

## Review

# Systematic review of risk factors for violence in psychosis: 10-year update

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## Background

Understanding risk factors for violence in people with psychosis can inform risk management and violence prevention. However, much of the evidence comes from cross-sectional studies, and previous reviews require updating.

## Aims

To synthesise evidence from longitudinal studies on risk factors for violence in people with schizophrenia-spectrum disorders, bipolar disorder or other affective psychoses.

## Method

We searched five bibliographic databases up to June 2022. We identified longitudinal studies reporting risk factors for violence in individuals diagnosed with schizophrenia or other psychoses using DSM or ICD criteria. If  $\geq 3$  independent samples reported a risk factor, we conducted random-effects meta-analyses to provide a pooled estimate. We also meta-analysed risk factors by major domains.

## Results

We identified 47 longitudinal studies on risk factors for violence in psychosis, representing 41 independent samples – 21 from the original and 20 from the updated review – and 203 297 individuals. A total of 30 risk factors were present in  $\geq 3$  independent samples. Criminal history factors were associated

with the greatest risk of violent outcomes (pooled odds ratio 3.50, 95% CI = 2.37, 5.16), followed by substance misuse factors (odds ratio 2.36, 95% CI = 1.99, 2.80). Many treatment-related factors were protective (odds ratio 0.54, 95% CI = 0.34, 0.85). Effect estimates were attenuated in inpatient settings. We also identified novel risk factors, including cannabis use, in a secondary analysis (odds ratio 3.34, 95% CI = 2.32, 4.82).

## Conclusions

Using longitudinal evidence, we have validated comorbid substance misuse and criminal history as major risk factors for violence in psychosis. Novel factors such as cannabis use need further replication. Several identified factors are possible intervention targets if associations are found to be causal.

## Keywords

Systematic review; psychosis; violence; risk factors.

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Given the high costs of violence perpetration to patients, victims and society,<sup>1,2</sup> preventing violent outcomes and improving risk management is a priority for clinical services.<sup>3</sup> Triangulated evidence shows a higher risk of violence among individuals with psychotic disorders than among those without.<sup>4–6</sup> Absolute risks of violence can be as high as 26% over 12 months in first-episode psychosis,<sup>7</sup> but are more typically less than 10% within 5 years of diagnosis by clinical services.<sup>5</sup> Identification of risk factors – particularly modifiable ones – in people with psychotic disorders is a next step in developing targeted interventions,<sup>8</sup> and could help develop more precise risk assessment tools that allow for risk stratification. Such tools are common in forensic mental health and criminal justice to aid clinical decision-making, but are of varying accuracy.<sup>9</sup> Updated evidence on risk factors is particularly relevant for treatment allocation in the context of limited resources, for example with the reduction of available psychiatric hospital beds in the UK and USA that has continued in recent years.<sup>10</sup>

A 2013 systematic review<sup>11</sup> of risk factors for violence in psychosis outlined a range of replicated risk markers, including criminal history, psychopathological symptoms and treatment-related factors. However, the previous review is now more than a decade old with its search ending in 2011, and many new investigations have since been published. The previous review also included a majority of cross-sectional studies where the temporal relationship between the studied factor and violent outcome is not clear. In this update, we have focused on longitudinal studies to improve the quality of the evidence. In addition, we have conducted separate analyses restricting studies to those using more severe violent

outcomes, and those where a majority of participants were recruited from inpatient settings. The latter can inform how to prevent and manage inpatient violence, which has been reported to occur in 21–32% of admitted individuals with psychotic disorders.<sup>12</sup> We consider risk factors in a broad way to include descriptive, causal and predictive associations.<sup>13</sup> We expected to replicate the strongest risk factors in the previous review – including criminal history and comorbid substance use disorders – while drawing on a decade of new evidence on emerging risk factors for violence in psychosis, including the misuse of individual substances.

## Method

The study was conducted following the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines.

## Protocol

The review methods are based on a previous review from our group,<sup>11</sup> published via open access in 2013. One deviation from this protocol was to limit the inclusion to longitudinal designs, which were included in the original protocol alongside cross-sectional designs. We further excluded studies in selected (e.g. solely offender) populations. There were no other material deviations from the original protocol.

## Search strategy

We implemented the same search strategy as the 2013 review,<sup>11</sup> using the following search term to identify studies examining psychiatric disorders and various violence measures: (*schiz\** AND

\* Joint second authors.

(*viol\* OR aggress\* OR crim\* OR offend\* OR danger\* OR hosti\**) OR (*psych\* AND (viol\* OR aggress\* OR crim\* OR offend\* OR danger\* OR hosti\*)*) OR (*mental\* AND (viol\* OR aggress\* OR crim\* OR offend\* OR danger\* OR hosti\*)*) (Supplementary Table 1 available at <https://doi.org/10.1192/bjp.2024.120>). We conducted the search in five databases – CINAHL, Embase, Global Health, PsycINFO and PubMed – to identify papers published from 1 January 2012 until 30 June 2022. We thus complemented the search from the original review, which searched evidence up to 31 December 2011. The extraction of references from the databases was carried out on 30 June 2022. We also conducted a manual search of reference lists in included or related papers. We translated non-English language publications using Google Translate and asked a native speaker for clarification where necessary. Search and eligibility assessment was carried out by the first author (T.L.). A second reviewer was not considered for systematic eligibility assessment, in line with the previous review.

### Study eligibility

We considered studies where the following apply: (a) at least 95% of the participants were diagnosed with schizophrenia, schizoaffective disorder, delusional disorder, other schizophrenia-spectrum disorder, schizotypal disorder, bipolar disorder or other affective psychosis, excluding drug-induced psychosis (e.g. if 94% of patients in a study had schizophrenia spectrum disorder and 6% had drug-induced psychosis or a non-psychotic disorder, then that study would be excluded); (b) diagnosis was made using DSM<sup>14</sup> or ICD<sup>15</sup> criteria; (c) individuals were aged 15 years or above; and (d) the study employed a design where risk factors preceded the outcome (e.g. cohort studies, nested case-control, randomised controlled trials (RCTs) and prediction studies); (e) further, we excluded studies that only considered repeat violence as an outcome – that is, any study examining violent outcomes in a selected population with violence histories. As in the previous review, bipolar disorder was included as psychosis is a feature in most individuals with bipolar disorder,<sup>16</sup> and to incorporate risk factors for violence in affective psychosis. We extracted information on each criterion (a)–(e) – the first criterion we noticed that made the study ineligible for inclusion was reported as the reason for that paper's exclusion.

### Data extraction

Data from included papers was extracted by the first author (T.L.). A second author (S.L.) independently extracted a randomly selected 20% subset of included studies to assess extraction accuracy. For both risk factors and violent outcomes, we extracted information on whether the variable was categorical or continuous. Risk factor definitions across studies were standardised where possible to ensure they could be pooled. In keeping with previous work,<sup>11</sup> we classified risk factors into the following broad and distinctive domains: criminal history, negative symptoms, neuropsychological, positive symptoms, premorbid, psychopathological, sociodemographic, substance misuse, suicidality and treatment-related.

To ensure comparability with previous work and allow for data pooling, we converted all effect estimates to odds ratios, using methods described in previous publications.<sup>11,17</sup> Hazard ratios and probit regression coefficients cannot strictly be converted to odds ratios. However, probit and logistic regression often give similar results, and hazard ratios can be relatively similar to odds ratios if the outcome event is rare over follow-up.<sup>18</sup> We therefore included these effect estimates in the main analysis. As a sensitivity analysis, we excluded all effect estimates measured by hazard ratios or probit regression coefficients.

Study quality was assessed by co-author A.P. using an adaptation of the Newcastle–Ottawa Scale<sup>19</sup> where we scored quality of exposure ascertainment separately for each risk factor domain

included in the study. The overall quality score was then produced by summing the quality score for each risk factor domain included with the other subscores (selection, comparability and outcome). The score was expressed as a percentage of the maximum quality score available given the risk factor domains included in the study.

### Statistical analyses

We assessed extraction accuracy by calculating interrater reliability for the extraction of the proportion of violent patients using Krippendorff's alpha.<sup>20</sup>

We pooled odds ratios using random-effects models owing to the heterogeneity in the design and predictors/outcomes of the studies. We only considered risk factors that occurred in at least three independent samples in the main analysis, but reported results for risk factors occurring in two samples in the supplement. We chose to pool the most minimally adjusted effect estimate available, as this was the most consistently reported and comparable.<sup>11,17</sup> Some included papers were based on the same original study. If a risk factor occurred in more than one paper that used the same study population, we included the risk factor that derived from the paper with the largest sample size to avoid double counting.<sup>11</sup> All analyses in the current paper were carried out at the level of independent samples rather than publications.

We additionally pooled odds ratios within each risk factor domain. If one paper contributed more than one risk factor per domain, we included the risk factor with the highest absolute *z*-score. The *z*-score takes into account the size of the effect estimate (strength of association) and its s.d. (precision).<sup>11</sup>

We assessed heterogeneity using the  $I^2$  statistic to quantify the proportion of the variance in the risk factor effect estimates due to between-study differences rather than random sampling error. To investigate the sources of between-study variability, we ran meta-regression models for risk factors that occurred in at least seven samples and had an  $I^2$  of  $\geq 75\%$ . We considered the following between-study factors in the meta-regression models: whether the study setting was in Europe or not (binary), whether a majority of study participants were recruited from an inpatient setting or not (binary), percentage of men (continuous) and whether the violent outcome was based on forensic care status/criminal record or not (binary).

### Sensitivity analyses

To account for different settings and violent outcome types across samples, we conducted analyses including risk factors from only the following: (a) samples where violence was defined by conviction/arrest (violence and/or homicide) or forensic psychiatric care (which typically requires a criminal offence); or (b) samples where 95% or more of the population was recruited from inpatient settings. We also conducted analyses where we excluded risk factors with effect estimates reported as HRs and probit regression coefficients. We further restricted analyses to those papers that were deemed to have a quality score of  $\geq 75\%$ . Finally, we conducted publication bias analyses for risk factors that occurred in at least seven independent samples using Peters' regression test.<sup>21</sup>

Data management and analyses, including effect estimate conversions, were carried out in R version 4.3.0 (R Core Team, Vienna, Austria; see <https://www.R-project.org/>).<sup>22</sup>

## Results

### Study characteristics

We identified 79 988 publications from the five listed databases after removal of duplicates; 288 remained after a title and abstract

screening. After full-text screening, 47 studies were included in this update, representing 41 independent samples (Supplementary Fig. 1, Supplementary Tables 2 and 3).<sup>5,23–68</sup> Four studies stratified their findings by gender and one further study was stratified by diagnostic category (schizophrenia *v.* bipolar disorder), and thus each contributed two independent sets of risk factors to the analyses.<sup>5,36,54,62,63</sup> Fourteen studies overall were overlapping, some of which incorporated aforementioned gender- or diagnosis-stratified samples.<sup>24–26,28,36–38,55,60,62–64,66,67</sup> Overall, 564 separate effect sizes were extracted, representing 387 unique risk factors after standardisation. Of these, 30 were examined in at least three independent samples, and an additional 29 factors occurred in two. The publication years ranged from 1983 to 2022, with the majority of samples originating from Europe ( $k = 20$ , 49%), followed by the USA ( $k = 10$ ) and the UK ( $k = 5$ ). Median sample size was 404 (Table 1). Ten samples (24%) recruited more than 95% of their participants from inpatient settings. The majority measured their violent outcome as a physical assault on another person ( $k = 12$ , 29%) or as a conviction for a violent crime ( $k = 12$ , 29%; Table 1).

We found good interrater reliability of the extraction of the proportion of violent individuals (Krippendorff's alpha 0.77).

Table 2 shows the effect estimates of risk factors that were reported in  $\geq 3$  independent samples, pooled over the individual risk factor. Fig. 1 illustrates the effect estimates when these risk factors were pooled over the risk factor domain. No risk factors in the neuropsychological domain were present in two or more samples, and hence were not included in the analyses. We allowed individual Positive and Negative Syndrome Scale (PANSS) items to be included as separate risk factors.

### Criminal history domain

Odds ratio estimates were similar across criminal history factors (Table 2), and were all associated with an increased risk of violence in people with psychosis. The most commonly reported criminal

history risk factor was 'violence history' ( $k = 15$ ), with an odds ratio of 2.91 (95% CI = 2.06, 4.10). After pooling across all criminal history factors, the overall domain was associated with more than a three-fold increased risk (odds ratio 3.50, 95% CI = 2.37, 5.16; Fig. 1). A history of non-violent crime and a family history of offending behaviour were also risk factors and were reported in two independent samples (Supplementary Table 4).

### Negative and positive symptom domains

Risk factors present in three or more independent samples were the negative symptom score on the PANSS (odds ratio 1.10, 95% CI = 0.95, 1.27) and positive symptom score on the PANSS (odds ratio 1.12, 95% CI = 0.45, 2.82) (Table 2), although neither showed a clear association with violence. In the positive symptom domain, paranoia and hostility were reported in two independent samples (Supplementary Table 4), with paranoia showing the greatest point estimate (odds ratio 3.07, 95% CI = 0.88, 10.69).

### Premorbid domain

Two premorbid factors were associated with an increased risk of violence: recent violent victimisation (odds ratio 5.81, 95% CI = 3.45, 9.78) and a history of parental violent crime (odds ratio 1.37, 95% CI = 1.15, 1.63; Table 2). Overall, the premorbid domain was associated with an increased risk of violence (odds ratio 1.79, 95% CI = 1.16, 2.77; Fig. 1). In addition, parental bereavement, ever being non-violently victimised, childhood abuse and a parental history of substance misuse were associated with violence risk, but were reported in only two independent samples (Supplementary Table 4).

### Psychopathological domain

A diagnosis of schizophrenia was associated with an increased risk of violence compared with other psychotic disorders in a given study (odds ratio 1.63, 95% CI = 1.17, 2.28; Table 2), as was

**Table 1** Characteristics of independent samples<sup>a</sup>

Characteristics	Original search (%) $k = 21$	Updated search (%) $k = 20$	Overall (%) $k = 41$
Publication year			
Median [min, max]	2004 [1983, 2010]	2016 [2013, 2022]	2011 [1983, 2022]
Region			
Asia	1 (4.8)	0 (0)	1 (2.4)
Australia	0 (0)	1 (5.0)	1 (2.4)
Europe	8 (38.1)	12 (60.0)	20 (48.8)
International collaboration	3 (14.3)	1 (5.0)	4 (9.8)
UK	2 (9.5)	3 (15.0)	5 (12.2)
USA	7 (33.3)	3 (15.0)	10 (24.4)
Sample size			
Median [min, max]	207 [16, 4035]	1221.5 [30, 58 771]	404 [16, 58 771]
Inpatient population			
No	12 (57.1)	19 (95.0)	31 (75.6)
Yes	9 (42.9)	1 (5.0)	10 (24.4)
Violence measure			
Aggression	1 (4.8)	0 (0)	1 (2.4)
Attack	10 (47.6)	2 (10.0)	12 (29.3)
Attack/verbal abuse	4 (19.0)	6 (30.0)	10 (24.4)
Forensic care	0 (0)	1 (5.0)	1 (2.4)
Homicide conviction	0 (0)	1 (5.0)	1 (2.4)
PANSS hostility (continuous)	0 (0)	2 (10.0)	2 (4.9)
Physical assault episodes (continuous)	1 (4.8)	0 (0)	1 (2.4)
Violent offence arrest	0 (0)	1 (5.0)	1 (2.4)
Violent offence conviction	5 (23.8)	7 (35.0)	12 (29.3)

PANSS, Positive and Negative Syndrome Scale.  
a. For overlapping samples, information from the sample with the greatest sample size is presented.

**Table 2** Pooled odds ratios of risk factors for violence in psychosis

Risk factor domain	Risk factor <sup>a</sup>	k	No. participants with violent outcomes	Total no. participants	Odds ratio (95% CI)	z-value	<i>i</i> <sup>2</sup>	
Criminal history	Violence history	15	5521	111 504	2.91 (2.06, 4.10)	6.07	96	
	Criminal history: prison	3	1756	10 140	3.04 (2.08, 4.44)	5.76	60	
	Violence history: recent	3	1564	9272	3.40 (1.65, 7.00)	3.32	78	
	Non-violent crime: history	6	2478	30 154	4.16 (1.66, 10.38)	3.05	99	
Negative symptoms	Negative: PANSS (continuous)	4	192	1767	1.10 (0.95, 1.27)	1.23	0	
Positive symptoms	Positive symptom score: PANSS (continuous)	5	192	2921	1.12 (0.45, 2.82)	0.25	99	
Premorbid	Victimisation: violent, recent	5	212	67 623	5.81 (3.45, 9.78)	6.62	67	
	Family history: violent crime, parent	4	4525	97 260	1.37 (1.15, 1.63)	3.47	47	
Psychopathological	Diagnosis: personality disorder	4	461	3246	2.30 (1.71, 3.09)	5.49	23	
	Admission to hospital: history, number (continuous)	3	128	386	2.65 (1.45, 4.84)	3.18	27	
	Diagnosis: schizophrenia	4	70	910	1.63 (1.17, 2.28)	2.87	0	
	Total: PANSS (continuous)	3	181	1655	0.67 (0.15, 2.95)	-0.53	97	
Sociodemographic	Education: no qualifications versus any	3	940	7697	1.46 (1.26, 1.69)	5.00	4	
	Black and minority ethnicity	4	259	2430	1.72 (1.08, 2.74)	2.28	68	
	Gender: male	12	2953	89 379	1.67 (1.06, 2.64)	2.19	96	
	Age: younger	4	983	59 827	1.63 (0.80, 3.32)	1.35	91	
	SES: low income	6	4110	49 864	1.11 (0.84, 1.47)	0.76	93	
	Living situation: homeless	3	293	1393	1.31 (0.59, 2.92)	0.67	0	
	Marital status: single	11	2917	31 384	1.07 (0.84, 1.35)	0.53	72	
	Living situation: living with others	3	202	1883	1.10 (0.72, 1.68)	0.45	0	
	Employment: unemployed	7	377	3353	0.86 (0.39, 1.91)	-0.36	90	
	Living situation: living alone	3	322	2378	0.75 (0.57, 1.00)	-1.97	0	
	Substance misuse	Substance misuse	11	2929	35 209	2.41 (1.84, 3.15)	6.42	87
		Drug misuse	11	3935	122 048	2.20 (1.72, 2.82)	6.28	86
		Alcohol misuse: history	3	1009	59 873	1.61 (1.31, 1.99)	4.45	19
Drug misuse: recent		4	254	2320	1.80 (1.35, 2.38)	4.06	0	
Alcohol misuse		7	3630	34 453	1.92 (1.38, 2.68)	3.86	87	
Suicidality	Self-harm: history	9	3324	155 356	1.74 (1.01, 2.98)	1.99	97	
Treatment-related	Medication: treatment adherence	4	261	2575	0.59 (0.33, 1.06)	-1.77	80	
	Medication: antipsychotic	4	2272	70 459	0.51 (0.27, 0.96)	-2.09	95	

PANSS, Positive and Negative Syndrome Scale; SES, socioeconomic status.  
a. Recent: within past year.

comorbid diagnosis of personality disorder (odds ratio 2.30, 95% CI = 1.71, 3.09) and number of past admissions to hospital (odds ratio 2.65, 95% CI = 1.45, 4.84). There was also an association for the overall psychopathological domain (odds ratio 1.66, 95% CI = 1.05, 2.62; Fig. 1). A lack of insight, comorbid diagnosis of antisocial personality disorder, a diagnosis of bipolar disorder, younger age at psychosis onset and traumatic brain injury were associated with violence in two independent samples (Supplementary Table 4).

### Sociodemographic domain

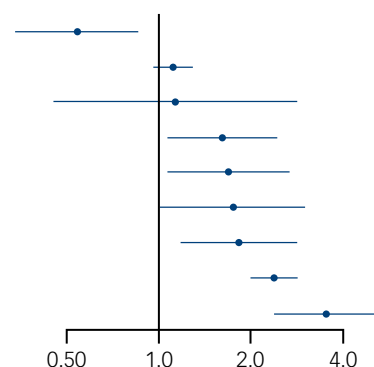
No educational qualifications versus any (odds ratio 1.46, 95% CI = 1.26, 1.69), Black and minority ethnicity (odds ratio 1.72, 95% CI = 1.08, 2.74) and male gender (odds ratio 1.67, 95% CI = 1.06, 2.64) were associated with higher violence risk

(Table 2), as was the overall sociodemographic domain (odds ratio 1.61, 95% CI = 1.07, 2.44).

### Substance misuse domain

All factors in the substance misuse domain were linked with an increased risk of violence. Substance misuse, drug misuse (current or recent), alcohol misuse and a history of alcohol misuse had similar effect sizes (odds ratios ranging from 1.61 to 2.41). The overall substance misuse domain was associated with an increased risk of violence (odds ratio 2.36, 95% CI = 1.99, 2.80; Fig. 1). Cannabis use history and recent alcohol misuse were reported in two samples, with the highest odds ratio estimated for a history of cannabis use (odds ratio 3.34, 95% CI = 2.32, 4.82; Supplementary Table 4).

Risk factor domain	k	N violent	N total	Odds ratio (95% CI)
Treatment-related	7	2533	72805	0.54 (0.34, 0.85)
Negative symptoms	4	192	1767	1.10 (0.95, 1.27)
Positive symptoms	5	192	2921	1.12 (0.45, 2.82)
Sociodemographic	25	6876	189791	1.61 (1.07, 2.44)
Psychopathological	12	760	5757	1.66 (1.05, 2.62)
Suicidality	9	3324	155356	1.74 (1.01, 2.98)
Premorbid	6	4655	98999	1.79 (1.16, 2.77)
Substance misuse	23	6841	157466	2.36 (1.99, 2.80)
Criminal history	19	5834	117 120	3.50 (2.37, 5.16)



**Fig. 1** Pooled odds ratios for violence in psychosis by overall risk factor domain.

### Suicidality domain

Self-harm history was associated with an elevated violence risk (odds ratio 1.74, 95% CI = 1.01, 2.98; Table 2; Fig. 1). Unintentional self-harm, reported in two samples, was also associated with violence (odds ratio 5.50, 95% CI = 4.26, 7.08; Supplementary Table 4).

### Treatment-related domain

Antipsychotic medication was associated with a reduced risk of violence (odds ratio 0.51, 95% CI = 0.27, 0.96; Table 2). Treatment adherence had a protective association, though not statistically significant. Antidepressant treatment was also associated with a protective effect (odds ratio 0.80, 95% CI = 0.66, 0.97; Supplementary Table 4) although it was only reported in two samples (Supplementary Table 4).

### Sensitivity analyses

When restricting analyses to those samples where a majority of patients were from inpatient settings ( $k = 10$  or 24% of the overall sample), risk factors occurring in two or more of these samples all had non-significant associations apart from the number of previous admissions, reported in two independent samples (odds ratio 2.07, 95% CI = 1.22, 3.53; Table 3). Overall, point estimates were also reduced as compared to the main analysis. Meanwhile, only including samples from community settings did not materially change results (Supplementary Table 5). When restricting analyses to samples where violence was defined by conviction or arrest for violence/homicide ( $k = 14$ ), there were 11 factors reported in  $k \geq 3$  independent samples, and findings were similar to the main analyses (Supplementary Table 6). There were no material changes when excluding risk factors where associations were reported as hazard ratios or probit regression coefficients (Supplementary Table 7). Results were also similar when restricting analyses to studies with independent samples that had quality scores  $\geq 75\%$  (Supplementary Table 8). We found no evidence of publication bias.

### Meta-regression

For the majority of included risk factors, heterogeneity as measured by the  $I^2$  statistic was high. Meta-regression was conducted for risk factors that occurred in at least seven independent samples and that had an  $I^2 \geq 75\%$ . These risk factors include unemployment, self-harm history, alcohol misuse, male gender, substance misuse, drug misuse and violence history. We only found statistically significant results for male gender. In univariate meta-regression, we found that if the study reported the outcome as violent arrest or violent conviction, then being male was statistically significantly associated with a greater risk of violence compared to when the

study used a less severe definition of violence. This was also the case for samples from Europe compared with other regions. In samples where there was a higher percentage of men, being male was associated with a lower violence risk as compared to samples where there was a lower percentage of men. None of these sample characteristics retained statistical significance when simultaneously entered into a multivariable meta-regression model.

## Discussion

In this updated systematic review and meta-analysis of risk factors for violence in psychosis based on longitudinal studies, we identified 41 independent samples comprising 203 297 individuals. We examined 30 individual risk factors that were reported in at least three independent samples. This synthesis identified novel risk factors and validates associations identified in previous work, providing information on several potentially modifiable risk factors. Importantly, the focus on longitudinal studies reduces the likelihood of reverse causation, meaning that findings can inform risk stratification and help identify potential treatment targets.

We found that criminal history was the risk factor domain with the strongest association with violence, followed by the substance misuse domain. In relation to criminal history, both previous violent and non-violent crime were risk factors – possibly because these measure a general propensity for criminality, or because engaging in criminality introduces individuals to social contexts and networks that increase the risk of subsequent violence. There are many explanations for the importance of comorbid substance misuse as a risk factor domain. Intoxication leads to poorer impulse control, which has been found to be a strong risk factor for violence in cross-sectional studies.<sup>11</sup> Individuals with schizophrenia (and other severe mental illnesses, including bipolar disorder) may also self-medicate with substances to manage symptoms,<sup>69</sup> possibly affecting treatment adherence and effectiveness,<sup>70</sup> while being an indicator of more severe symptomatology. Drug misuse may additionally be an entry route into criminality and expose individuals to violent environments.<sup>71,72</sup> Another important risk factor domain is suicidality. It is likely that common processes underpin violence and suicidality – for example, they may both be outward expressions to regulate intense internal states.<sup>71</sup> The strength of the association for these and other studied risk factor domains were broadly similar to the previous review,<sup>11</sup> validating the evidence using longitudinal designs.

An unexpected finding in this review was the lack of a clear association between overall positive symptom scores and violence, despite previous research finding that positive symptoms are an important risk factor for violence.<sup>73</sup> However, the CI for our result was wide (odds ratio 1.12, 95% CI = 0.45, 2.82), and with a

**Table 3** Pooled odds ratios of risk factors for violence in psychosis in inpatient settings

Risk factor domain	Risk factor	<i>k</i>	No. participants with violent outcomes	Total no. participants	Odds ratio (95% CI)	<i>z</i> -value	<i>I</i> <sup>2</sup>
Criminal history	Violence history	3	215	586	1.23 (0.35, 4.33)	0.33	83
Negative symptoms	Negative: PANSS (continuous)	2	104	220	1.42 (0.62, 3.23)	0.83	64
Positive symptoms	Paranoia: BPRS (continuous)	2	7	223	3.07 (0.88, 10.69)	1.76	52
	Positive symptom score: PANSS (continuous)	2	104	220	0.48 (0.07, 3.29)	-0.75	92
Psychopathological	Admission to hospital: history, number (continuous)	2	106	201	2.07 (1.22, 3.53)	2.69	0
	Total: PANSS (continuous)	2	104	220	0.41 (0.05, 3.64)	-0.80	94
Sociodemographic	Marital status: single	3	131	1922	1.22 (0.84, 1.77)	1.03	0
	Gender: male	3	267	3105	1.59 (0.44, 5.81)	0.70	92
	Living situation: living with others	2	125	448	1.28 (0.46, 3.55)	0.47	0
	Living situation: homeless	2	125	448	0.54 (0.08, 3.88)	-0.61	0
	Employment: unemployed	2	69	367	0.33 (0.01, 10.01)	-0.64	94
Suicidality	Self-harm: history	2	69	367	1.64 (0.88, 3.05)	1.57	0

PANSS, Positive and Negative Syndrome Scale; BPRS, Brief Psychiatric Rating Scale.

high  $I^2$  statistic. We also found a modest, although statistically non-significant, association for negative symptom score (odds ratio 1.10, 95% CI = 0.95, 1.27). While this may be consistent with the null association found in the previous review (odds ratio 1.00, 95% CI = 0.90, 1.20),<sup>11</sup> another possibility is that negative symptoms are a marker of disease severity and partial response to treatment. It is also possible that other symptoms associated with violence, such as hostility, are misclassified as negative symptoms owing to their overlap with blunted affect and asociality.<sup>74</sup>

Antipsychotic treatment was statistically significantly associated with reduced violence (odds ratio 0.51, 95% CI = 0.27, 0.96), in line with the protective effect expected from interventions that reduce psychotic symptoms. We also found that treatment with an antidepressant (reported in two independent samples) was a protective factor (odds ratio 0.80, 95% CI = 0.66, 0.97). It is possible that the affective component of a psychosis is being treated with the antidepressant, including anger, which has been shown to be one mechanism for violence in schizophrenia.<sup>73</sup> Alternatively, being under medical treatment may be a marker of closer contact with clinical services. Another novel finding is that cannabis use (reported in two independent samples) was identified as a risk factor for violence in the current review with an odds ratio of 3.34 (95% CI = 2.32, 4.82), whereas it had a null association in the previous review. This suggests that cannabis use disorder could be considered as part of an individualised violence risk assessment.<sup>75,76</sup> It is also possible that cannabis drives at least part of the overall association of substance misuse with violence, given that many studies did not separate out the effect of cannabis misuse from that of other substances. Previous literature has found that cannabis use is associated with earlier onset of schizophrenia and other psychotic disorders,<sup>77–79</sup> and that earlier onset of psychosis is associated with greater risk of violent crime.<sup>80</sup> We further identified traumatic brain injury as a potential risk factor, which was not reported in the previous review. While organic brain disorders have been implicated in violence,<sup>81</sup> further research is necessary – especially as we only identified this risk factor in two independent samples.

### Clinical implications

Our results confirm the importance of previous crime and previous or current substance misuse, and also identify potential novel risk factors including cannabis use. These findings can inform more precise stratification of violence risk for patients with psychotic disorders. However, the clinical impact of accurate risk prediction models depends on whether effective interventions are available. Many of the identified risk factors, such as comorbid substance misuse, are potential targets for clinical intervention if found to be causal. Others may be markers of modifiable risk factors. For example, criminal history could be a marker of pro-criminal beliefs that may be reduced using cognitive or behavioural interventions.<sup>71</sup> Further research is recommended to test the impact of different types of interventions in individuals with psychosis.


We have included multiple study settings, and it is possible that certain risk factors vary in their association depending on the context in which the assessment is made. When we restricted our analysis to inpatient settings, risk factors were attenuated and became statistically non-significant. Patients from inpatient settings are likely to be a selected sample owing to greater severity of symptoms and may have more similar risk factor profiles, leading to lower variance in violence risk. Consequently, more in-depth risk assessment may be necessary.

### Strengths and limitations

We have focused on longitudinal studies, which reduces the likelihood that reverse causation explains the results. However, several

limitations should be noted. First, we relied on minimally adjusted effect estimates, given that covariate adjustment varied across studies. This may help explain why the effect estimates are quite similar within certain domains (notably the criminal history domain). Different types of criminality often co-occur within individuals – the presence of a given type may therefore be a marker of another, obscuring the relationship of specific risk factors with violence. We did not have individual participant data, which may have allowed us to provide consistently adjusted effect estimates.<sup>82</sup> Second, causal inferences are not possible, given that we extract minimally adjusted effect estimates and include a mix of study designs with different aims (descriptive, causal and predictive). Nonetheless, our results may offer insight into risk factors that are important to consider in predictive and causal models, and – if found to be causal – that are modifiable in the sense that they can be improved through clinical intervention. Both types of risk factors are important – causal factors would allow for interventions to mitigate risk, while predictive ones would provide better risk stratification. Third, our decision to pool individual risk factors and risk factor domains needs to be considered in the context of substantial heterogeneity, for example, in terms of study designs and measures. However, pooling can provide a broad overview of a large body of evidence.<sup>11</sup> Fourth, some risk factors were only measured in a cross-sectional setting, notably those based on detailed neuropsychological or cognitive tests or biomarkers, which meant that they were not included in our review. Fifth, we did not use a second reviewer for screening, given the large number of references at the screening phase and because we followed a previous protocol. However, the original search and the updated one were carried out by independent reviewers. Sixth, included studies predominantly come from high-income Western countries and further research is needed to better understand risk factors in other regions.

In this updated systematic review and meta-analysis using longitudinal designs, we have validated the importance of comorbid substance misuse and criminal history as risk factors for violence in psychosis, while identifying potential novel risk factors such as cannabis misuse that require replication. If found to be causal, several of these factors represent modifiable intervention targets.

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### Supplementary material

Supplementary material is available online at <https://doi.org/10.1192/bjp.2024.120>.

### Data availability

Data availability is not applicable to this article as no new data were created or analysed in this study.

### Author contribution

T.L., R.Y. and S.F. designed the study. T.L. carried out the search and screening with support from R.Y. and S.F. T.L. carried out data extraction and S.L. conducted quality control. A.P. assessed the quality of included studies with support from T.L., R.Y. and S.F. T.L. carried out data management and analysis. T.L. drafted the manuscript. All authors provided detailed comments on the manuscript and contributed to the interpretation of results.

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## Declaration of interest

None.

## Transparency declaration

T.L. acts as guarantor and affirms that the manuscript is an honest, accurate and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned (and, if relevant, registered) have been explained.

## Analytic code availability

The analytic code for running the meta-analysis is available upon request.

## Research material availability

Tabular data is available upon reasonable request.

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