

Robust Huber-Based Cubature Kalman Filter for GPS Navigation Processing

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A robust state estimation technique based on the Huber-based Cubature Kalman Filter (HCKF) is proposed for Global Positioning System (GPS) navigation processing. The Cubature Kalman Filter (CKF) employs a third-degree spherical-radial cubature rule to compute the Gaussian weighted integration, such that the numerical instability induced by round-off errors can be avoided. In GPS navigation, the filter-based estimation of the position and velocity states can be severely degraded due to contaminated measurements caused by outliers or deviation from a Gaussian distribution assumption. For the signals contaminated with non-Gaussian noise or outliers, a robust scheme combining the Huber M-estimation methodology and the CKF framework is beneficial where the Huber M-estimation methodology is used to reformulate the measurement information of the CKF. GPS navigation processing using the HCKF algorithm has been carried out and the performance has been compared to those based on the Extended Kalman Filter (EKF), Unscented Kalman Filter (UKF) and CKF approaches. Simulation and experimental results presented in this paper confirm the effectiveness of the method.

KEYWORDS

1. GPS navigation.
2. Unscented Kalman filter.
3. Cubature Kalman filter
4. Huber M-estimation.

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1. INTRODUCTION. The Global Positioning System (GPS) is a satellite-based navigation system that provides a user access to useful and accurate positioning information anywhere on the globe (Axelrad and Brown, 1996). The well-known Kalman Filter (KF) (Gelb, 1974; Brown and Hwang, 1997), which provides an optimal (minimum mean square error) estimate of the system state vector, has been widely applied to the field of navigation.

In navigation filter processing, the linear model increases modelling errors since the vehicle motions are normally nonlinear. It has been common that additional fictitious process noise can be added to the system process model, nevertheless, the more suitable

cure for the divergence problem caused by mis-modelling is to correct the dynamics model. To achieve improved estimation accuracy, the designers are required to possess complete *a priori* knowledge on both the dynamic process and measurement models. The Extended Kalman Filter (EKF) relies on the first order linearization of the system model to propagate the mean and covariance of the state. The EKF is sensitive to outliers in the measurements, thus it is likely to yield poor performance if the measurements include outliers.

The Unscented Kalman Filter (UKF) (Julier, 2002; Julier et al., 1995; Julier et al., 2000) has been developed to deal with highly nonlinear state estimation problems. The UKF performs a Gaussian approximation with a limited number of points (sigma points) using the Unscented Transform (UT). The sigma points are used to propagate the probability of state distribution through the nonlinear dynamics of the system. The standard EKF is generally a first order approximation, whereas the UKF based on a second-order Taylor expansion is shown to be quite closely related to UKF. Gustafsson and Hendeby (2012) indicated that the UT gives a good approximation of many standard sensor models in tracking and navigation applications, however, UT generally does not give the same elements in the covariance as the second-order Taylor expansion. A simple counterexample was demonstrated in Gustafsson and Hendeby (2012). Furthermore, for high-dimensional systems, the UT or the UKF is believed to be unstable due to the possible negative weights on the centre point.

As a general approach to approximating the Bayesian solution, the Cubature Kalman Filter (CKF) proposed by Arasaratnam and Haykin (2009) has been an increasing interest in the development of nonlinear Bayesian filters. The CKF uses a third-degree spherical-radial cubature rule to solve the integration in the Bayesian filtering problem for numerically computing the multivariate moment integrals encountered in the nonlinear filtering framework. The CKF is considered to be a more accurate and stable nonlinear filtering approach than the UKF and has been used in many engineering applications (Fernandez-Prades and Vila-Valls, 2010). The spherical-radial cubature rule is composed of two different integrals which include spherical and radial integrals. This is based on the spherical-radial transformation and generates an even number of equally weighted cubature points. These integrals are then numerically computed by the spherical cubature rule and the Gaussian quadrature rule, respectively. The performance in terms of estimation accuracy, numerical stability and computational costs can be improved.

Gustafsson and Hendeby (2012) pointed out that the CKF can be viewed as a special case of UKF. The viewpoint provided by Wu et al. (2006) has validated the statement. As stated in this paper, the UT is essentially a Gaussian quadrature rule and other similar rules can also be applied. The spherical-radial cubature rule used in CKF is a special case of the quadrature rules. The higher-degree CKF was studied in Jia et al. (2013), and higher-order UKF has already been studied in Wu et al. (2006). The possible negative weights of the centre point in UT can be avoided through tuning the parameters (Julier et al., 2000). Using the additional tuning parameters of UKF compared to CKF provides flexibility. If one does not want to bother to tune them, we can fix them as their default values, or just exclude the centre point, and the CKF is automatically obtained.

The differences between UKF and CKF are summarised as follows: (1) The CKF uses cubature points with the same weights through the spherical-radial integration

rule, which employs an analytical Probability Density Function (PDF) to capture the mean and covariance of the posterior distribution using the total probability theorem and subsequently uses the measurement to update with Bayes' rule. From the perspective of numerical analysis, the third-degree cubature rule can be viewed as a UT of special form with better numerical stability; (2) the CKF follows directly from the spherical-radial cubature rule for numerically computing Gaussian-weighted integrals whose important property is that they do not entail any free parameters. The UKF introduces a non-zero scaling parameter, which defines the non-zero centre point and is often associated with a set of weighted samples higher than that of the minimal set of sigma points (Julier et al., 1995; Wan and Van der Merwe, 2001; Van der Merwe and Wan, 2004); (3) the UKF-calculated estimation covariance matrix is not always guaranteed to be positive definite, hence the decomposition of the covariance matrix is unavailable. To improve the stability of the nonlinear filter, the CKF can effectively avoid round-off errors of numerical computation, and it is more stable than the UKF and EKF (Jia et al., 2013).

Investigation of the CKF algorithm on the navigation problem presented by Xu et al. (2014) assumed that the Kalman-type filters are concerned with estimation of the dynamics model from noisy measurements where the dynamic and measurement processes can be approximated by Gaussian state-space models. Those nonlinear filters are very sensitive to measurement errors when the distribution of the true noise deviates from the assumed Gaussian model, often being characterised by high-intensity measurement noise and heavy tail distributions. There is considerable motivation to consider robust statistical procedures to improve the performance of the CKF in non-Gaussian environments, particularly in the presence of outliers.

Masreliez and Martin (1977) used the M-estimator to "robustify" the KF. Boncelet and Dickinson (1983) recast the measurement update as a linear regression problem and solved it using the M-estimator. Using the M-estimator, Yang (1991) robustified the general Bayes estimator by which KF had been seen as a special case. An adaptively Robust Kalman Filter (RKF), capable of dealing with uncertainties in both process and measurement models, was proposed in Yang et al. (2001) and Yang and Gao (2006a; 2006b), in which the influences of measurement outliers were controlled using the M-estimator. The Sigma Point Kalman Filter (SPKF) was robustified by Karlgaard et al. (2007), focusing on the divided difference filter. As the statistical linearization perspective of the UT is incorrectly employed in the work by Karlgaard et al. (2007) and Wang et al. (2010), Chang et al. (2012a) robustified the measurement update of the UKF using the M-estimator in a straightforward and intuitive manner. Recently, the robust sigma point Kalman smoother was studied by Karlgaard (2015). It should be noted that two kinds of nonlinearity exist, i.e., the original nonlinearity in the measurement equation, and that brought by the re-weighting procedure of the M-estimator, and the iteration becomes rather complex. If both nonlinearities were to be accounted for in the iteration, it involves issues of iterative UKF (Chang et al., 2012b; 2012c).

The Huber M-estimation methodology (Huber, 1981; Maronna et al., 2006; Gandhi and Mili, 2010; Chang and Liu, 2015) is one of the robust techniques that is based on the combination of minimum ℓ_1 - and ℓ_2 -norm estimator. The method has been successfully used for robust state estimation, inertial navigation system and visual tracking applications (Wang et al., 2010; Gao et al., 2014; Hou, 2014). The development of a RKF uses the Huber technique applied to a linear regression problem at each

measurement update (Kovacevic et al. 1992; Durgaprasad and Thakur, 1998). The Huber M-estimation methodology is essentially based on modifying the quadratic cost function in the filter framework, and works between smooth ℓ_2 – norm properties for small residuals and robust ℓ_1 – norm properties for large residuals (Petrus, 1999). Alternative functional equations, such as the Huber criterion and the hybrid ℓ_1/ℓ_2 criterion have also been considered (Huber, 1973; Bube and Langan, 1997; Guitton and Symes, 2003).

The robust Huber-based Cubature Kalman Filter (HCKF) algorithm integrates the merits of the Huber M-estimation methodology and CKF. The CKF is used to avoid the numerical instability and the Huber M-estimation methodology is incorporated to handle outliers through reformulating the measurement information in the CKF framework to provide robustness against deviations from Gaussian behaviour.

This paper is organised as follows. In Section 2, the preliminary background to the nonlinear filter is introduced, where the UKF and the CKF will be reviewed. The Huber M-estimate regression filtering strategy is presented in Section 3. In Section 4, some results from simulation and experiments are presented to evaluate GPS navigation performance for the HCKF approach in comparison to those by the conventional approach. Conclusions are given in Section 5.

2. THE NONLINEAR FILTERS. The nonlinear filtering deals with the case governed by the nonlinear stochastic difference equations:

$$\mathbf{x}_{k+1} = \mathbf{f}(\mathbf{x}_k, k) + \mathbf{w}_k \quad (1a)$$

$$\mathbf{z}_k = \mathbf{h}(\mathbf{x}_k, k) + \mathbf{v}_k \quad (1b)$$

where the state vector is $\mathbf{x}_k \in \mathfrak{X}^n$, process noise vector $\mathbf{w}_k \in \mathfrak{X}^n$, measurement vector $\mathbf{z}_k \in \mathfrak{X}^m$, and measurement noise vector $\mathbf{v}_k \in \mathfrak{X}^m$. The vectors \mathbf{w}_k and \mathbf{v}_k are zero mean Gaussian white sequences having zero cross-correlation with each other:

$$\mathbf{E}[\mathbf{w}_k \mathbf{w}_i^T] = \begin{cases} \mathbf{Q}_k, & i = k \\ 0, & i \neq k \end{cases}; \quad \mathbf{E}[\mathbf{v}_k \mathbf{v}_i^T] = \begin{cases} \mathbf{R}_k, & i = k \\ 0, & i \neq k \end{cases}; \quad \mathbf{E}[\mathbf{w}_k \mathbf{v}_i^T] = \mathbf{0} \quad (2)$$

for all i and k

where \mathbf{Q}_k is the process noise covariance matrix, \mathbf{R}_k is the measurement noise covariance matrix.

2.1. Unscented Kalman Filter. The first step in the UKF involves the UT to generate a set of weighted samples/sigma points, which are deterministically chosen so that they adequately capture the true mean and covariance of the random variable. Consider an n dimensional random variable \mathbf{x} , which has the mean $\hat{\mathbf{x}}$ and covariance \mathbf{P} , and suppose that it propagates through an arbitrary nonlinear function \mathbf{f} . The UT creates $2n + 1$ sigma vectors \mathbf{x} (a capital letter) and weighted points W ,

$$\mathbf{X}_{(0)} = \hat{\mathbf{x}}; \quad \mathbf{X}_{(i)} = \hat{\mathbf{x}} + (\sqrt{(n + \lambda)\mathbf{P}})_i^T; \quad \mathbf{X}_{(i+n)} = \hat{\mathbf{x}} - (\sqrt{(n + \lambda)\mathbf{P}})_i^T, \quad i = 1, \dots, n \quad (3)$$

$$W_0^{(m)} = \frac{\lambda}{(n + \lambda)}; \quad W_0^{(c)} = W_0^{(m)} + (1 - \alpha^2 + \beta); \quad W_i^{(m)} = W_i^{(c)} = \frac{1}{2(n + \lambda)}, \quad (4)$$

$$i = 1, \dots, 2n$$

where $(\sqrt{(n + \lambda)\mathbf{P}})_i$ is the i th row of the matrix square root. $\sqrt{(n + \lambda)\mathbf{P}}$ can be obtained from the lower-triangular matrix of the Cholesky factorisation; $\lambda = \alpha^2(n + \kappa) - n$ is a scaling parameter; α determines the spread of the sigma points around $\hat{\mathbf{x}}$; κ is a secondary scaling parameter; β is used to incorporate prior knowledge of the distribution of $\bar{\mathbf{x}}$; $W_i^{(m)}$ and $W_i^{(c)}$ are the weights for the mean and covariance, respectively, associated with the i th point.

The sigma vectors are propagated through the nonlinear function to yield a set of transformed sigma points,

$$\mathbf{y}_i = f(\mathbf{X}_i) \quad i = 0, \dots, 2n \tag{5}$$

The mean and covariance of \mathbf{y}_i are approximated by a weighted average mean and covariance of the transformed sigma points:

$$\bar{\mathbf{y}}_u = \sum_{i=0}^{2n} W_i^{(m)} \mathbf{y}_i \tag{6}$$

$$\mathbf{P}_u = \sum_{i=0}^{2n} W_i^{(c)} (\mathbf{y}_i - \bar{\mathbf{y}}_u)(\mathbf{y}_i - \bar{\mathbf{y}}_u)^T \tag{7}$$

To implement the UKF, the set of sigma points are firstly created by Equation (3). After that, the algorithm for implementation of the UKF is summarized in Table 1 (Tseng et al., 2016). The UKF involves the time update (prediction) stage, i.e., Equations (8)–(12), and the measurement update (correction) stage, i.e., Equations (13)–(17).

2.2. *The Cubature Kalman Filter.* Like the UKF, the CKF is another type of nonlinear filtering approach without linearization of the nonlinear model. The first step of the CKF algorithm is to approximate the mean and variance of the probability distribution through a set of $2n$ (n is the dimension of the system model) cubature points. The cubature points are then propagated through the nonlinear functions. The mean and variance of the current approximate Gaussian distribution by the propagated cubature points can then be calculated. A set of $2n$ cubature points are given by $[\xi_i, \omega_i]$, where ξ_i is the i th cubature point and ω_i is the corresponding weight:

$$\xi_i = \begin{cases} \sqrt{n}[1]_i, & i = 1, 2, \dots, n \\ -\sqrt{n}[1]_{i-n}, & i = n + 1, n + 2, \dots, 2n \end{cases}; \quad \omega_i = \frac{1}{2n}, \quad i = 1, 2, \dots, 2n \tag{18}$$

where $[1]_i \in \mathbb{R}^n$ denotes the i th column vector of the identity matrix $\mathbf{I}_{n \times n}$. Under the assumption that the posterior density at time $k - 1$ is known, the steps involved in the time and measurement updates are given by Equations (19)–(32), summarised in Table 2. See Tseng et al. (2016) for further discussion.

3. THE HUBER-BASED REGRESSION FILTERING STRATEGY. As a recursive minimum ℓ_2 -norm estimation technique, the KF exhibits sensitivity to deviations from the assumed underlying probability distributions.

3.1. *The Huber M-estimate regression filtering.* The technique that relies on Huber’s generalised Maximum Likelihood (ML) methodology exhibits robustness against deviations from the commonly assumed Gaussian probability density functions and can solve the non-Gaussian distribution problem efficiently. The

Table 1. Implementation algorithm for the unscented Kalman filter.

- Initialisation: Initialise state vector $\hat{\mathbf{x}}_{0|0}$ and state covariance matrix $\mathbf{P}_{0|0}$

- Time update

(1) The transformed set is given by instantiating each point through the process model

$$\zeta_{i,k|k-1} = \mathbf{f}(\mathbf{X}_{i,k-1}); i = 0, \dots, 2n \quad (8)$$

(2) Predicted mean

$$\hat{\mathbf{x}}_{k|k-1} = \sum_{i=0}^{2n} W_i^{(m)} \zeta_{i,k|k-1} \quad (9)$$

(3) Predicted covariance

$$\mathbf{P}_{k|k-1} = \sum_{i=0}^{2n} W_i^{(c)} [\zeta_{i,k|k-1} - \hat{\mathbf{x}}_{k|k-1}] [\zeta_{i,k|k-1} - \hat{\mathbf{x}}_{k|k-1}]^T + \mathbf{Q}_{k-1} \quad (10)$$

(4) Instantiate each of the prediction points through observation model

$$\mathbf{Z}_{i,k|k-1} = \mathbf{h}(\zeta_{i,k|k-1}) \quad (11)$$

(5) Predicted observation

$$\hat{\mathbf{z}}_{k|k-1} = \sum_{i=0}^{2n} W_i^{(m)} \mathbf{Z}_{i,k|k-1} \quad (12)$$

- Measurement update

(6) Innovation covariance

$$\mathbf{P}_{zz} = \sum_{i=0}^{2n} W_i^{(c)} [\mathbf{Z}_{i,k|k-1} - \hat{\mathbf{z}}_{k|k-1}] [\mathbf{Z}_{i,k|k-1} - \hat{\mathbf{z}}_{k|k-1}]^T + \mathbf{R}_k \quad (13)$$

(7) Cross covariance

$$\mathbf{P}_{xz} = \sum_{i=0}^{2n} W_i^{(c)} [\zeta_{i,k|k-1} - \hat{\mathbf{x}}_{k|k-1}] [\mathbf{Z}_{i,k|k-1} - \hat{\mathbf{z}}_{k|k-1}]^T \quad (14)$$

(8) Performing update

$$\mathbf{K}_k = \mathbf{P}_{xz} \mathbf{P}_{zz}^{-1} \quad (15)$$

$$\hat{\mathbf{x}}_{k|k} = \hat{\mathbf{x}}_{k|k-1} + \mathbf{K}_k (\mathbf{z}_k - \hat{\mathbf{z}}_{k|k-1}) \quad (16)$$

$$\mathbf{P}_{k|k} = \mathbf{P}_{k|k-1} - \mathbf{K}_k \mathbf{P}_{zz} \mathbf{K}_k^T \quad (17)$$

Table 2. Implementation algorithm for the cubature Kalman filter.

- Initialisation: Initialise state vector $\hat{\mathbf{x}}_{0|0}$ and state covariance matrix $\mathbf{P}_{0|0}$

- Time update

(1) Factorise the covariance

$$\mathbf{P}_{k-1|k-1} = \mathbf{S}_{k-1|k-1} \mathbf{S}_{k-1|k-1}^T \tag{19}$$

(2) Evaluate the cubature points

$$\mathbf{X}_{i,k-1|k-1} = \mathbf{S}_{k-1|k-1} \boldsymbol{\xi}_i + \hat{\mathbf{x}}_{k-1|k-1} \tag{20}$$

(3) Evaluate the propagated cubature points through the process model

$$\mathbf{X}_{i,k|k-1}^* = \mathbf{f}(\mathbf{X}_{i,k-1|k-1}) \tag{21}$$

(4) Estimate the predicted mean

$$\hat{\mathbf{x}}_{k|k-1} = \sum_{i=1}^{2n} \omega_i \mathbf{X}_{i,k|k-1}^* \tag{22}$$

(5) Estimate the predicted error covariance

$$\mathbf{P}_{k|k-1} = \sum_{i=1}^{2n} \omega_i \mathbf{X}_{i,k|k-1}^* \mathbf{X}_{i,k|k-1}^{*T} - \hat{\mathbf{x}}_{k|k-1} \hat{\mathbf{x}}_{k|k-1}^T + \mathbf{Q}_{k-1} \tag{23}$$

- Measurement update

(6) Factorize the covariance

$$\mathbf{P}_{k|k-1} = \mathbf{S}_{k|k-1} \mathbf{S}_{k|k-1}^T \tag{24}$$

(7) Evaluate the cubature points

$$\mathbf{X}_{i,k|k-1} = \mathbf{S}_{k|k-1} \boldsymbol{\xi}_i + \hat{\mathbf{x}}_{k|k-1} \tag{25}$$

(8) Evaluate the propagated cubature points through observation model

$$\mathbf{Z}_{i,k|k-1} = \mathbf{h}(\mathbf{X}_{i,k|k-1}) \tag{26}$$

(9) Evaluate the propagated observation

$$\hat{\mathbf{z}}_{k|k-1} = \sum_{i=1}^{2n} \omega_i \mathbf{Z}_{i,k|k-1} \tag{27}$$

(10) Estimate the innovation covariance

$$\mathbf{P}_{zz} = \sum_{i=1}^{2n} \omega_i \mathbf{Z}_{i,k|k-1} \mathbf{Z}_{i,k|k-1}^T - \hat{\mathbf{z}}_{k|k-1} \hat{\mathbf{z}}_{k|k-1}^T + \mathbf{R}_k \tag{28}$$

Table 2. (Cont.)

- Initialisation: Initialise state vector $\hat{\mathbf{x}}_{0|0}$ and state covariance matrix $\mathbf{P}_{0|0}$

- Measurement update

(11) Estimate the cross-covariance

$$\mathbf{P}_{xz} = \sum_{i=1}^{2n} \omega_i \mathbf{X}_{i,k|k-1} \mathbf{Z}_{i,k|k-1}^T - \hat{\mathbf{x}}_{k|k-1} \hat{\mathbf{z}}_{k|k-1}^T \tag{29}$$

(12) Perform update state vector $\hat{\mathbf{x}}_{k|k}$ and its covariance matrix $\mathbf{P}_{k|k}$

$$\mathbf{K}_k = \mathbf{P}_{xz} \mathbf{P}_{zz}^{-1} \tag{30}$$

$$\hat{\mathbf{x}}_{k|k} = \hat{\mathbf{x}}_{k|k-1} + \mathbf{K}_k (\mathbf{z}_k - \hat{\mathbf{z}}_{k|k-1}) \tag{31}$$

$$\mathbf{P}_{k|k} = \mathbf{P}_{k|k-1} - \mathbf{K}_k \mathbf{P}_{zz} \mathbf{K}_k^T \tag{32}$$

measurement update can be modified to enhance robustness using the Huber M-estimation methodology.

To incorporate this mechanism into the CKF framework, it is necessary to reformulate the measurement update as a nonlinear regression problem between the state prediction and the observed quantity, and then the nonlinear regression model has the form (Chang et al., 2012a; Tsakonas et al., 2014)

$$\begin{bmatrix} \mathbf{z}_k \\ \hat{\mathbf{x}}_{k|k-1} \end{bmatrix} = \begin{bmatrix} \mathbf{h}(\mathbf{x}_k) \\ \mathbf{x}_k \end{bmatrix} + \begin{bmatrix} \mathbf{v}_k \\ \delta \mathbf{x}_k \end{bmatrix} \tag{33}$$

where \mathbf{z}_k is the measurement; $\hat{\mathbf{x}}_{k|k-1}$ is the predicted state, which is obtained from the time update of the CKF, and $\delta \mathbf{x}_k$ is the error between the true state and its prediction. Denote

$$\mathbf{e}_k = \begin{bmatrix} \mathbf{v}_k \\ \delta \mathbf{x}_k \end{bmatrix} \tag{34}$$

with

$$E[\mathbf{e}_k] = 0 \text{ and } E[\mathbf{e}_k \mathbf{e}_k^T] = \begin{bmatrix} \mathbf{R}_k & \mathbf{0} \\ \mathbf{0} & \mathbf{P}_{k|k-1} \end{bmatrix} = \mathbf{\Lambda}_k \mathbf{\Lambda}_k^T$$

where the term $\mathbf{\Lambda}_k$ may be obtained by Cholesky decomposition, \mathbf{R}_k is assumed to be a known noise covariance of \mathbf{v}_k , and $\mathbf{P}_{k|k-1}$ is the error covariance propagation.

Multiplying Equation (33) with $\mathbf{\Lambda}_k^{-1}$ yields the nonlinear regression model

$$\tilde{\mathbf{z}}_k = \mathbf{g}(\mathbf{x}_k) + \mathbf{v}_k \tag{35}$$

where

$$\tilde{\mathbf{z}}_k = \mathbf{\Lambda}_k^{-1} \begin{bmatrix} \mathbf{z}_k \\ \hat{\mathbf{x}}_{k|k-1} \end{bmatrix} \tag{36}$$

and

$$\mathbf{g}(\mathbf{x}_k) = \Lambda_k^{-1} \begin{bmatrix} \mathbf{h}(\mathbf{x}_k) \\ \mathbf{x}_k \end{bmatrix}; \mathbf{v}_k = \Lambda_k^{-1} \begin{bmatrix} \mathbf{v}_k \\ \delta \mathbf{x}_k \end{bmatrix} \tag{37}$$

The covariance matrix of \mathbf{v}_k is the identity matrix, as can be seen from expanding the expectation $E[\mathbf{v}_k \mathbf{v}_k^T]$.

In Equation (35), the nonlinear regression problem can be solved efficiently for robustness using Huber’s generalised ML method, in which the solution is found by minimising the cost function

$$\hat{\mathbf{x}}_{k|k} = \arg \min \sum_{i=1}^{m+n} \rho(v_{k,i}) \tag{38}$$

where $v_{k,i}$ is the i th component of the residual vector, \mathbf{v}_k , and $\rho(\cdot)$ is the score function used to handle the outlier measurements or contaminated Gaussian; m is the dimension of the measurement. The Huber ρ – function used in the M-estimation is given by

$$\rho(v_{k,i}) = \begin{cases} \frac{1}{2} v_{k,i}^2 & \text{for } |v_{k,i}| < \gamma \\ \gamma |v_{k,i}| - \frac{1}{2} \gamma^2 & \text{for } |v_{k,i}| \geq \gamma \end{cases} \tag{39}$$

where γ is a tuning parameter. The ρ function is a blend of the ℓ_1 – and ℓ_2 – norm estimation functions, which provides ℓ_2 – norm properties for small residuals to ensure both quality and efficiency simultaneously with normal distribution and ℓ_1 – norm properties for large residuals to suppress the impact of outliers or contaminated Gaussian noises.

If the ρ – function is differentiable, the solution to the nonlinear regression problem can be found by deriving the cost function as

$$\sum_{i=1}^{m+n} \phi(v_{k,i}) \frac{\partial v_{k,i}}{\partial \mathbf{x}_k} = 0 \tag{40}$$

where $\phi(v_{k,i}) = \rho'(v_{k,i})$. By defining the ψ – function as $\psi(v_{k,i}) = \phi(v_{k,i})/v_{k,i}$, and corresponding to the Huber density

$$\psi(v_{k,i}) = \begin{cases} 1 & \text{for } |v_{k,i}| \leq \gamma \\ \gamma \text{sgn}(v_{k,i})/v_{k,i} & \text{for } |v_{k,i}| > \gamma \end{cases} \tag{41}$$

The Huber M-estimation methodology makes use of the weighting matrix $\Psi_k = \text{diag}[\psi(v_{k,i})]$ to recast the measurement information. The weighting matrix Ψ_k can be interpreted in terms of re-weighting the residual error covariance and constructing the measurement process. The modified covariance is given by

$$\bar{\mathbf{S}}_k = \Lambda_k^{1/2} \Psi_k^{-1} \Lambda_k^{T/2} \tag{42}$$

The decomposition of the $\bar{\mathbf{S}}_k$ matrix into two submatrices $\bar{\mathbf{S}}_{k,x}$ and $\bar{\mathbf{S}}_{k,z}$ corresponding to the state update and state correction with measurements so that

$$\bar{\mathbf{S}}_k = \begin{bmatrix} \bar{\mathbf{S}}_{k,z} & 0 \\ 0 & \bar{\mathbf{S}}_{k,x} \end{bmatrix} \tag{43}$$

where the $\bar{\mathbf{S}}_{k,x} \in \mathfrak{R}^{n \times n}$ and $\bar{\mathbf{S}}_{k,z} \in \mathfrak{R}^{m \times m}$, and the robust measure of correlation can be expressed in the recursive filter as the CKF measurement-update.

3.2. *The Huber M-estimation methodology-based cubature Kalman filter.* The Huber M-estimation methodology is robust to the outliers and ensures the accuracy of the state estimation. To apply this method, the ψ -function is the component of the residual vector \mathbf{v}_k . Specifically, when $|\mathbf{v}_{k,i}| \leq \gamma$, the accurate estimation of the system state reaches the expected behaviour demand and Huber-based filtering is equal to the CKF measurement update because $\Psi_k = \mathbf{I}$. On the other hand, when $|\mathbf{v}_{k,i}| > \gamma$, the CKF has a divergent trend due to the noisy signal contaminated with outliers. The measurement update is modified to exhibit robustness using the Huber M-estimation methodology. Defining

$$\tilde{\mathbf{R}}_k = \bar{\mathbf{S}}_{k,z} \quad (44)$$

as the modified measurement covariance, and replacing \mathbf{R}_k with $\tilde{\mathbf{R}}_k$ for operating the CKF framework in the measurement update process leads to the HCKF.

The M-estimation problem in Huber-based regression filtering aims to minimise Equation (38), which is essentially nonlinear. In general, the calculation procedure of the RKF based on the M-estimation should be solved iteratively because the reweight matrix is constructed based on the residuals which are further related to the state parameter estimates. By setting the weighting matrix $\Psi_k = \text{diag}[\psi(\mathbf{v}_{k,i})]$, an iterative solution to Equation (40) can be obtained using statistical approximation of the nonlinear function $\mathbf{M}_k(\cdot)$. The implicit Equation (40) can be written in matrix form,

$$\mathbf{M}_k^T \Psi_k (\mathbf{M}_k \mathbf{x}_k - \tilde{\mathbf{z}}_k) = 0 \quad (45)$$

which can be solved for \mathbf{x}_k to give $\mathbf{x}_k = (\mathbf{M}_k^T \Psi_k \mathbf{M}_k)^{-1} \mathbf{M}_k^T \Psi_k \tilde{\mathbf{z}}_k$. An iterative solution to Equation (45) is expressed as

$$\mathbf{x}_k^{(j+1)} = (\mathbf{M}_k^T \Psi_k^{(j)} \mathbf{M}_k)^{-1} \mathbf{M}_k^T \Psi_k^{(j)} \tilde{\mathbf{z}}_k \quad (46)$$

where the superscript (j) is the iteration index. The method can be initialised using the iteratively reweighted least squares solution: $\mathbf{x}_k^{(0)} = (\mathbf{M}_k^T \mathbf{M}_k)^{-1} \mathbf{M}_k^T \tilde{\mathbf{z}}_k$. The iteration will converge if the ψ function is non-increasing when using the Huber ρ function. The Huber M-estimation methodology can be iterated until convergence or can be carried out through only one fixed iteration step to save the computational costs.

The one-step Huber update is summarised as follows: (1) Calculate $\mathbf{X}_{i,k|k-1}$, $\mathbf{Z}_{i,k|k-1}$ and $\hat{\mathbf{z}}_{k|k-1}$ using Equations (24)–(27); (2) Form a redundant observation vector shown as in Equation (33); (3) Solve the nonlinear regression model, Equation (35), using the solution based on Equation (46); (4) Update the modified covariance matrix using Equation (42); (5) Perform the modified measurement covariance $\tilde{\mathbf{R}}_k = \bar{\mathbf{S}}_{k,z}$, and replace \mathbf{R}_k with $\tilde{\mathbf{R}}_k$ for operating the CKF framework in Equation (28). A high level operation of the HCKF is shown in Figure 1(a). Figure 1(b) presents the architecture of the GPS navigation processing using the HCKF implementation. The block indicated by the dashed-line is the Huber M-estimation mechanism for reformulating the measurement information of the CKF framework.

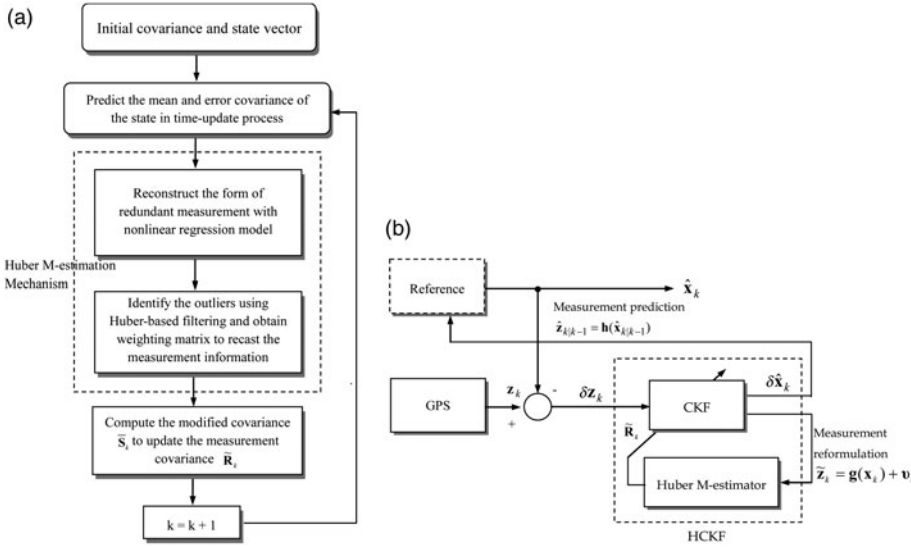


Figure 1. (a) High level operation of the HCKF. (b) Configuration of the GPS navigation processing using the HCKF.

4. GPS NAVIGATION PROCESSING AND PERFORMANCE EVALUATION. Both the Position-Velocity (PV) model and the Position-Velocity-Acceleration (PVA) model are linear. The nonlinear dynamic model that provides more adequate description of the vehicle is adopted from Jwo and Lai (2008).

4.1. *Nonlinear dynamic modelling for GPS navigation.* Consider a vehicle moving at the velocity represented as $V_b = u_b \vec{i} + v_b \vec{j} + w_b \vec{k}$. The velocity in the fixed frame in terms of Euler angles and body velocity components has the relation

$$V_b = \begin{bmatrix} \dot{x} \\ \dot{y} \\ \dot{z} \end{bmatrix} = \begin{bmatrix} C_\theta C_\psi & S_\Phi S_\theta C_\psi - C_\Phi S_\psi & C_\Phi S_\theta C_\psi + S_\Phi S_\psi \\ C_\theta S_\psi & S_\Phi S_\theta S_\psi + C_\Phi C_\psi & C_\Phi S_\theta S_\psi - S_\Phi C_\psi \\ -S_\theta & S_\Phi C_\theta & C_\Phi C_\theta \end{bmatrix} \begin{bmatrix} u_b \\ v_b \\ w_b \end{bmatrix}$$

where the following notations are used: $S_\Phi \equiv \sin(\Phi)$, $C_\Phi \equiv \cos(\Phi)$, $S_\theta \equiv \sin(\theta)$, $S_\psi \equiv \sin(\psi)$, $S_\psi \equiv \sin(\psi)$, and $C_\psi \equiv \cos(\psi)$. Suppose that the non-holonomic constraint is applied, only the longitudinal movement is considered and the lateral slippage is neglected. In case the velocity in the x-component of the body frame is considered, $\|V_b\| \approx \|u_b \vec{i}\| \approx V$, the dynamic process model of the GPS receiver, when written in the form $\dot{x} = f(x, t) + u(t)$, can be represented as

$$\begin{bmatrix} \dot{x} \\ \dot{y} \\ \dot{z} \\ \dot{\Phi} \\ \dot{\theta} \\ \dot{\psi} \\ \dot{V} \\ \dot{b} \\ \dot{d} \end{bmatrix} = \begin{bmatrix} \dot{x}_1 \\ \dot{x}_2 \\ \dot{x}_3 \\ \dot{x}_4 \\ \dot{x}_5 \\ \dot{x}_6 \\ \dot{x}_7 \\ \dot{x}_8 \\ \dot{x}_9 \end{bmatrix} = \begin{bmatrix} V \cos \theta \cos \psi \\ V \cos \theta \sin \psi \\ -V \sin \theta \\ 0 \\ 0 \\ 0 \\ 0 \\ d \\ 0 \end{bmatrix} + \begin{bmatrix} u_1 \\ u_2 \\ u_3 \\ u_4 \\ u_5 \\ u_6 \\ u_7 \\ u_8 \\ u_9 \end{bmatrix} \tag{47}$$

The process noise covariance matrix for this nonlinear model has the form

$$\mathbf{Q}_k = \begin{bmatrix} \text{diag}(q_{11}, \dots, q_{77}) & 0 \\ 0 & \begin{matrix} q_{88} & q_{98} \\ q_{89} & q_{99} \end{matrix} \end{bmatrix} \tag{48}$$

where $\text{diag}(q_{11}, \dots, q_{77})$ denotes the 7×7 diagonal matrix, and $q_{88} = S_f \Delta t + S_g \Delta t^3 / 2$; $q_{98} = S_g \Delta t^2 / 2$; $q_{89} = S_g \Delta t^2 / 2$; $q_{99} = S_g \Delta t$.

If only the pseudorange observables are available, the expected pseudorange $\mathbf{h}_k(\hat{\mathbf{x}}_k^-)$ based on the GPS satellite position and the *a priori* estimate $\hat{\mathbf{x}}_k^-$ is given by

$$\mathbf{h}_k(\hat{\mathbf{x}}_k^-) = \sqrt{(\hat{x}_k^- - x_i)^2 + (\hat{y}_k^- - y_i)^2 + (\hat{z}_k^- - z_i)^2} \tag{49}$$

The elements of the measurement model \mathbf{H}_k are the partial derivatives of the predicted measurements with respect to each state, which is an $n \times 9$ matrix:

$$\mathbf{H}_k = \frac{\partial \mathbf{h}_k}{\partial \mathbf{x}_k} = \begin{bmatrix} e_x^{(1)} & e_y^{(1)} & e_z^{(1)} & 0 & 0 & 0 & 0 & 1 & 0 \\ e_x^{(2)} & e_y^{(2)} & e_z^{(2)} & 0 & 0 & 0 & 0 & 1 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots & \vdots & \vdots & 0 \\ e_x^{(3)} & e_y^{(3)} & e_z^{(3)} & 0 & 0 & 0 & 0 & 1 & 0 \end{bmatrix}$$

where the line-of-sight vector from the user to the satellites is $(e_x^{(i)}, e_y^{(i)}, e_z^{(i)})$

$$e_x^{(i)} = \frac{\hat{x}_k^- - x_i}{\hat{r}_i}; \quad e_y^{(i)} = \frac{\hat{y}_k^- - y_i}{\hat{r}_i}; \quad e_z^{(i)} = \frac{\hat{z}_k^- - z_i}{\hat{r}_i}$$

4.2. *Simulation test.* The commercial software Satellite Navigation (SATNAV) Toolbox by GPSof LLC was employed for generating the satellite positions and pseudoranges. The scenario for simulation used in this paper is adopted from Jwo and Lai (2008). The three dimensional vehicle trajectory and the vehicle velocity in the east, north, and altitude components is shown in Figure 2. The description of the vehicle motion is provided in Table 3. The dynamic process model used in the EKF is the PV model, while the nonlinear model has been applied to the other approaches. The spectral amplitudes for the clock model, $S_f = 10^{-2} \text{ sec}$ and $S_g = 10^{-2} \text{ sec}^{-1}$, were used to find the \mathbf{Q}_k matrix as in Equation (48) for the nonlinear dynamic model.

It is assumed that the Differential GPS (DGPS) mode is used, and most of the receiver-dependent errors are removed. We assume the remaining errors, mostly caused by multipath errors and thermal noise, follow a Gaussian mixture distribution of the form, defined by the PDF of the form

$$f_{\mathbf{X}}(x) = \left(\frac{1 - \varepsilon}{\sigma_1 \sqrt{2\pi}} \right) \exp \left[- \left(\frac{x^2}{2\sigma_1^2} \right) \right] + \left(\frac{\varepsilon}{\sigma_2 \sqrt{2\pi}} \right) \exp \left[- \left(\frac{x^2}{2\sigma_2^2} \right) \right] \tag{50}$$

or equivalently by $\mathbf{X} \sim (1 - \varepsilon)N(0, \sigma_1^2) + \varepsilon N(0, \sigma_2^2)$, where σ_1 and σ_2 are the standard deviations of the individual Gaussian distributions, and ε is a perturbing parameter that represents error model contamination. In order to demonstrate the validity of the robustness of the proposed algorithm in non-Gaussian distributions, the perturbed conditions $\varepsilon = 0.6$ (i.e., with 60% contamination), and $\sigma_2 = 10\sigma_1 = 10 \times 1$ (m) have been used. Figure 3 shows the Gaussian mixture model for the pseudorange errors in the

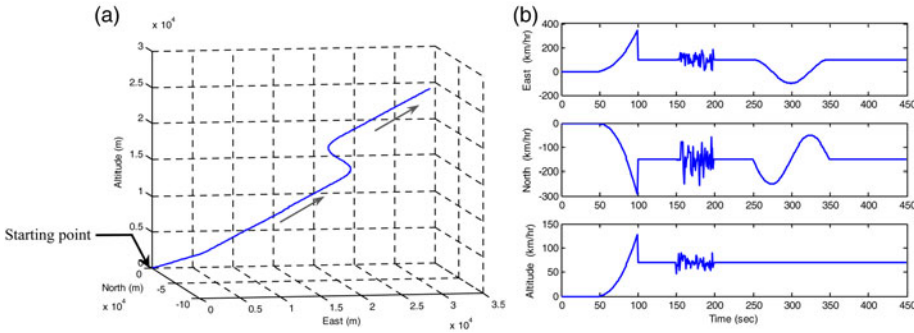


Figure 2. The scenario for simulation: (a) three dimensional trajectory of the vehicle and (b) the east, north, and altitude components of the vehicle velocity.

Table 3. Description of the vehicle motion.

Time interval (sec)	Motion
[0–50] [101–150] [201–250] [351–450]	Constant velocity
[51–100]	Constant acceleration
[151–200]	Variable acceleration
[251–350]	Circular motion, clockwise turn

simulation. The initial measurement noise variances r_{ρ_i} value are set to be 8 m^2 . The parameters utilised are $\alpha = 1.4$, $\beta = 2.5$, $\kappa = 0$ in the UKF; $\gamma = 1.345$ in the CKF, for which the same value of the bounded-influence estimation selected by Huber (1981) is used.

The results are shown in Figures 4–6. Prior to the incorporation of the Huber M-estimation mechanism, the GPS navigation accuracies based on the conventional nonlinear approaches: EKF, UKF and CKF are presented, shown in Figure 4, where the merits of the CKF have been shown. After that, evaluation of the results based on the nonlinear approaches compared to the Huber-based M-estimation CKF approach is conducted. Figure 5 provides comparison for the three types of filters: UKF, CKF, and HCKF. To show the benefit of the Huber M-estimation methodology, the comparison among Huber-based EKF, Huber-based UKF and Huber-based CKF has also be presented, shown in Figure 6. The HCKF, which inherits the virtue of robustness, shows the best performance among the various approaches. The simulation results show the superiority of the HCKF in highly dynamic situations which lead to the model to be mismatched to the actual situation. The improvement in robustness for the non-Gaussian measurement errors has also been presented.

Several important remarks are given as follows.

- (1) In the four time intervals, 0–50, 101–150, 201–250, 350–450 s, the vehicle does not conduct manoeuvring and is performing constant-velocity straight-line motion. The navigation errors of all the nonlinear filters are not significantly distinguishable between the six methods.
- (2) In the three time intervals, 51–100, 151–200, 251–350 s, the vehicle is conducting manoeuvring (constant acceleration, variable acceleration, or circular motion).

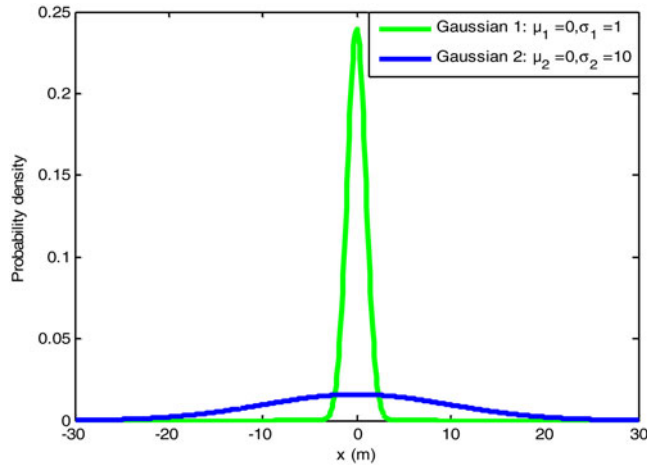


Figure 3 The Gaussian mixture model used in the simulation of pseudorange errors.

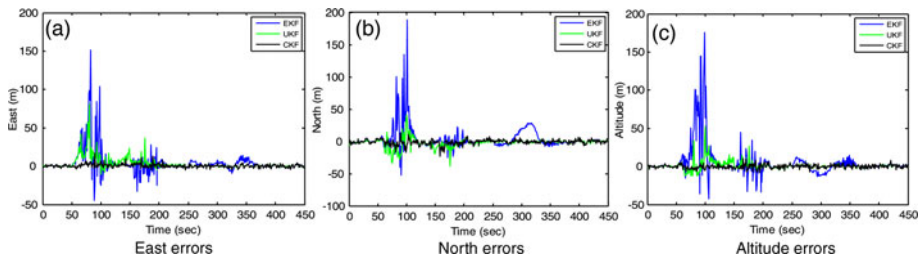


Figure 4. Comparison of positioning errors for EKF, UKF and CKF.

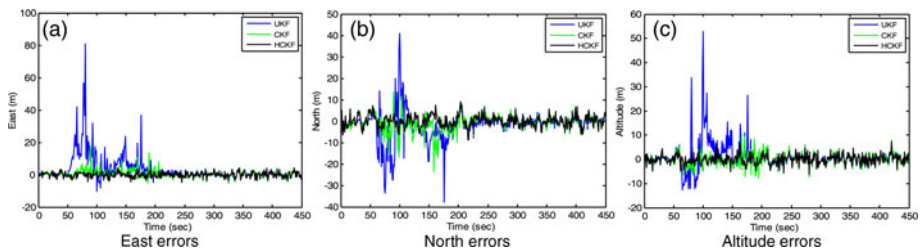


Figure 5. Comparison of positioning errors for UKF, CKF and HCKF.

The model mismatch to the actual situation leads to increased errors, as can be seen from the solution using the standard EKF. The CKF is about to converge in the high dynamic regions where there still exist noticeably large errors for the UKF-based solutions. However, performances based on the nonlinear filters are sensitive when the actual noise distribution deviates from the assumed Gaussian model. To further improve the performance of the CKF, a robust technique where the Huber M-estimation mechanism is incorporated can be utilised.

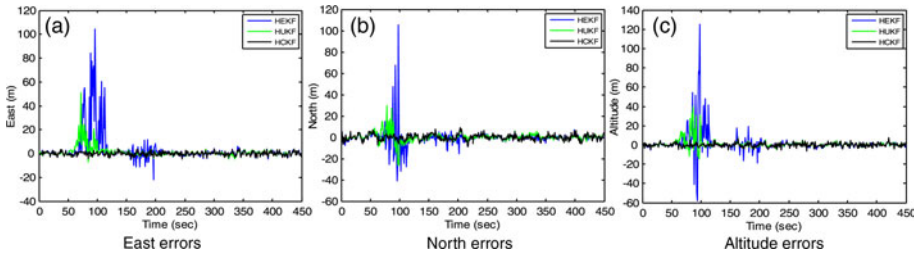


Figure 6. Comparison of positioning errors for HEKF, HUKF and HCKF.

- (3) When the GPS pseudorange observables are contaminated with non-Gaussian measurement errors, the CKF can adequately capture the non-Gaussianity and demonstrate noticeably better performance. Due to the appropriate tuning via the Huber M-estimation methodology, the HCKF exhibits robust behaviour and therefore outperforms the CKF when the non-Gaussian measurement errors are involved.
- (4) The HCKF associates the contaminated measurements with an increase in the measurement covariance, causing the reformulated error covariance. This modified covariance inflation is known to cause an increase in the state estimation error due to the fact that all measurements are processed as if they were outliers.
- (5) The HCKF and CKF filters give comparable results to each other for the positioning error during the non-maneuvring intervals, while the HCKF provides a substantially improved performance in the high dynamic regions (51–100, 151–200, and 251–350 s), where it adequately captures both nonlinearity and non-Gaussianity. The HCKF is expected to provide further improved performance since the residual error covariance and measurement information used in computing the Huber M-estimation is the ML estimates for perfectly compensating against deviations from Gaussianity. The navigation results show that the HCKF demonstrates superior performance as compared to the other approaches.

4.3. *Outdoor tests.* Two tests conducted outdoors were carried out to further validate the effectiveness of the proposed algorithm. The two tests conducted at the National Taiwan Ocean University (NTOU) campus involved a field test and a rooftop experiment. The data collected by the Ashtech GPS receiver (DG14 module) was used for processing to perform a comparative study on navigation performance using various nonlinear estimators. The sampling rate of the measurements is 1 Hz. The measurement noise variances r_{p_i} value are set to be 100 m^2 . The test results were obtained by post-processing of logged sensor data using the Matlab[®] software. The high accuracy relative positioning solution was used as the reference trajectory.

4.3.1. *Field test.* The first dataset was collected on the sports field at the university campus, shown in Figure 7. The raw data was collected for 140 seconds. In this test, the number of visible satellites varied between eight and nine in the clear open sky. To investigate the effect due to the contaminated non-Gaussian measurement errors, additional errors in the GPS pseudorange observables are added. The errors are assumed to possess the distribution again in the form as in Figure 3. Figure 8 shows the positioning errors for EKF, UKF (b) and CKF. The CKF framework possesses some



Figure 7. The trajectory for the field test in Google map.

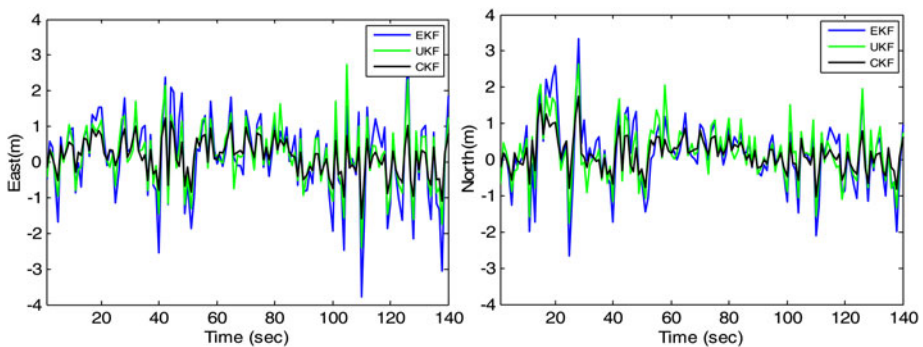


Figure 8. Positioning errors for EKF, UKF and CKF – the field test.

capability to deal with the non-Gaussian measurement errors/outliers. A comparison of the positioning errors for the three types of filters: UKF, CKF, and HCKF, is shown in Figure 9.

4.3.2. *Rooftop experiment.* The second dataset selected for performance evaluation was collected on the rooftop of the department building. The trajectory for the rooftop experiment is shown in Figure 10, where a close look at the trajectory is also provided using the data collected by the Ashtech GPS receiver. The picture shown inside the trajectory based on the Ashtech GPS receiver is a zoomed-in shot for the upper right corner where a red tower is located. The starting point of the trajectory is at one side of the red tower, and the terminated point at the other side. In this experiment, the raw data was collected for 124 seconds. The rooftop environment presented a negative impact on the quality of GPS signals in such a challenging multipath environment. In this environment, there are several reflective objects on the rooftop,

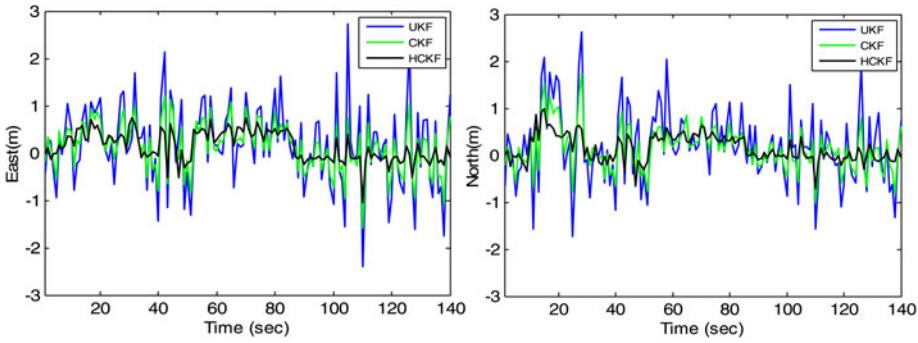


Figure 9. Positioning errors for UKF, CKF and HCKF– the field test.

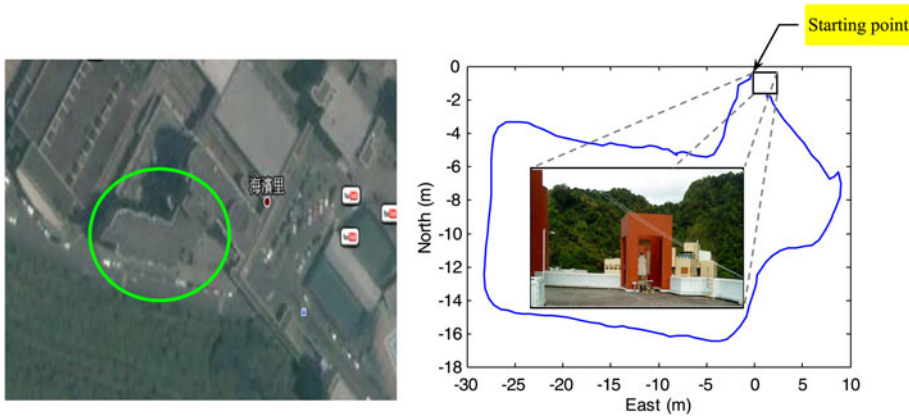


Figure 10. The trajectory for the rooftop experiment by Google map (left, circled area) and by the data collected by the Ashtech GPS receiver (right), respectively.

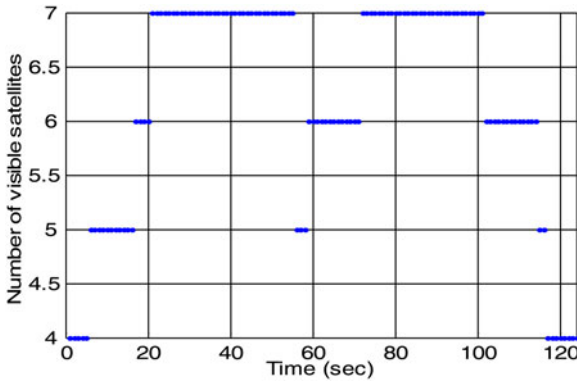


Figure 11. Number of visible satellites for the rooftop experiment.

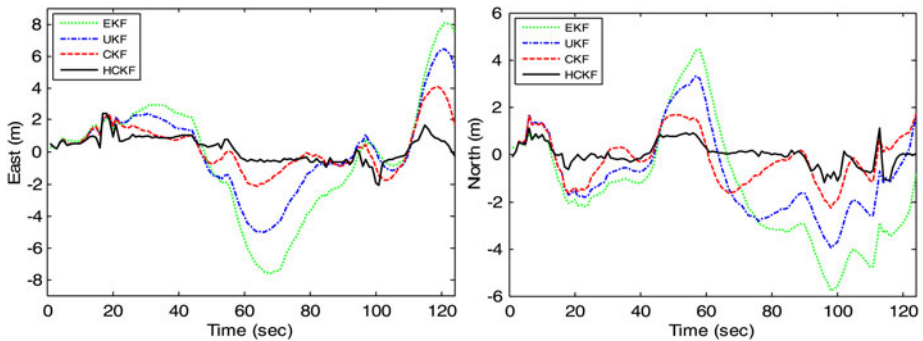


Figure 12. Positioning errors for EKF, UKF, CKF and HCKF—the rooftop experiment.

including the door that provides access to the roof. In addition, there are several other antennae in the vicinity. Multipath reflections from rooftop objects introduced a significant number of non-Gaussian errors to the GPS pseudorange observables. Furthermore, some of the GPS signals were blocked by the mountains in the south such that the number of visible satellites varied between four and seven, as shown in Figure 11. Figure 12 presents the positioning errors for the four approaches: EKF, UKF, CKF and HCKF.

The results for the two outdoor tests do not try to demonstrate the superiority of the HCKF in high dynamic environments, but exhibit effectiveness against the non-Gaussian measurement errors. In the above three illustrative examples, the improvement due to the incorporation of the Huber M-estimation mechanism has been demonstrated. The GPS positioning errors clearly show the superiority of the HCKF algorithm.

5. CONCLUSIONS. This paper has presented a Huber-based Cubature Kalman Filter (HCKF) for GPS navigation processing. The measurements are GPS pseudorange observables contaminated with non-Gaussian errors. The cubature Kalman filter is based on the third-degree spherical-radial cubature rule to propagate the cubature points through the nonlinear functions, so as to solve the integration in the Bayesian filtering problem for numerically computing the multivariate moment integrals. The Huber M-estimation methodology is essentially based on Huber's generalised ML estimation theory, which exhibits robustness against the deviations from the Gaussian distribution.

The HCKF algorithm integrates the merits of the Huber M-estimation methodology to handle non-Gaussian errors or outliers and the cubature Kalman filter to avoid the numerical instability problem and improve the estimation accuracy in high-dimensional systems. The CKF framework was recast in the form of a sequence of nonlinear regression problems to be solved, at each measurement update using the robust Huber M-estimation methodology so as to enhance the CKF performance. In addition to the simulation test, the real GPS navigation data is also applied to validate the proposed algorithm. Performance comparisons on the six types of filters, EKF, UKF, CKF, HEKF, HUKF, and HCKF have been presented, among which the HCKF algorithm shows superior performance in GPS navigational accuracy.

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REFERENCES

- Axelrad, P. and Brown, R.G. (1996). *GPS navigation algorithms*, in: Parkinson, B.W., Spilker, J.J., Axelrad, P. and Enga, P. (Ed.), *Global Positioning System: Theory and Applications*, Volume I, AIAA, Washington DC, Chap. 9.
- Arasaratnam, I. and Haykin, S. (2009). Cubature Kalman filters, *IEEE Transactions on Automatic Control*, **54**(6), 1254–1269.
- Boncellet, C.G. and Dickinson, B.W. (1983). *An approach to robust Kalman filtering*. in: The 22nd IEEE Conference on Decision and Control, San Antonio, USA, 304–305.
- Brown, R.G. and Hwang, P.Y.C. (1997). *Introduction to Random Signals and Applied Kalman Filtering*. John Wiley & Sons, New York.
- Bube, K.P. and Langan, R.T. (1997). Hybrid ℓ_1/ℓ_2 minimization with applications to tomography, *Geophysics*, **62**(4), 1183–1195.
- Chang, L., Hu, B., Chang, G. and Li, A. (2012a). Huber-based novel robust unscented Kalman filter, *IET Science, Measurement & Technology*, **6**(6), 502–509.
- Chang, L., Hu, B., Chang, G. and Li, A. (2012b). Marginalised iterated unscented Kalman filter, *IET Control Theory & Applications*, **6**, 847–854.
- Chang, L., Hu, B., Chang, G. and Li, A. (2012c). Multiple outliers suppression derivative-free filter based on unscented transformation, *Journal of guidance, control, and dynamics*, **35**, 1902–1907.
- Chang, G. and Liu, M. (2015). M-estimator-based robust Kalman filter for systems with process modeling errors and rank deficient measurement models, *Nonlinear Dynamics*, **80**(3), 1431–1449.
- Durgaprasad, G. and Thakur, S.S. (1998). Robust dynamic state estimation of power systems based on M-estimation and realistic modeling of system dynamics, *IEEE Transactions on Power Systems*, **13**(4), 1331–1336.
- Fernandez-Prades, C. and Vila-Valls, J. (2010). Bayesian nonlinear filtering using quadrature and cubature rules applied to sensor data fusion for positioning, in: *Proceeding of IEEE international conference on communications*, 1–5.
- Gandhi, M.A. and Mili, L. (2010). Robust Kalman filter based on a generalized maximum-likelihood-type estimator, *IEEE Transactions on Signal Processing*, **58**(5), 2509–2520.
- Gao, W., Liu, Y. and Xu, B. (2014). Robust Huber-based iterated divided difference filtering with application to cooperative localization of autonomous underwater vehicles, *Sensors*, **14**(12), 24523–24542.
- Gelb, A. (1974). *Applied Optimal Estimation*. M. I. T. Press, MA.
- Guitton, A. and Symes, W.W. (2003). Robust inversion of seismic data using the Huber norm, *Geophysics*, **68**(4), 1310–1319.
- Gustafsson, F. and Hendeby, G. (2012). Some relations between extended and unscented Kalman filters. *IEEE Transactions on Signal Processing*, **60**, 545–555.
- Hou, Y. (2014). Multiple-sensor fusion tracking based on square-root cubature Kalman filtering, *Journal of networks*, **9**(7), 1955–1961.
- Huber, P.J. (1973). Robust regression: asymptotics, conjectures and Monte Carlo, *Annals of Statistics*, **1**, 799–821.
- Huber, P.J. (1981). *Robust statistics*. Wiley, New York.
- Jia, B., Xin, M. and Cheng, Y. (2013). Higher-degree cubature Kalman filter, *Automatica*, **49**(2), 510–518.
- Julier, S.J. (2002). The scaled unscented transformation, in: *Proceedings of the American Control Conference*, Anchorage, USA, 4555–4559.
- Julier, S.J., Uhlmann, J.K. and Durrant-Whyte, H.F. (1995). A new approach for filtering nonlinear system, in: *Proceeding of the American Control Conference*, 1628–1632.
- Julier, S.J., Uhlmann, J.K. and Durrant-Whyte, H.F. (2000). A new method for the nonlinear transformation of means and covariances in filters and estimators, *IEEE Transactions on Automatic Control*, **5**(3), 477–482.
- Jwo, D. -J. and Lai, C. -N. (2008). Unscented Kalman filter with nonlinear dynamic process modeling for GPS navigation, *GPS Solutions*, **12**(4), 249–260.

- Karlggaard, C.D. (2015). Nonlinear regression Huber–Kalman filtering and fixed-interval smoothing. *Journal of Guidance, Control and Dynamics*, **38**, 322–330.
- Karlggaard, C.D., Schaub, H., Crassidis, J.L., Psiaki, M., Chiang, R., Ohlmeyer, E.J. and Norgaard, M. (2007). Huber-based divided difference filtering, *Journal of guidance, control, and dynamics*, **30**, 885–891.
- Kovacevic, B., Durovic, Z. and Glavaski, S. (1992). On robust Kalman filtering, *International Journal of Control*, **56**(3), 547–562.
- Maronna, R.A., Martin, R.D. and Yohai, V.J. (2006). *Robust statistics: theory and methods*. Wiley, West Sussex, England.
- Masreliez, C. and Martin, R. (1977). Robust Bayesian estimation for the linear model and robustifying the Kalman filter, *IEEE Transactions on Automatic Control*, **22**, 361–371.
- Petrus, P. (1999). Robust Huber adaptive filter, *IEEE Transactions on Signal Processing*, **47**(4), 1129–1133.
- Tsakonas, E., Jaldén, J., Sidiropoulos, N.D. and Ottersten, B. (2014). Convergence of the Huber regression M-estimate in the presence of dense outliers, *IEEE Signal Processing Letters*, **21**(10), 1211–1214.
- Tseng, C. -H, Lin, S. -F. and Jwo, D.J. (2016). Fuzzy Adaptive Cubature Kalman Filter for Integrated Navigation Systems, *Sensors*, **16**(8), 1167.
- Van der Merwe, R. and Wan, E.A. (2004). Sigma-point Kalman filters for nonlinear estimation and sensor-fusion: applications to integrated navigation, in: *Proceedings of the AIAA Guidance, Navigation & Control Conference and Exhibit*, 16–19.
- Wan, E.A. and Van der Merwe, R. (2001). The Unscented Kalman Filter, in: Simon Haykin (Ed.), *Kalman Filtering and Neural Networks*, Wiley, (Chapter 7).
- Wang, X., Cui, N. and Guo, J. (2010). Huber-based unscented filtering and its application to vision-based relative navigation, *IET Radar, Sonar & Navigation*, **4**(1), 134–141.
- Wu, Y., Hu, D., Wu, M. and Hu, X. (2006). A numerical-integration perspective on Gaussian filters, *IEEE Transactions on Signal Processing*, **54**, 2910–2921.
- Xu, B., Xiao, Y.P., Gao, W., Zhang, Y.G., Liu, Y.L. and Liu, Y. (2014). Dual-model reverse CKF algorithm in cooperative navigation for USV, *Mathematical Problems in Engineering*, **2014**, 1–17.
- Yang, Y. (1991). Robust Bayesian estimation, *Bulletin Géodésique*, **65**, 145–150.
- Yang, Y. and Gao, W. (2006a). An optimal adaptive Kalman filter, *Journal of Geodesy*, **80**, 177–183.
- Yang, Y. and Gao, W. (2006b). A new learning statistic for adaptive filter based on predicted residuals, *Progress in Natural Science*, **16**(8), 833–837.
- Yang, Y., He, H. and Xu, G. (2001). Adaptively robust filtering for kinematic geodetic positioning, *Journal of Geodesy*, **75**, 109–116.