

Original Article

Cite this article: Ju C., Wu R., Zhang B., You X., and Luo Y. (2020) Parenting style, coping efficacy, and risk-taking behavior in Chinese young adults. *Journal of Pacific Rim Psychology*, Volume 14, e3. <https://doi.org/10.1017/prp.2019.24>

Received: 28 December 2017
Revised: 18 November 2019
Accepted: 26 November 2019


Keywords:

parenting style; coping efficacy; risk-taking behavior; Chinese young adults

Author for correspondence:

Chengting Ju, Email: beautyjuting@126.com

Parenting style, coping efficacy, and risk-taking behavior in Chinese young adults

Chengting Ju¹ , Rina Wu², Baoshan Zhang², Xuqun You² and Yun Luo³

¹Psychological Counseling Center, Chang'an University, Xi'an, China; ²School of Psychology, Shaanxi Normal University, Xi'an, China and ³Psychological Counseling Center, Zhuhai Campus of Zunyi Medical University, Zhuhai, China

Abstract

The goal of this study was to examine the relationship between three parenting styles (warmth, rejection, and overprotection), coping efficacy and risk-taking behavior in Chinese young adults. A total of 719 subjects completed three instruments: the Egna Minne av Barndomsnauppförstran, the Coping Efficacy Questionnaire and the Adolescent Risk-Taking Questionnaire. Structural equation modeling results showed that maternal warmth (negatively) and rejection (positively), but not overprotection, correlated with risk-taking behavior via coping efficacy, whereas paternal rejection and overprotection (positively), but not warmth, correlated with risk-taking behavior via coping efficacy. These results suggest that parenting styles indirectly associate with risk-taking behavior through coping efficacy in young adults.

Risk-taking behavior refers to engaging in actions that are associated with potentially adverse consequences (Boyer, 2006). Adolescents and young adults have more frequent and prevalent risk-taking behavior than individuals in any other age group (Arnett, 1996). This may be the result of imbalances between an increased susceptibility to motivational stimuli (e.g., rewards) and the ongoing development of cognitive control (e.g., inhibition) during adolescence (Steinberg, 2008). Studies in both Eastern and Western cultures have found that young adults have a tendency to engage in risky behaviors, such as speeding, drinking, smoking and having unsafe sex. For example, Vassallo et al. (2007) found that young adults tended to speed, which was a major contributing factor to their having the highest rates of injury and death in road traffic accidents compared with Australian drivers from other age groups. The situation is similar in China. A study by An (2015) found that 17.6% of high and 39.3% of moderate sensation seekers among 786 Chinese young adults were punished for speeding every year. Moreover, Xu and Deng (2016) reported that about 11.2% of Chinese young adults had a dangerous and harmful drinking problem, which can impede healthy physical development and hinder academic performance. With so many young adults taking risks involving so many negative consequences, it is very important to form a comprehensive model by which to understand risk-taking behaviors in young adults.

According to social cognitive theory, the development of risk-taking behavior might be related to both external stimuli (e.g., family, school, peers) and internal factors (e.g., personal characteristics, self-efficacy, self-regulation; Perry, Baranowski, & Parcel, 1990). As a typical and unavoidable external stimulus, parenting style is closely associated with behavioral development in adolescents and young adults. Parenting style, which refers to the typical ways that parents think, feel and behave in terms of child-rearing (Levin, 2011), has been divided into different types, such as warmth, rejection and overprotection. These parenting styles predict the propensity for behavior problems in adolescents and young adults. For instance, parental warmth (understood as physical affection, praise and other forms of emotional support) is associated with lower levels of delinquent and aggressive behavior in adolescents (Stright & Yeo, 2014). In contrast, parental rejection, which is characterized by hostility, punishment and derogation, is associated with increased externalizing problems, such as smoking (Fuemmeler et al., 2012), suicide (Sobrinho, Campos, & Holden, 2016) and aggression (Barbot, Crossman, Hunter, Grigorenko, & Luthar, 2014). Finally, parental overprotection, which refers to excessively intrusive and highly regulative parenting, is associated with an increased lifetime likelihood of smoking (Huver, Engels, Vermulst, & De, 2007). Based on these studies, we predict that parental warmth is negatively correlated with risk-taking behavior, whereas parental rejection and overprotection are positively correlated with risk-taking behavior in young adults.

As an internal factor, coping efficacy is also associated with behavior problems in adolescents and young adults. Coping efficacy refers to the global belief that one can deal with the demands made and the emotions aroused by a stressful or problematic situation (Sandler, Tein, Mehta, Wolchik, & Ayers, 2000). Researchers have postulated that adolescents with higher coping efficacy use better coping strategies, greater effort and prolonged

© The Author(s) 2020. This is an Open Access article, distributed under the terms of the Creative Commons Attribution licence (<http://creativecommons.org/licenses/by/4.0/>), which permits unrestricted re-use, distribution, and reproduction in any medium, provided the original work is properly cited.

persistence to overcome barriers or life stressors, resulting in fewer externalized behavior problems (Hampel & Petermann, 2006). Consistently, empirical studies have found that young adults who lack confidence in their coping abilities (i.e., poor coping efficacy) are more likely to have problematic eating attitudes and behaviors and may be at risk for disordered eating problems (MacNeil, Esposito-Smythers, Mehlenbeck, & Weismore, 2012). These findings suggest that coping efficacy may be negatively related to risk-taking behavior in young adults.

There is also a significant relationship between parenting styles and coping efficacy in adolescents and young adults. Specifically, parental warmth is likely to include opportunities for instruction about and reinforcement of adaptive coping efforts, which may in turn increase coping efficacy (Camisasca, Miragoli, Blasio, & Grych, 2017). In contrast, parental rejection may lead to withdrawal in young adults when confronted with stressors, and this withdrawal may in turn decrease coping efficacy (Stevens, Bardeen, & Murdock, 2015). Similarly, parental overprotection may prevent the learning of action-oriented coping strategies (Spada *et al.*, 2012), which may in turn decrease coping efficacy. Thus, we can predict that parental warmth is positively correlated with coping efficacy, whereas parental rejection and overprotection are negatively correlated with coping efficacy in young adults.

In addition to these direct relationships, could there be indirect relationships between parenting styles, coping efficacy and risk-taking behavior? Some theorists have asserted that parenting may affect subsequent coping efficacy in adolescents, which in turn is related to externalization problems (Zhou *et al.*, 2008). There is some evidence for this hypothesis. Research has indicated that improvement in the mother-child relationship (e.g., increased warmth, acceptance and support) leads to increased coping efficacy in teenagers (Vélez, Wolchik, Tein, & Sandler, 2011), and that adolescents with higher coping efficacy can reduce the prevalence of problematic behaviors through effective coping with life difficulties (Smith *et al.*, 2006). Moreover, in a longitudinal study, Yaacob Idris, and Wan (2015) reported that coping efficacy played a mediating role in the relationship between maternal attachment and antisocial behavior in adolescents from divorced families. These studies suggest that parenting styles are probably associated with risk-taking behavior through coping efficacy.

A review of the literature shows that many studies have examined the relationships between parenting styles, coping efficacy and risk-taking behaviors. However, some gaps need to be addressed further. First, previous studies only focused on each parenting style separately and did not try to integrate the three parenting styles into a single, unified model that could be used to explore the relationships between these styles and coping efficacy and risk-taking behavior. Thus, whether there is a unique effect of each parenting style on coping efficacy and risk-taking behavior above and beyond the effects of the other two parenting styles is unknown. Second, most of the previous studies focused solely on maternal parenting, ignoring paternal parenting (e.g., Yaacob *et al.*, 2015; Vélez *et al.*, 2011). To the authors' knowledge, no research has examined the parenting styles of both mothers and fathers. Previous studies have indicated that maternal and paternal parenting styles are associated with variations in adolescent adjustment. For example, Milevsky, Schlechter, Netter, and Keehn (2007) found that authoritative mothering was related to higher self-esteem, life-satisfaction and lower depression, whereas authoritative fathering was only related to

lower depression. Moreover, both maternal and paternal parenting are critical dimensions in predicting adolescent functioning (Forehand & Nousiainen, 1993). For instance, researchers found that child development was influenced by fathers' supportiveness, which acted as a potential buffer against unsupportive maternal parenting, but mothers' supportiveness did not buffer against unsupportive paternal parenting (Martin, Ryan, & Brooks-Gunn, 2010). Thus, examining the respective roles of both maternal and paternal parenting styles on risk-taking behavior is necessary.

To address these gaps, the goal of the current study was to integrate both mothers' and fathers' three parenting styles (i.e., warmth, rejection or overprotection), coping efficacy and risk-taking behavior into a model and explore the respective roles of both mothers' and fathers' three parenting styles on coping efficacy and risk-taking behavior in young adults. Based on findings that the three parenting styles all have direct effects on coping efficacy and risk-taking behavior and that parenting has an indirect effect on externalizing behavior via coping efficacy, we tentatively expected that all three parenting styles (i.e., warmth, rejection and overprotection) of both mothers and fathers might have indirect effects on risk-taking behavior via coping efficacy. The three main hypotheses of this study were as follows:

- H1: Maternal (and paternal) warmth negatively correlate with risk-taking behavior, while maternal (and paternal) rejection and overprotection positively correlate with risk-taking behavior.
- H2: Higher levels of coping efficacy are associated with lower levels of risk-taking behavior.
- H3: The three parenting styles (i.e., warmth, rejection and overprotection) of both mothers and fathers all have indirect effects on risk-taking behavior via coping efficacy.

Methods

Participants

The data were collected through a college student self-report survey during 2016 in a western Chinese city, Xi'an. College students who lived in two-parent families were recruited from 16 institutions. The local institutional review board approval was obtained for this study. Incentives in the form of money (10 yuan, which is approximately equivalent to 1.50 USD in 2019) were provided to the participants. The participant sample varied with respect to socio-demographic characteristics (age, gender and family demographic variables) and seemed to be a reasonable representation of college students from the west of China. We initially recruited 720 college students to participate in our survey; however, one participant reported that she was 15 years old. Therefore, we eliminated her data in the final statistical analysis as she had not reached adulthood. Thus, the final analytic sample included 719 participants (women $n = 399$; men $n = 320$). The sample's mean age was 19.96 years ($SD = 0.97$, with an age range of 18 to 21 years). 61.05% and 55.91% of the participants' fathers and mothers, respectively, had a bachelor's degree or above.

Measures

Parenting style

The Chinese version of the short Eigna Minne av Bardndosnaupporstran (s-EMBU-C; Jiang, Lu, Jiang, & Xu, 2010) was used to measure the perceived maternal (and paternal)

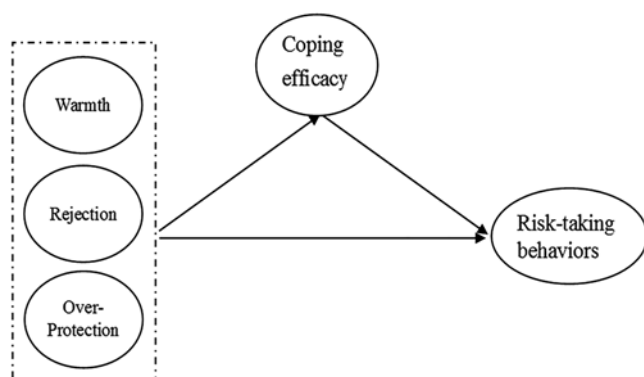


Figure 1. Theoretical model of the relationship between parenting styles, coping efficacy, and risk-taking behavior.

parenting styles. The s-EMBU-C consists of 21 self-report items with three dimensions; namely, emotional warmth (seven items), rejection (six items) and overprotection (eight items). Each item consists of two declarative sentences relating to fathering and mothering respectively. For example, the assessment form uses “Father comforts me” and “Mother comforts me” to assess warmth, “Father is strict to me” and “Mother is strict to me” to assess rejection, and “Father interferes with anything I do” and “Mother interferes with anything I do” to assess overprotection. The participants responded to each item using a 4-point Likert scale ranging from 1 = *never* to 4 = *always*. We recorded the total scores of the three dimensions for the mother and father separately. Jiang et al. (2010) reported that this scale achieved good reliability when used on a Chinese sample. In their study, the Cronbach’s alpha for both the maternal and paternal scales ranged from .74 to .84, the split-half reliability ranged from .73 to .84, and the retest reliability after 10 weeks ranged from .70 to .81. Jiang et al. (2010) also reported that the questionnaire showed good structural validity. High correlations were found between the s-EMBU-C and the original Chinese version of the EMBU (Yue, Li, Jin, & Ding, 1993): .82 for rejection, .92 for emotional warmth and .89 for overprotection. Because of its good reliability, validity and accessibility, the s-EMBU-C has been widely used with Chinese college students (Lu et al., 2015; P. Wu, Liu, Lu, & Tian, 2013). In the current study, the Cronbach’s alpha coefficients for the three subscales were acceptable: .82 for paternal and .79 for maternal emotional warmth, .77 for paternal and .74 for maternal rejection, and .67 for paternal and .68 for maternal overprotection.

Coping efficacy

Coping efficacy in the participants was measured by a Chinese version of the 17-item Coping Efficacy Questionnaire (CEQ-17; Tong, 2005). The CEQ-17 includes three dimensions: competency (nine items), cognition (five items) and confidence (three items). Examples of items pertaining to each dimension are as follows: “I am able to turn pressure into motivation” (competency); “I think trouble is a friend” (cognition); and “I’m afraid I cannot cope with these problems” (confidence; reversed items). The participants were asked, “When things are not going well for you or when you are having problems, do you have the following experience?” Then, they were instructed to provide a rating on a 4-point Likert scale ranging from 1 = *strongly disagree* to 4 = *strongly agree*. A global coping efficacy score was computed by adding the ratings on the 17 items with reverse coding of the relevant items. The scores ranged from 17 to 68, with a higher score indicating better

coping efficacy. Tong (2005) reported that the CEQ-17 had good reliability in a sample of 1806 Chinese college students. The Cronbach’s alpha coefficient was .86, the split-half reliability coefficient was .79, and the CEQ-17 also correlated with the General Self-Efficacy Questionnaire, $r(1804) = .63, p < .001$ (Schwarzer, Bäßler, Kwiatek, Schröder, & Zhang, 1997), Rosenberg’s Self-Esteem Questionnaire, $r(1804) = .53, p < .001$ (Schmitt & Allik, 2005), and the Center for Epidemiological Studies Depression Scale $r(1804) = -.48, p < .001$. In the current study, the Cronbach’s alpha coefficients were .89 for competency, .82 for cognition and .54 for confidence. Since the Cronbach’s alpha coefficient for confidence was not acceptable, this dimension was excluded from further analysis. After deleting the confidence subscale, the total Cronbach’s alpha coefficient for coping efficacy in the current study was .91. The split-half reliability coefficient was .81. The coping efficacy questionnaire used in the current study is presented in the Appendix.

Risk-taking behavior

The Chinese short version of the Adolescent Risk-taking Questionnaire — Risk Behavior Scale (ARQ-RB; Zhang, Zhang, & Shang, 2011), revised based on the Adolescent Risk-Taking Questionnaire (Gullone, Moore, Moss, & Boyd, 2000), was used to measure the levels of risk-taking behavior in college students. The ARQ-RB is a 17-item questionnaire used to rate the frequency of participation in risky behaviors. Thrill-seeking (five items), rebellious (six items), reckless (two items), and antisocial behaviors (four items) make up the four subscales of the ARQ-RB. Examples of each dimension are as follows: snow skiing (thrill-seeking), smoking (rebellious), having unprotected sex (reckless), and cheating (antisocial). Participants indicated the frequency of experiences on a 5-point Likert scale (0 = *never done*, 1 = *hardly ever done*, 2 = *done sometimes*, 3 = *done often*, and 4 = *done very often*). A total behavior score is calculated by summing up the frequency rating for all items, with a high score indicating a high overall level of participation in a risky activity. The ARQ-RB has been reported to have good reliability and validity when utilized on a Chinese adolescent sample (Zhang et al., 2011). In the current study, the Cronbach’s alpha coefficient was .66 for thrill-seeking, .76 for rebellious, .47 for reckless, and .72 for antisocial behavior. Since the Cronbach’s alpha coefficient for reckless behavior was not acceptable, this dimension was excluded from further analysis. After deleting the reckless behavior subscale, the overall alpha coefficient for the entire scale was .80, demonstrating good reliability.

Data analysis

We used structural equation modeling (SEM) to test the theoretical model (see Figure 1). In the theoretical model, parenting warmth, rejection and overprotection were three latent variables, and the items on each subscale were used to make up two parcels as indicators for each latent variable. We constructed the parcels by randomly assigning a similar number of items from each dimension to each parcel. This approach, which has been shown to have satisfactory statistical properties, is widely used in SEM (Y. Wu & Wen, 2011). Coping competence and coping cognition were the two indicators for coping efficacy; while thrill-seeking behaviors, rebellious behaviors and antisocial behaviors were the three indicators for risk-taking behavior. For SEM, a good fit is obtained when Bentler’s comparative fit index (CFI) is larger than .95, and the root mean square error of approximation (RMSEA) is

Table 1. Means, standard deviations, and bivariate correlations of variables ($N = 719$)

	1	2	3	4	5	6	7	8
1. Warmth	-							
2. Rejection	-.30** (-.40**)	-						
3. Overprotection	-.04 (-.01)	.51** (.50**)	-					
4. Coping competency	.26** (.29**)	-.20** (-.20**)	-.09* (-.08*)	-				
5. Coping cognition	.25** (.28**)	-.29** (-.25**)	-.06 (-.03)	.64**	-			
6. Thrill-seeking behaviors	.08* (.08*)	.16** (.16**)	.14** (.15**)	.03	-.04	-		
7. Rebellious behaviors	-.16** (-.13**)	.31** (.33**)	.23** (.22**)	-.23**	-.26**	.42**	-	
8. Antisocial behaviors	-.14** (-.13**)	.29** (.31**)	.19** (.21**)	-.14**	-.15**	.28**	.54**	-
<i>M</i>	2.74 (2.92)	1.49 (1.55)	1.97 (2.12)	3.14	3.44	1.96	1.68	1.83
<i>SD</i>	0.64 (0.63)	0.52 (0.51)	0.48 (0.51)	0.54	0.56	0.61	0.58	0.68

Note: *M* = mean; *SD* = standard deviation; The data outside parentheses refer to paternal parenting style, while the data inside parentheses refer to maternal parenting style. * $p < .05$; ** $p < .01$.

below .05 (Hu & Bentler, 1999). We considered the model to have an adequate fit if the CFI was larger than .90 and the RMSEA was between .06 and .08 (Yu, 2002).

To test our hypotheses, we performed the following data analysis. First, we conducted a correlation analysis using SPSS 13.0 to examine the relationships between the three maternal (and paternal) parenting styles and risk-taking (H1), and the relationships between coping efficacy and risk-taking behaviors (H2). Then we used the two-step procedure recommended by Anderson and Gerbing (1988) to analyze the indirect effect of maternal (and paternal) parenting styles on risk-taking behavior via coping efficacy using AMOS 20.0. The measurement model was first tested to assess the extent to which each latent variable was represented by its indicators. If the measurement model was shown to have an acceptable fit, then the structural model was tested using the maximum likelihood estimation. Both the measurement and structural analyses were done for the maternal and paternal parenting styles separately. The goodness of fit of the theoretical model was compared with that of two competing models to validate H3. Specifically, in Competing Model 1, the three parenting styles were specified as independent variables, while coping efficacy and risk-taking behavior were specified as dependent variables; in Competing Model 2, the three parenting styles and coping efficacy were specified as independent variables, while risk-taking behavior was specified as a dependent variable. Differences in the chi-square and fit indices were used to compare the models and to determine those that best fit the data. We used a bootstrapping analysis in AMOS 20.0 to examine the indirect effects because this method has adequate control over type I errors (Preacher & Hayes, 2008). We considered a significant indirect effect to be identified if the 95% bias-corrected bootstrap confidence intervals (BC CIs) for the indirect effect did not contain a zero or end at zero. Finally, we conducted multiple group analyses to examine the gender differences in the model.

Results

Descriptive statistics

Table 1 presents the means, standard deviations, and bivariate correlations for all the variables. Maternal (and paternal) warmth

was negatively related to all the indicators of risk-taking behavior, whereas maternal (and paternal) rejection and overprotection were positively related to all the indicators of risk-taking behavior. These results supported H1. In addition, the results of the correlation analysis showed that of the three indicators of risk-taking behavior, only antisocial behaviors and rebellious behaviors were negatively correlated with indicators of coping efficacy, while thrill-seeking behaviors had no correlations with any indicator of coping efficacy. That is, higher levels of coping efficacy were associated with lower levels of antisocial behaviors and rebellious behaviors but not associated with thrill-seeking behaviors. Thus, H2 was partially supported. We also found that maternal parenting styles and paternal parenting styles strongly correlated with each other. Specifically, maternal warmth correlated with paternal warmth, $r(717) = .67, p < .001$; maternal rejection correlated with paternal rejection, $r(717) = .76, p < .001$; and maternal overprotection with paternal overprotection, $r(717) = .71, p < .001$. These results indicate that there are some similarities between perceived maternal and paternal parenting styles within the same home, a finding which is consistent with previous studies (Winsler, Madigan, & Aquilino, 2005).

Measurement model for maternal and paternal parenting

The measurement model for all the latent variables provided a good fit to maternal parenting, $\chi^2(34, N = 719) = 111.04, p < .001, CFI = .97, RMSEA = .05$; also, the measurement model for all the latent variables provided a good fit to paternal parenting, $\chi^2(34, N = 719) = 120.27, p < .001, CFI = .96, RMSEA = .06$. In both the maternal and paternal measurement models, all the factor loadings for the indicators on the latent variables were significant ($p < .001$), indicating that all the latent factors were well represented by their respective indicators (see Table 2).

Structural model for maternal parenting

For perceived maternal parenting styles, the results showed that the theoretical model fit the data well, $\chi^2(34, N = 719) = 111.04, p < .001, CFI = .97, RMSEA = .06$. We compared the theoretical model with two other competing models. The

Table 2. Factor loadings for the indicators on the latent variables in both maternal and paternal measurement models

Items	Standardized factor loadings (Mother) <i>B</i>	Standardized factor loadings (Father) <i>B</i>
Warmth		
Warmth Parcel 1	.84	.86
Warmth Parcel 2	.69	.79
Rejection		
Rejection Parcel 1	.97	.84
Rejection Parcel 2	.68	.75
Overprotection		
Protect Parcel 1	.84	.75
Protect Parcel 2	.77	.71
Coping Efficacy		
Competency	.76	.72
Cognition	.83	.89
Risk-taking behavior		
Thrill-seeking behaviors	.47	.46
Rebellious behaviors	.89	.90
Antisocial behaviors	.62	.61

model comparison results showed that the overall fit of Competing Model 1, $\chi^2(35, N=719) = 127.10, p < .001, CFI = .96, RMSEA = .06$ was worse than the fit of the theoretical model, $\Delta\chi^2(\Delta df = 3) = 60.33, p < .001$. Additionally, Competing Model 2, $\chi^2(37, N=719) = 171.37, p < .001, CFI = .95, RMSEA = .06$, had a worse fit than did the theoretical model, $\Delta\chi^2(\Delta df = 1) = 16.06, p < .001$. As a result, the theoretical model was retained (Figure 2).

Significant relationships between maternal parenting styles, coping efficacy and risk-taking behavior were found. Specifically, coping efficacy was positively associated with maternal warmth ($B = .26, p = .01$) and negatively associated with maternal rejection ($B = -.28, p = .04$). Risk-taking behavior was negatively associated with maternal warmth ($B = -.17, p = .01$) and positively associated with maternal rejection ($B = .52, p = .01$). However, maternal overprotection was not associated with coping efficacy ($B = -.10, p = .23$) or risk-taking behavior ($B = .09, p = .18$). Coping efficacy was negatively associated with the young adults' risk-taking behavior ($B = -.21, p < .01$).

Bootstrapping analyses revealed significant indirect effects only for maternal warmth and rejection but not for overprotection. Specifically, the indirect effect of maternal warmth on risk-taking behavior via coping efficacy was significant (indirect effect = $-.05, 95\% BC CI [-.14, -.02]$). Similarly, the indirect effect of maternal rejection on risk-taking behavior via coping efficacy was significant (indirect effect = $.06, 95\% BC CI [.01, .13]$).

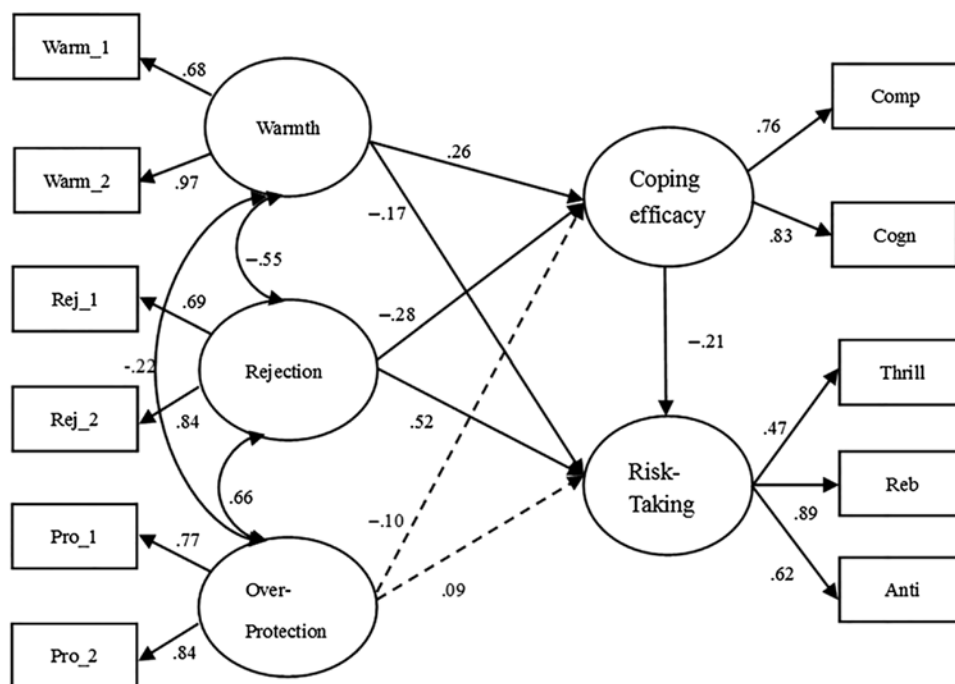


Figure 2. Structural model for the association between participants' report of their mothers' parenting, coping efficacy, and risk-taking behaviors; model fit: $\chi^2(34, N = 719) = 111.04, p < .001, CFI = .97, RMSEA = .06$. Comp = Competency; Cogn = cognition; Thrill = thrill seeking; Reb = rebellious; Anti = anti-social. Standardized path coefficients are presented. The non-significant paths are presented in dashed lines.

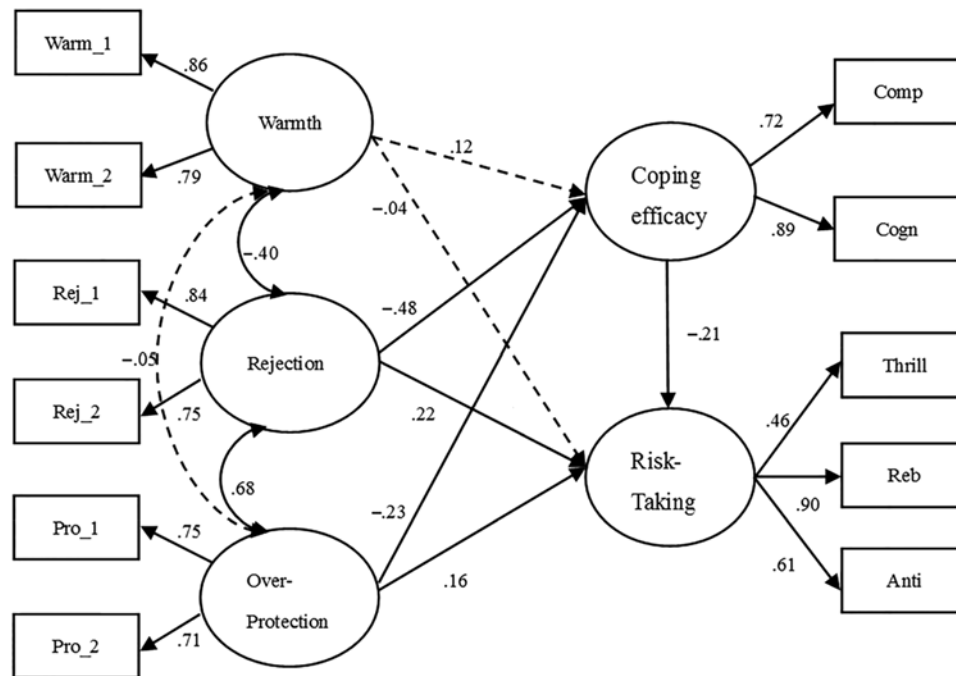


Figure 3. Structural model for the association between participants' reports of their fathers' parenting, coping efficacy and risk-taking behaviors; model fit: $\chi^2(34, N = 719) = 120.26, p < .001, CFI = .96, RMSEA = .06$. Comp = Competency; Cogn = cognition; Thrill = thrill seeking; Reb = rebellious; Anti = anti-social. Standardized path coefficients are presented. The non-significant paths are presented in dashed lines.

Structural model for paternal parenting

For perceived paternal parenting styles, the theoretical model also fit the data best, $\chi^2(34, N = 719) = 120.26, p < .001, CFI = .96, RMSEA = .06$. The model comparison results indicated that both Competing Model 1, $\chi^2(35, N = 719) = 135.24, p < .001, CFI = .95, RMSEA = .06; \Delta\chi^2(\Delta df = 1) = 14.98, p < .001$, and Competing Model 2, $\chi^2(37, N = 719) = 189.56, p < .001, CFI = .95, RMSEA = .06; \Delta\chi^2(\Delta df = 3) = 69.30, p < .001$, were worse than the theoretical model. Thus, the theoretical model was retained (see Figure 3).

Significant relationships were also found between paternal parenting styles, coping efficacy and risk-taking behavior. Specifically, coping efficacy was negatively associated with paternal rejection ($B = -.48, p < .01$) and paternal overprotection ($B = -.23, p = .04$) but was not associated with paternal warmth ($B = .12, p = .14$). In addition, risk-taking behavior was positively associated with paternal rejection ($B = .22, p = .05$) and paternal overprotection ($B = .16, p < .01$) but not associated with paternal warmth ($B = -.04, p = .57$). Coping efficacy was negatively associated with young adults' risk-taking behavior ($B = -.21, p < .01$).

The bootstrapping analyses only revealed significant indirect effects for paternal rejection and overprotection but not for warmth. Specifically, the indirect effect of paternal rejection on risk-taking behavior via coping efficacy was significant (indirect effect = .10, 95% BC CI [.04, .21]) as was the indirect effect of paternal overprotection on risk-taking behavior via coping efficacy (indirect effect = $-.05$, 95% BC CI [$-.13, -.01$]).

Therefore, with respect to maternal parenting, warmth and rejection (but not overprotection) had indirect effects on risk-taking behavior via coping efficacy, whereas with respect to paternal parenting, rejection and overprotection (but not warmth)

had indirect effects on risk-taking behavior via coping efficacy. These results partially support H3.

Gender differences

A multigroup analysis was used to identify whether the path coefficients differed significantly between the male and female participants in the theoretical model. Two models were contrasted for the maternal parenting variables. The first allowed the structural paths to vary across gender, and the second constrained all the structural paths across genders to be equal. The chi-square differences between the two models were not significant, $\Delta\chi^2(\Delta df = 7) = 2.80, p = .90$. Similarly, no gender differences in the theoretical model for paternal parenting variables were found, $\Delta\chi^2(\Delta df = 7) = 2.88, p = .89$. These results suggest that the theoretical model did not differ according to gender.

Discussion

This study examined the relationships between three parenting styles (i.e., warmth, rejection, and overprotection), coping efficacy, and risk-taking behavior using a sample of Chinese college students. The results of the correlation analysis showed that the three parenting styles of both mothers and fathers all had a direct association with risk-taking behavior in young adults. Moreover, higher levels of coping efficacy were associated with lower levels of risk-taking behavior (only antisocial and rebellious dimensions). Structural equation modeling results showed that maternal warmth (negatively) and rejection (positively), but not overprotection, correlated with risk-taking behavior via coping efficacy, whereas paternal rejection and overprotection (positively), but not warmth, correlated with risk-taking behavior via coping efficacy. Our findings suggest that perceived parenting styles of

mothers (or fathers) associate with risk-taking behavior through coping efficacy in young adults.

As expected, we found that mothers' and fathers' three parenting styles (i.e., warmth, rejection and overprotection) all had direct associations with risk-taking behavior in young adults. Specifically, maternal (and paternal) warmth were negatively related to risk-taking behavior, whereas maternal (and paternal) rejection and overprotection were positively related to risk-taking behavior. These results are consistent with earlier studies that reported a relationship between parenting style and risk-taking behavior (Coley, Carrano, & Lewin-Bizan, 2011; Davids, Roman, & Leach, 2016; Hoeve et al., 2009).

In addition, the current study found that the perceived parenting styles of mothers (or fathers) relate to risk-taking behavior through coping efficacy in young adults. This result was consistent with previous studies that found that parents shape their children's social behavior indirectly through influences on the latter's self-regulation and autonomy (Schneider, Atkinson, & Tardif, 2001). According to social cognitive theory, human behavior is explained in terms of a triadic, dynamic and reciprocal model of interrelations among behaviors, personal factors (including cognitions) and environmental influences (Perry et al., 1990). In this study, we consistently found support for an association between parenting styles (as environmental influence) and young adults' risk-taking behaviors through their coping efficacy (as personal factor).

The most interesting finding of this study was that only mothers' (not fathers') warmth was found to negatively correlate with risk-taking behavior via coping efficacy, while only fathers' (not mothers') overprotection was found to positively correlate with risk-taking behavior via coping efficacy. Why would this be the case? We assume that these distinct patterns may be explained by gender role differentiation in Chinese families. Chinese mothers are expected to be kind and to assume the role of caregivers, whereas Chinese fathers are expected to be strict and assume the role of disciplinarians, as illustrated in the Chinese phrase "yan-fu-ci-mu" ("strict father, kind mother"). In addition, mothers and fathers show their warmth and overprotection for their children in different ways. For example, mothers often show their warmth via emotional support, whereas fathers often show their warmth via instrumental support (Liu, Li, Lv, & Li, 2013). As a result, maternal warmth may be more likely to be perceived as a more positive emotional experience than paternal warmth. Thus, maternal (not paternal) warmth was found to negatively correlate with risk-taking behavior via coping efficacy. Similarly, mothers often overprotect in what the child may perceive as minor aspects of life (e.g., mothers' over-control of their children's eating or clothing or making friends), whereas fathers often overprotect in what the child may perceive to be the more important aspects of life (e.g., father's over-control of their children's choice of a professional direction, a school, a job career). As a result, paternal overprotection might be more likely to be perceived as an obstacle to self-identity than maternal overprotection (Wang, 2005). Thus, paternal (not maternal) overprotection was found to positively correlate with risk-taking behavior via coping efficacy.

To our knowledge, our study is the first to focus on the three maternal and paternal parenting styles (i.e., warmth, rejection and overprotection) and ways in which they may be associated with risk-taking behavior via coping efficacy in Chinese young adults. We highlight the importance of coping efficacy, as it reflects

an individual's capability to have confidence in performing a behavior (including overcoming problems in performing the behavior). It can help young adults cope with stress more positively, thus reducing the propensity toward risk-taking behavior. Moreover, based on our results, we think that parents can be encouraged to adopt appropriate parenting styles to prevent risk-taking behavior in young adults. Specifically, it is important to encourage Chinese mothers to show more warmth and less rejection, while encouraging Chinese fathers to show less overprotection and rejection in the rearing of their children. Such parenting styles could help increase coping efficacy in young adults, which could further prevent young adults from risk-taking behavior.

Certain limitations of this study might influence the interpretation of the results. First, the study used a cross-sectional design, which means that causal inferences cannot be obtained from these data. Thus, the results of the indirect association analysis presented here should be interpreted with caution. Future studies should use a longitudinal design to test the mediating effect of coping efficacy in the relationship between parenting style and risk-taking behavior in adolescents. Second, the data in this study were collected only through self-report measures, which could have led to subjective social desirability bias. Future studies should use multiple methods to evaluate parenting practices and risk-taking behavior, to reduce the impact of subjectivity. Third, the study only sampled college students who live in two-parent homes, while those who live with a single parent were not included. Therefore, whether the results obtained from this study could be applied to those with other family structures cannot be conclusively ascertained. Future studies should examine the relationships between parenting styles, coping efficacy, and risk-taking behaviors in young adults who grew up in divorced or other single-parent families.

Acknowledgments. The authors would like to thank the Fundamental Research Funds for the Central Universities (300102508638) and the MOE (Ministry of Education in China) Project of Humanities and Social Sciences (17XJC190009) for funding this research.

Financial Support. None.

References

- An G. (2015). Investigation and intervention research on the status of sensation seeking and risky behavior among college students [in Chinese]. *Modern Education Science: High Education Research Edition*, 2, 45–52.
- Anderson J.C. and Gerbing D.W. (1988). "Structural Equation Modeling in Practice: A Review of the Two-Step Approach". *Psychological Bulletin*, 103, 411–423.
- Arnett J.J. (1996). Sensation seeking, aggressiveness, and adolescent reckless behavior. *Personality & Individual Differences*, 20, 693–702.
- Barbot B., Crossman E., Hunter S.R., Grigorenko E.L. and Luthar S.S. (2014). Reciprocal influences between maternal parenting and child adjustment in a high-risk population: A 5-year cross-lagged analysis of bidirectional effects. *American Journal of Orthopsychiatry*, 84, 567–580.
- Boyer T. (2006). The development of risk-taking: A multi-perspective review. *Developmental Review*, 26, 291–345.
- Camasca E., Miragoli S., Blasio P.D. and Grych J. (2017). Children's coping strategies to inter-parental conflict: The moderating role of attachment. *Journal of Child & Family Studies*, 26, 1–13.
- Coley R.L., Carrano J. and Lewin-Bizan S. (2011). unpacking links between fathers' antisocial behaviors and children's behavior problems: Direct, indirect, and interactive effects. *Journal of Abnormal Child Psychology*, 39, 791–804.

- Daivids E.L., Roman N.V. and Leach L. (2016). Decision making styles: A systematic review of their associations with parenting. *Adolescent Research Review*, 1, 1–22.
- Forehand R. and Nousiainen S. (1993). Maternal and paternal parenting: Critical dimensions in adolescent functioning. *Journal of Family Psychology*, 7, 213–221.
- Fuemmeler B.F., Yang C., Costanzo P., Hoyle R.H., Siegler I.C., Williams R.B. and Ostbye T. (2012). Parenting styles and body mass index trajectories from adolescence to adulthood. *Health Psychology*, 31, 441.
- Gullone E., Moore S., Moss S. and Boyd C. (2000). The Adolescent Risk-Taking Questionnaire: Development and psychometric evaluation. *Journal of Adolescent Research*, 15, 231–250.
- Hampel P. and Petermann F. (2006). Perceived stress, coping, and adjustment in adolescents. *Journal of Adolescent Health*, 38, 409–415.
- Hoeve M., Dubas J.S., Eichelsheim V.I., van der Laan P.H., Smeenk W. and Gerris J.R. (2009). The relationship between parenting and delinquency: A meta-analysis. *Journal of Abnormal Child Psychology*, 37, 749–775.
- Hu L.-T. and Bentler P.M. (1999). Cutoff criteria for fit indexes in covariance structure analysis: Conventional criteria versus new alternatives. *Structural Equation Modeling*, 6, 1–55.
- Huiver R.M., Engels R.C., Vermulst A.A. and De V.H. (2007). Bi-directional relations between anti-smoking parenting practices and adolescent smoking in a Dutch sample. *Health Psychology*, 26, 762–768.
- Jiang J., Lu Z., Jiang B. and Xu Y. (2010). Revision of the short-form Egna Minnen av Barndoms Uppfostran for Chinese [in Chinese]. *Psychological Development and Education*, 26, 94–99.
- Levin E. (2011). Parenting styles. In S. Goldstein & J.A. Naglieri (Eds.), *Encyclopedia of child behavior and development* (pp. 1063–1064). Boston, MA: Springer US.
- Liu L., Li Y., Lv Y. and Li Y. (2013). The effects of father's involvement and parenting on children's early social skills [in Chinese]. *Psychological Development and Education*, 29, 38–45.
- Lu Y., He J., Zhu H., Wu S., Cai T., Hu X. and Mao W. (2015). Effects of parenting styles on eating disorders in teenagers: Mediating role of self-control [in Chinese]. *Chinese Journal of Clinical Psychology*, 23, 473–476.
- MacNeil L., Esposito-Smythers C., Mehlenbeck R. and Weismore J. (2012). The effects of avoidance coping and coping self-efficacy on eating disorder attitudes and behaviors: A stress-diathesis model. *Eating Behaviors*, 13, 293–296.
- Martin A., Ryan R.M. and Brooks-Gunn J. (2010). When fathers' supportiveness matters most: Maternal and paternal parenting and children's school readiness. *Journal of Family Psychology*, 24, 145–155.
- Milevsky A., Schlechter M., Netter S. and Keehn D. (2007). Maternal and paternal parenting styles in adolescents: Associations with self-esteem, depression and life-satisfaction. *Journal of Child & Family Studies*, 16, 39–47.
- Perry C.L., Baranowski T. and Parcel G.S. (1990). How individuals, environments, and health behavior interact: Social learning theory. In K. Glanz, P.M. Lewis, and B. Rimer (Eds.), *Health Behavior and Health Education: Theory, Research, and Practice*. (pp. 161–186). San Francisco, CA: Jossey-Bass.
- Preacher K.J. and Hayes A.F. (2008). Asymptotic and resampling strategies for assessing and comparing indirect effects in multiple mediator models. *Behavior Research Methods*, 40, 879–891.
- Sandler I.N., Tein J.Y., Mehta P., Wolchik S. and Ayers T. (2000). Coping efficacy and psychological problems of children of divorce. *Child Development*, 71, 1099–1118.
- Schmitt D.P. and Allik J. (2005). Simultaneous administration of the Rosenberg Self-Esteem Scale in 53 nations: Exploring the universal and culture-specific features of global self-esteem. *Journal of Personality and Social Psychology*, 89, 623–642.
- Schneider B.H., Atkinson L. and Tardif C. (2001). Child–parent attachment and children's peer relations: A quantitative review. *Developmental Psychology*, 37, 86–100.
- Schwarzer R., Bäßler J., Kwiatek P., Schröder K. and Zhang J.X. (1997). The assessment of optimistic self-beliefs: Comparison of the German, Spanish, and Chinese versions of the General Self-efficacy Scale. *Applied Psychology*, 46, 69–88.
- Smith C.L., Eisenberg N., Spinrad T.L., Chassin L., Morris A.S., Kupfer A. . . . Kwok O.M. (2006). Children's coping strategies and coping efficacy: Relations to parent socialization, child adjustment, and familial alcoholism. *Development & Psychopathology*, 18, 445–469.
- Sobrinho A.T., Campos R.C. and Holden R.R. (2016). Parental rejection, personality, and depression in the prediction of suicidality in a sample of nonclinical young adults. *Psychoanalytic Psychology*, 33, 554–570.
- Spada M.M., Caselli G., Manfredi C., Rebecchi D., Rovetto F., Ruggiero G.M. . . . Sassaroli S. (2012). Parental overprotection and metacognitions as predictors of worry and anxiety. *Behavioural & Cognitive Psychotherapy*, 40, 287–296.
- Steinberg L. (2008). A social neuroscience perspective on adolescent risk-taking. *Developmental Review*, 28, 78–106.
- Stevens E.N., Bardeen J.R. and Murdock K.W. (2015). Parenting behaviors and anxiety in young adults: Effortful control as a protective factor. *Journal of Individual Differences*, 36, 170–176.
- Stright A.D. and Yeo K.L. (2014). Maternal parenting styles, school involvement, and children's school achievement and conduct in Singapore. *Journal of Educational Psychology*, 106, 301–314.
- Tong H. (2005). Coping efficacy: Establishment questionnaire and theoretical models' construction [in Chinese]. *Acta Psychologica Sinica*, 37, 413–419.
- Vélez C.E., Wolchik S.A., Tein J.Y. and Sandler I. (2011). Protecting children from the consequences of divorce: A longitudinal study of the effects of parenting on children's coping processes. *Child Development*, 82, 244–257.
- Vassallo S., Smart D., Sanson A., Harrison W., Harris A., Cockfield S. and McIntyre A. (2007). Risky driving among young Australian drivers: Trends, precursors and correlates. *Accident Analysis & Prevention*, 39, 444–458.
- Wang L. (2005). The history, present condition and trend of the research on fathering pattern in the Western [in Chinese]. *Advances in Psychological Science*, 13, 290–297.
- Winsler A., Madigan A.L. and Aquilino S.A. (2005). Correspondence between maternal and paternal parenting styles in early childhood. *Early Childhood Research Quarterly*, 20, 1–12.
- Wu P., Liu H., Lu L. and Tian M. (2013). The relationship between adolescent immoral behaviours on internet and parenting style [in Chinese]. *Journal of Psychological Science*, 36, 372–377.
- Wu Y. and Wen Z. (2011). Item parceling strategies in structural equation modeling [in Chinese]. *Advances in Psychological Science*, 19, 1859–1867.
- Xu Z. and Deng Y. (2016). Dangerous and harmful drinking and influencing factors in college students [in Chinese]. *Chinese Journal of Clinical Psychology*, 24, 36–40.
- Yaacob S.N., Idris F.A. and Wan G.S. (2015). Parental attachment, coping efficacy and antisocial behavior among adolescents from divorced family in Selangor, Malaysia. *Journal of Management Research*, 7, 364.
- Yu C.-Y. (2002). Evaluating cutoff criteria of model fit indices for latent variable models with binary and continuous outcomes. *Analyzing Discourse Text & Talk*, 28, 14.
- Yue D., Li M., Jin K. and Ding B. (1993). Preliminary revision of EMBU and its application in neurotic patients [in Chinese]. *Chinese Mental Health Journal*, 7, 97–101.
- Zhang C., Zhang L. and Shang L. (2011). Reliability and validity of Adolescent Risk-taking Questionnaire-Risk Behavior Scale (ARQ-RB) in middle school students [in Chinese]. *Chinese Mental Health Journal*, 25, 636–640.
- Zhou Q., Wang Y., Deng X., Eisenberg N., Wolchik S.A. and Tein J.Y. (2008). Relations of parenting and temperament to Chinese children's experience of negative life events, coping efficacy, and externalizing problems. *Child Development*, 79, 493–513.

Appendix

Coping Efficacy Questionnaire (Tong, 2005)	
English	Chinese
*1. I think I'm fragile.	*1、我觉得自己是脆弱的。
*2. I'm afraid I can't handle it.	*2、我担心我应付不了。
*3. I feel too tired to live.	*3、我觉得自己活得太累了。
4. I can withstand setbacks.	4、我能承受挫折与打击。
5. I can turn pressure into motivation.	5、我能将压力变成动力。
6. I think I'm tough.	6、我觉得自己是坚强的。
7. I dare to challenge life's setbacks.	7、我敢于向挫折与打击挑战。
8. I do not easily succumb to setbacks.	8、我是不轻易向挫折与打击屈服的。
9. I dare to control my own destiny.	9、我敢于扼住命运的咽喉。
10. I have the confidence to overcome any difficulty.	10、我有信心战胜任何困难。
11. I won't be defeated by life's setbacks.	11、挫折与打击不可能是我垮掉。
12. I can reasonably analyze setbacks.	12、我会理智地分析挫折和打击。
13. I actively seek solutions.	13、我会积极地寻求解决的办法。
14. I believe that life is like a boat with ups and downs.	14、我认为人生不如意是常有的事。
15. I believe that every cloud has a silver lining.	15、我认为天无绝人之路。
16. I believe that every bad thing will pass.	16、我相信一切都会过去的。
17. I will not give up on myself.	17、我不会自暴自弃的。

Note: *refers to reverse item. The confident dimension of coping efficacy includes three items: item 1-3; the competence dimension of coping efficacy includes nine items: item 4-12; the cognitive level dimension of coping efficacy includes five items: item 13-17.