\tilde{\lambda})^{-1} \lambda e^{-\lambda y} (2\pi)^{-1/2} \ell^{-3/2} d\ell dy, \quad 0 < y < \ell < \infty,$$

where the normalisation constant $\tilde{\lambda}$ is given by (5.3). Consequently,

$$\begin{aligned} E(e^{-\alpha\mathcal{A}''} \mid \zeta) &= E(\exp\{-\alpha\zeta(T_1 - g_{v^*})\} \mid \zeta) \\ &= \int_{y=0}^{\infty} \int_{\ell=y}^{\infty} e^{-\alpha\zeta y} (\tilde{\lambda})^{-1} \lambda e^{-\lambda y} (2\pi)^{-1/2} \ell^{-3/2} d\ell dy \\ &= \pi^{-1/2} \lambda^{1/2} \int_{y=0}^{\infty} e^{-(\lambda+\alpha\zeta)y} y^{-1/2} dy \\ &= \lambda^{1/2} (\lambda + \alpha\zeta)^{-1/2}. \end{aligned} \tag{5.6}$$

Again, since Ξ' and Ξ'' are independent, \mathcal{A}' and \mathcal{A}'' are conditionally independent given ζ . Thus, (5.5) and (5.6) yield

$$\begin{aligned} E(\exp\{-\alpha\mathcal{A}(T_1)\} \mid \zeta) &= E(e^{-\alpha\mathcal{A}'} \mid \zeta) E(e^{-\alpha\mathcal{A}''} \mid \zeta) \\ &= \left(\frac{\lambda}{\lambda + \alpha\zeta} \right)^{1/2} \exp\left\{ \sqrt{2\lambda}\zeta - \frac{2\sqrt{2}}{3\alpha} ((\lambda + \alpha\zeta)^{3/2} - \lambda^{3/2}) \right\}. \end{aligned}$$

By (5.4), ζ has the density $\sqrt{2\lambda}e^{-\sqrt{2\lambda}x}$, $x > 0$, and it follows that

$$E(\exp\{-\alpha\mathcal{A}(T_1)\}) = \lambda\sqrt{2} \int_0^{\infty} (\lambda + \alpha x)^{-1/2} \exp\left\{ -\frac{2\sqrt{2}}{3\alpha} ((\lambda + \alpha x)^{3/2} - \lambda^{3/2}) \right\} dx.$$

Finally, the substitution

$$\frac{2\sqrt{2}}{3\alpha\lambda} ((\lambda + \alpha x)^{3/2} - \lambda^{3/2}) \mapsto s$$

provides the slightly simpler formula

$$E(\exp\{-\alpha \mathcal{A}(T_1)\}) = \lambda \int_0^\infty \left(1 + \frac{3\alpha s}{2\sqrt{2\lambda}}\right)^{-2/3} e^{-\lambda s} ds.$$

The result now follows by a comparison with (5.1).

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