Deciphering the concurrent phenomenon of childhood malnutrition by using the extended composite index of anthropometric failure (ECIAF): facts from the BESLEN project

Gözde DUMLU BİLGİN^{1*}, İrem KAYA CEBİOĞLU², Hasan Kaan KAVSARA^{3,4}, Aybüke SARIOĞLU⁵, Melis KEKÜLLÜOĞLU TAN⁶, Sema AYDIN⁷, Pınar USTA⁸, Binnur Okan BAKIR⁹

¹PhD, Assistant Professor, Yeditepe University, Faculty of Health Sciences, Department of Nutrition and Dietetics, İnönü Mah. Kayışdağı Cad. 34755 İstanbul, Turkey, ORCID: 0000-0001-6385-0190, P. +902165780560 F. +902165780496

²PhD, Associate Professor, Yeditepe University, Faculty of Health Sciences, Department of Nutrition and Dietetics, İnönü Mah. Kayışdağı Cad. 34755 İstanbul, Turkey, ORCID ID: 0000-0002-0135-2011, irem.cebioglu@yeditepe.edu.tr

³Yeditepe University, Faculty of Health Sciences, Department of Nutrition and Dietetics, İnönü Mah. Kayışdağı Cad. 34755 İstanbul, Turkey

⁴Istanbul Medipol University, Graduate School of Health Sciences, Nutrition and Dietetics, Istanbul, Türkiye. ORCID: 0000-0002-0322-0397, <u>kaan.kavsara@yeditepe.edu.tr</u>

⁵PhD candidate, Research Assistant, Yeditepe University, Faculty of Health Sciences, Department of Nutrition and Dietetics, İnönü Mah. Kayışdağı Cad. 34755 İstanbul, Turkey, ORCID: 0000-0002-0700-3744, <u>aybuke.sarioglu@yeditepe.edu.tr</u>



This is an Accepted Manuscript for Public Health Nutrition. This peer-reviewed article has been accepted for publication but not yet copyedited or typeset, and so may be subject to change during the production process. The article is considered published and may be cited using its

DOI 10.1017/S1368980024002520

Public Health Nutrition is published by Cambridge University Press on behalf of The Nutrition Society. This is an Open Access article, distributed under the terms of the Creative Commons Attribution licence (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted re-use, distribution and reproduction, provided the original article is properly cited.

⁶PhD candidate, Research Assistant, Yeditepe University, Faculty of Health Sciences, Department of Nutrition and Dietetics, İnönü Mah. Kayışdağı Cad. 34755 İstanbul, Turkey, ORCID: 0000-0002-2007-650X, melis.kekulluoglu@yeditepe.edu.tr

⁷PhD candidate, Research Assistant, Yeditepe University, Faculty of Health Sciences, Department of Nutrition and Dietetics, İnönü Mah. Kayışdağı Cad. 34755 İstanbul, Turkey, ORCID: 0000-0003-4262-7007, <u>sema.aydin@yeditepe.edu.tr</u>

⁸MSc, PhD student, Research Assistant, Yeditepe University, Faculty of Health Sciences, Department of Nutrition and Dietetics, İnönü Mah. Kayışdağı Cad. 34755 İstanbul, Turkey, ORCID: 0000-0001-8760-0233, <u>pinar.usta@yeditepe.edu.tr</u>

⁹PhD, Associate Professor, Yeditepe University, Faculty of Health Sciences, Department of Nutrition and Dietetics, İnönü Mah. Kayışdağı Cad. 34755 İstanbul, Turkey, ORCID ID: 0000-0002-0448-4300, <u>binnur.bakir@yeditepe.edu.tr</u>

*Corresponding Author: Gözde DUMLU BİLGİN, E-mail: gozde.dumlu@yeditepe.edu.tr

Running title: Childhood malnutrition using ECIAF: BESLEN project

Acknowledgements: We are grateful to Pendik Municipality Directorate of Cultural Affairs for supporting the study to be carried out, for their cooperation at each step, and to Yeditepe University Department of Nutrition and Dietetics for their contributions. We also thank all mothers and children who agreed to participate in the study.

Financial Support: This study was not funded or supported by governmental, commercial, or non-profit organizations.

Conflict of Interest: The authors declared no conflict of interest.

Authorship: GDB is the author who establishes communication with the organisation and takes part in the planning of the entire flow of the study. Moreover, İKC, and BOB participated in the conception, design of the study. Data collection was conducted by AS, MKT, SA, and PU. İKC performed the data analysis, and also contributed to the interpretation of the findings, drafting and revising the manuscript together with GDB, HKK, and BOB. All authors have critically reviewed and aggreed the final version of the manuscript, and are responsible for all aspects of the study.

Abstract

Objective: To investigate the co-existence of single and multiple anthropometric failures among children using an extended composite index of anthropometric failure (ECIAF). This study aims to elucidate the complex interplay between child-specific and maternal factors, highlighting the multifaceted nature of childhood malnutrition.

Design: A multicentre cross-sectional study as part of the BESLEN project

Setting: Mother-Child Education Centre in the Pendik district of Istanbul, Türkiye

Participants: 1283 children (preschool children, *n* 822, school-aged children, *n* 462) and 1044 mothers

Results: Almost 1/3 of the children included in the study had an anthropometric failure as determined by ECIAF. Weight excess was the leading cause of the total anthropometric failures, most of which were observed to be slightly higher in boys, except for stunting only and cooccurrence of stunting and underweight. Among the mother-related factors, including higher BMI and WC, low maternal age at delivery, low number of children in the household, and being a single parent may be considered predisposing factors to any phenomenon of childhood malnutrition. Among child-related factors, birth weight being \geq 3500 grams had a higher risk for ECIAF failure, and children aged \geq 60 months were more likely to experience stunting and underweight, while those < 60 months had a higher prevalence of weight excess.

Conclusions: The co-existence of stunting and overweight, the occurrence of weight excess in one in three stunted children, and the high risk of central obesity are public health concerns. Also, ECIAF can better assess all aspects of childhood malnutrition than conventional measures.

Keywords: Childhood malnutrition, ECIAF, Obesity, Maternal factors, Anthropometry

1. Introduction

Childhood malnutrition is a persistent global health concern, significantly impacting both individual well-being and societal development. Using anthropometric measures has long been recognized as essential for assessing nutritional status, offering a nuanced understanding of malnutrition's diverse manifestations. Svedberg (2005) and Nandy et al. (2005) introduced the Composite Index of Anthropometric Failure (CIAF)^(1,2), which further expanded to address overnutrition as the extended CIAF (ECIAF) that is a notable contribution, evaluating anthropometric failures in children aged under and over 60 months⁽³⁾. These indices are superior to conventional measures, enabling the distinction between single and multiple anthropometric failures and guiding more targeted policy interventions⁽⁴⁾.

Undernutrition and overweight have traditionally been viewed as distinct issues with different risk factors. Poverty, food insecurity, and infection were associated with undernutrition, while affluence, dietary abundance, and sedentary lifestyles were linked to obesity⁽⁵⁾. However, a growing recognition emphasizes that these two forms of malnutrition often coexist within communities, families, and individuals, exemplified by cases of individuals experiencing both stunted growth and overweight conditions⁽⁶⁾. World Health Organisation (WHO) is gradually acknowledging the significant challenge faced by low-income and middle-income countries (LMICs) dealing with the double burden of malnutrition^(7,8). This phenomenon involves the simultaneous occurrence of undernutrition, including micronutrient deficiencies, underweight, childhood stunting, and wasting, alongside issues of overweight, obesity, and diet-related non-communicable diseases. Current estimates indicate that globally, 2·28 billion children and adults are affected by overweight conditions, while over 150 million children experience stunted growth, underscoring the urgent need for coordinated efforts to address this dual nutritional burden⁽⁹⁾.

A growing body of evidence explores the co-existence of undernutrition and overweight in transition countries^(3,10,11). Findings emphasize the necessity of holistic approaches to address the double burden of malnutrition for sustainable improvements in child health parallel to Sustainable Developmental Goal 2, which aims to end all forms of malnutrition by 2030.

Shifting the focus to maternal factors, various studies have documented the influence of maternal anthropometric features on childhood malnutrition^(12,13). Katoch (2022) underscores the critical role of maternal nutritional status, emphasizing its significant contribution to the risk of stunting and underweight in children⁽¹⁴⁾. This study further contributes to the existing body of evidence by reaffirming the importance of maternal anthropometrics and identifying specific parameters, such as the number of children, maternal age at delivery, marital status, and birth and feeding characteristics, as significant contributors to childhood anthropometric failures.

The objective of this study is to utilize the ECIAF to investigate the prevalence, gender-specific patterns, and co-existence of single and multiple anthropometric failures among children. In addition, this study aims to explore the intricate interplay between child-specific and maternal factors, shedding light on the multifaceted nature of childhood malnutrition. The research also explores the role of child-specific factors and assesses the impact of maternal and birth-related variables, with the overarching goal of enhancing our understanding of childhood malnutrition and informing targeted public health interventions. To the best of the authors' knowledge, this study marks the first attempt in Türkiye to use the ECIAF to determine single and multiple anthropometric failures among children under and over 60 months of age.

2. Material and Methods

2.1.Participants and Data Collection

The BESLEN project is a multicentre cross-sectional study screening the prevalence of childhood malnutrition and examining maternal predisposing factors that may lead to malnutrition. This study aimed to reach all 1300 children and their mothers registered in the seven different Mother-Child Education Centres in the Pendik district of Istanbul, the first of the BESLEN Project centers.

The data collection was carried out in two phases between October 2022 and June 2023. In the first phase, the anthropometrics of 1283 children were measured after their mothers signed the informed consent (17 children's mothers did not volunteer to participate in any phase of the study). Subsequently, the mothers of 1044 of these children volunteered to participate in the second phase of the study in which data on the mother-related and child-related factors were collected face-to-face by the researchers. Apart from the overall sample, the findings were also

presented by dividing the sample into two age groups: < 60 months, preschool children, and ≥ 60 months school-aged children.

2.2. Measures

The trained researchers measured children's anthropometrics using calibrated tools in the first phase. Body weight was measured via a portable digital scale sensitive to 100 grams (g), and the standing height was measured via a portable stadiometer sensitive to 1 millimeter, with participants wearing minimum clothing. Additionally, the children's mid-upper arm circumferences (MUAC) and waist circumferences (WC) were measured in centimeters (cm) with an inflexible tape, and the waist-to-height ratio (WHtR) was calculated. The prevalence of central obesity was reported in the case of children with WHtR $\ge 0.5^{(15)}$. The nutritional status of children was assessed using WHO criteria, and the nutritional status of preschool children was evaluated via the WHO Anthro software, while the WHO AnthroPlus software Version 1.0.4 was used for school-aged children. All children were assessed and grouped according to the criteria of the composite index of anthropometric failure (CIAF) and extended composite index of anthropometric failure (ECIAF)⁽³⁾. Malnutrition was considered in the case of ≤ -2 standard deviation (SD) for weight-for-height (WFH), for height-for-age (HFA), and for weight-for-age (WFA). Following the WHO instructions, for school-aged children (> 60 months), the body mass index-for-age (BMI-for-age) was used to evaluate wasting, and malnutrition was diagnosed with a BMI-for-age \leq -2 SD. For both age groups, the weight excess was also calculated using WFH for preschoolers and BMI-for-age for school-aged children with respect to > 1 SD cut-point⁽³⁾.

In the second phase, apart from mothers' demographic (age, number of children, maternal age at delivery, marital, education, and employment status) and anthropometric characteristics (BMI and WC), information was collected on their children's birth (birth weight and mode of delivery) and feeding status. Only six mothers' anthropometrics were considered invalid because they were pregnant at the data collection, so the BMI and WC measurements were presented for 1038 mothers.

Moreover, mothers completed a scale developed and validated among adults between the ages 18 and 65 by Batmaz H. in 2018 in Türkiye named the Nutrition Knowledge Level Scale for Adults (NKLSA)⁽¹⁶⁾. This scale is comprised of two parts, each of which assesses a distinct aspect of

adult knowledge. The first part, entitled "Basic Nutrition," comprises 20 items, while the second part, "Food Preference," comprises 12 items. The internal consistency of the scale is indicated by Cronbach's alpha coefficients of 0.72 and 0.70, respectively⁽¹⁶⁾, which were found to be 0.78 and 0.95 in our study, respectively. Higher scores indicate higher knowledge levels.

2.3. Statistics

The continuous variables were presented with mean and SD, and percentages were used for qualitative variables. In the first part of the study, the results were presented in two age groups: "< 60 months" and " \geq 60 months." The Chi-square test was used to compare the qualitative variables, and the student *t*-test or the non-parametric equivalent Mann-Whitney U test was used for univariate analysis of continuous outcomes. The same tests were used in the second part of the study as the results were presented in two groups regarding the children's ECIAF status. Moreover, the odds ratio (OR) and confidence intervals (CI) for OR were computed for failures by age groups, maternal demographics, and birth and feeding characteristics of children concerning their ECIAF status. The binary logistic regressions were performed for the significant variables to examine their effect on the likelihood of having ECIAF failure. The Hosmer–Lemeshow test and Omnibus test of model coefficients were evaluated for the goodness of fit. All statistical analyses were conducted using Statistical Package for Social Sciences (SPSS version 23, IBM, Illinois, California) for Windows. The CI was 95 %, and findings with a *P*-value of less than 0.05 were considered statistically significant.

3. Results

The findings of the study were presented and discussed in two phases. Accordingly, the first phase presented data pertaining to children, and the second phase included data pertaining to the mothers.

3.1. Findings of the first phase

The mean age of the 1283 children (preschool children, *n* 822, school-aged children, *n* 462) was 57.9 (SD 6.5) months, of which 48.2 % were girls and 51.8 % were boys. Considering the anthropometrics, the mean WC was 54.4 (SD 5.1) cm, and the mean MUAC was 17.5 (SD 1.9) cm among all children, and the results of these two measurements were quite similar between the

two age groups. However, the WHtR was slightly but significantly higher among preschool children (P < 0.001) (Table 1).

As seen in Table 2, the anthropometric failure prevalence determined by CIAF among overall children was 3.7 %, while it was 31.3 % by ECIAF. The prevalence of CIAF among preschool children was 2.6 %, even significantly higher among school-aged children, with a prevalence of 5.6 % (P < 0.01). However, the failure prevalence was remarkably higher for both age groups considered by ECIAF: 32.5 % for preschoolers and 29.3 % for school-aged children. When the prevalence distribution was rearranged regarding sex, it was observed that boys had a higher prevalence of anthropometric failure by ECIAF in preschoolers and school-aged children (34.3 % and 31.2 %, respectively) (Figure 1).

All categories for single failures were observed to be more prevalent among school-aged children than preschoolers, except category G (weight excess). Wasting (category B) was determined with a 1.3 % prevalence among all children, while it was 1.1 % for preschoolers. Although not statistically significant, it was found that the odds of being wasted were 1.6 times higher among school-aged children than among preschoolers (CI: 0.6 - 4.2), with a prevalence of 1.7 %. Regarding sex distribution, it was observed that the prevalence of wasting was twice as high among preschool boys, while it was even higher among school-aged boys than among girls (Figure 1). The prevalence of stunting (category F) was 2.3 % for all children, 1.7 % for preschoolers, and significantly higher at 3.5 % for school-aged children (P < 0.05), with odds of $2 \cdot 10$ [CI: $1 \cdot 0 - 4 \cdot 3$]. Contrary to other categories, stunting was more prevalent among schoolaged girls than boys, with a prevalence of 4.3 % and 2.8 %, respectively, but equal among preschoolers (Figure 1). Underweight (category Y) prevalence was also 1 % among all children, which was doubled for school-aged children (2.0 %, P < 0.05). In addition, school-aged children were 4.1 times more likely to be underweight than preschool children [CI: 1.25 - 13.3]. The failure Y was more prevalent among boys than girls in both age groups (Figure 1). According to the data, 28.4 % of all children were diagnosed with category G (weight excess only). Contrary to the prevalence of other single failures, preschool children had a significantly higher prevalence of being overweight or obese than school-aged children at 30.9 %. They are 1.41 times more likely to be overweight or obese (CI: 0.52 - 1.17). In addition, in both age groups, boys were observed to have a higher prevalence of weight excess failure than girls (Figure 1).

Considering the simultaneous dual failures, the prevalence was higher among school-aged children for categories C (wasting + underweight) and E (stunting + underweight) than preschoolers (for category C, 0.4 % and 0.2 % and for category E, 1.1 % and 0.5 %, respectively). On the contrary, 1.0 % of preschoolers had a simultaneous dual failure of category H (stunting + weight excess), while it was 0.4 % of school-aged children. Additionally, no school-aged children were determined to be at risk for triple failure (category D), but this category has a 0.2 % prevalence among preschoolers. All children identified in categories C and D consisted of boys (Figure 1). The dual failure of stunting + underweight was more prevalent among school-aged girls than boys, but this prevalence was more common among preschool boys (Figure 1).

The anthropometrics of the children were also considered regarding their ECIAF status. As shown in Table 3, children who have ECIAF failure have significantly higher mean WC (57.8 cm, SD 6.2), higher mean WHtR (0.53, SD 0.05), and higher MUAC (18.8 cm, SD 2.1) (P < 0.001 for all).

3.2. Findings of the second phase

The mean age of 1044 mothers was 34 (SD 5.6) years, ranging from 21 to 56. In terms of anthropometric measures, the mean BMI was $28 \cdot 0 \pm 5 \cdot 6 \text{ kg/m}^2$, indicating overweight, and the range was quite wide, from 15.9 to $60 \cdot 2 \text{ kg/m}^2$. On the other hand, the WC falls within the desired range with a mean of 87.7 (SD 12.9) cm. Moreover, there was statistical significance in the ECIAF status of their children; BMI and WC were significantly higher in mothers of children with at least one failure (P < 0.001 for both) (Table 4).

As presented in Table 4, the mean number of children of the mothers was 2·3 (SD 0·94), ranging from 1 to 7. This mean was the same for the ECIAF no failure group, however slightly but significantly lower for the ECIAF failure group (P < 0.05). A similar pattern was observed for the maternal age at delivery with a mean of 28·4 (SD 5·7) years, almost the same as in the ECIAF no failure group, but mothers of children in the ECIAF failure group had a significantly lower maternal age at delivery (P < 0.05).

The vast majority of the women (98.3 %) were married at the time of data collection, and 95.6 % of them were unemployed/housewives, although more than half of them reported having a high school or university degree (35.2 %, and 18.4 %, respectively). Regarding the ECIAF status of their children, no statistically significant risk was observed for their employment or educational status (P > 0.05). However, it was observed that the odds of having ECIAF failure in children were significantly increased by 3.89 [CI: 0.5 - 10.1] for single mothers (P < 0.01) (Table 4).

In the second part of the study, the mothers' nutrition knowledge level was also examined and presented in Table 4 in two parts. Accordingly, the mean total basic nutrition score of mothers was 44.6 (SD 10.9), with a range from 17 to 71, and the mean total food preference score was 30.3 ± 12.9 , ranging from 0 to 48. These scores did not differ significantly between ECIAF groups (P > 0.05 for both).

In Table 5, the birth and feeding characteristics of children were presented. Of the 1044 children, 1015 (97.2 %) were breastfed, and the mean duration of breastfeeding was 19.3 (SD 9.5) months, ranging from 0.5 to 60 months. However, breastfed children were slightly but insignificantly more likely to have ECIAF failure (29.4 %, P > 0.05). Moreover, although a slightly longer duration of breastfeeding was reported in children without ECIAF failure, this difference was not significantly different (P > 0.05). The proportion of formula users among children was 43.2 %, and 28.8 % of formula users had ECIAF failure compared to 29.7 % of non-users. However, this difference was insignificant (P > 0.05).

The proportion of children born by cesarean section was $55 \cdot 1$ %, of whom $29 \cdot 9$ % had ECIAF failure, compared with a slightly but insignificantly lower rate of $28 \cdot 6$ % in children born vaginally (P > 0.05). When the birth weight of the children was analyzed, the mean of overall 1044 children was $3229 \cdot 4$ (SD $624 \cdot 5$) kg which were quite similar in groups according to the ECIAF failure status of the children. However, although the mean birth weights of children with and without ECIAF failure were so close and fell within the clinically normal birth weight range, this small difference was statistically significant, indicating that those with ECIAF had a higher birth weight with a mean of $3288 \cdot 4$ (SD $629 \cdot 2$) kg (P < 0.05). Due to the similar means between the two groups of ECIAF, this significant finding was investigated further. Accordingly, the cut-

off value was determined as 3500 grams^(17,18) and accordingly, it was determined that children with a birth weight above this value had a 1.5-fold higher risk of ECIAF (Table 5).

Binary logistic regression was conducted for the variables significantly related to children's ECIAF status. There was one categorical explanatory variable, marital status, and two continuous explanatory variables, maternal BMI and WC, for the binary logistic regression model to explain the factors contributing to children's ECIAF status. Accordingly, the binary-logistic regression results showed that maternal BMI, number of children, and marital status of the mother significantly affected the likelihood of having ECIAF failure (Table 6).

4. Discussion

Determining the prevalence of malnutrition in communities is important to increase targeted interventions to reduce malnutrition globally, in line with the 2030 Sustainable Development Goals⁽¹⁹⁾. As is common in transition countries, Türkiye has undergone structural changes over decades, and obesity, along with malnutrition, has become a public health concern. Several studies also emphasized obesity as an alarming risk, with an 11.6-fold increase in prevalence over the last 20 years in our country^(10,20). Moreover, according to a recent report by WHO (2022), Türkiye has been listed as the 10th in obesity and overweight prevalence among WHO European region countries, with a total prevalence of nearly 50 %⁽²¹⁾. This study is noteworthy because it is the first to use the ECIAF index to assess both overnutrition and undernutrition in children under and over 60 months of age.

Almost one in three children included in the study had an anthropometric failure determined by ECIAF. This finding falls within the range identified among Ghanaian children by Kuwornu et al. $(2022)^{(4)}$, although another study data of 10 879 children from six different regions of Argentina indicated a lower prevalence $(15 \cdot 1 \ \%)^{(3)}$ and also a higher prevalence $(37 \cdot 5 \ \%)$ was reported from El Salvador conducted in 334 school children⁽²²⁾. When looking at the sex and age groups, this prevalence shows a similar distribution, although it is slightly higher in boys, in line with the literature^(3,22), and also in preschoolers. Contrarly, Bejarano et al. (2019) also indicated a lower prevalence among boys with ages lower than 60 months⁽³⁾. When all ECIAF categories were evaluated, weight excess categories (overweight and obesity) were found to be the main contributor to total anthropometric failure, similar to other studies^(3,22). In addition, more than

half of the children, which is higher than the literature⁽²²⁾, had an elevated risk of central obesity on the basis of their WHtR, and this risk was 6·11 times higher in children with ECIAF failure, in support of the findings from ECIAF. A recent review of the 25-year trend in overweight and obesity in Türkiye reported a doubling of the prevalence in preschool children, from 6·6 % to $14\cdot4$ %⁽¹¹⁾. The current study reported an even higher prevalence (28·4 %) in both age groups, which once again underlines the increasing trend in the prevalence of obesity. Supportively, studies reported high levels of obesity in children from households with a lower socioeconomic level in high-income countries and children living in urban areas in low and middle-income countries⁽¹¹⁾. It should be highlighted that this study was conducted in a low socioeconomic area of the country's largest metropolitan city, home to diverse demographic groups.

Stunting, wasting, and underweight failures were the most common single failures after the category of weight excess among overall children, indicating the risk of chronic and acute malnutrition as well as obesity in the population. The co-existence of overnutrition and undernutrition phenomenon can be explained by the double burden of malnutrition, which is common in transition countries^(3,10,11); although Türkiye is not listed among them, this phenomenon persists as a public health concern in our country.

Among the single failures following category G (weight excess), undernutrition (wasting, stunting, and underweight) was more prevalent in school-aged children than in preschoolers, parallel to the literature⁽²³⁾. Moreover, the odds of being stunted were 2.10 times, and being underweight 4.1 was significantly higher among school-aged children. This trend could be elucidated by an increase in energy consumption due to being physically more active and the persistence of poor nutrition practices, which are an extension of related maternal and childcare factors. In terms of sex distribution, the only single failure is stunting, which is higher among girls, particularly those of school age, contrary to the literature⁽²⁴⁾. However, wasting and underweight were higher in boys than in girls, in keeping with the various studies indicating that malnutrition/retarded growth (stunting + underweight) was more prevalent in boys^(22–25). This observation has been attributed to the female eco-stability hypothesis, which states that girls' smaller body size and naturally greater adiposity confer an advantage in being more resistant to external environmental factors that modulate ontogenetic development^(25,26). The multiple anthropometric failure prevalence of overall children was not at a remarkable level in our study.

However, the prevalence of simultaneous existence of stunting and weight excess was at a comparable level to Pedrero et al. findings (2022), indicating 0.9 % ECIAF-H failure⁽²²⁾. Moreover, based on the findings that 10 out of 30 stunted children were overweight/obese in this study, we should emphasize that one in every three stunted children has a condition related to weight excess. This observation is supported by the literature, as short stature is considered a notable factor in the obesity epidemic among children and adults⁽³⁾. Furthermore, children with multiple failures are more prone to have ill health and an increased mortality risk⁽⁴⁾.

Besides the environmental and genetic factors that contribute to childhood malnutrition, the impact of parental influence should not be overlooked. In many societies, mothers are generally the primary caregivers of children and make an important contribution to children's healthy eating patterns. Thus, understanding the role of maternal factors is of great importance in controlling childhood malnutrition.

Current anthropometric characteristics of mothers have been identified as one of the strongest predictors of malnutrition levels in children, especially excess weight⁽²⁷⁾. Our study findings also confirmed that BMI and WC are significantly higher in mothers of children with ECIAF failure. Parallel to the present findings, a study by Vehapoğlu et al. revealed that mothers who were overweight were significantly more likely to have obese children compared to normal-weight mothers⁽²⁸⁾. Another study has reported that the BMI of the mother can account for 25-40 % of the BMI of the children⁽²⁹⁾. This effect may be attributed to heavier mothers' lower levels of physical activity, their preference for high-fat foods, and their less healthy food choices⁽³⁰⁾.

The mother's education level and employment status may be counted as potential contributors to childhood obesity; this study could not display any statistical significance regarding the anthropometric failures of children. Interestingly, although single mother population was quite low in this study, a remarkable finding has demonstrated that the risk of ECIAF failure in children of single mothers was significantly increased by 3.89 times. Consistently, several researchers have pointed out that children of single parents are more likely to be obese^(31,32). Single-parent households' lack of resources, preference for foods high in saturated fat and sugar, low levels of physical activity, lack of time to prepare healthy meals, increased exposure to television, and uncertain housing conditions can be cited as potential arguments for this association^(11,33). Moreover, we observed a significantly lower maternal age at delivery among

the ECIAF group, similar to a study among the Chinese population indicating maternal age under 28 as a remarkable factor in predicting childhood obesity⁽³⁴⁾.

Maternal feeding practices have been described as an important variable contributing to childhood obesity, and approaches may differ between primiparous and multiparous mothers⁽³⁵⁾. The present study revealed that mothers' mean number of children was significantly lower in the ECIAF group. Consistent with this outcome, several researchers have documented higher obesity prevalence in children born to primiparous mothers compared to children born to multiparous women^(36,37). The concept of parity and childhood obesity can be mediated by sibling interactions, higher physical activity opportunities, and the tendency for household income to be shared equally among all children, which reduces access to large amounts of food and/or increases the need for restrictions in large families⁽³⁸⁾.

Considering the processes that begin with the birth of the child, it is reported that infant feeding practices, as well as mode of delivery and birth weight, may have an impact on childhood obesity. Several studies reported the protective effect of prolonged breastfeeding practices on children's risk of obesity^(27,39). However, in addition to the limitations of the study designs, environmental, genetic, cultural, and maternal and infant variables have all contributed to the inconsistent findings in the literature regarding the effect of breastfeeding on obesity^(40,41). Thus, this study failed to observe any statistical significance in relation to the ECIAF status of the children and their breastfeeding status, duration, and formula use, supporting the literature. Regarding the mode of delivery, WHO recommends no more than 10 to 15 percent of the population to have C-sections; over the past few years, there has been a dramatic increase in global C-section rates, including Türkiye, which is among the top five countries with the highest rate (50.8 %) in the world $^{(42,43)}$. Although this study did not determine any statistical difference regarding ECIAF failure of the children, a mounting body of literature supports a relationship between C-sections and an increased risk of overweight and obesity in offspring^(34,44,45). Birth weight predominantly reflects intrauterine developmental conditions, particularly maternal-fetal nutrition, and recently, the "fetal origins hypothesis" has emerged as one of the most potent theoretical frameworks in medicine regarding the early causes of later disease⁽⁴⁶⁾. Supportingly, a substantial body of research is evidence that high birth weight is a risk factor for childhood obesity^(47,48). Consistent with this observation, the current study has demonstrated that children

with ECIAF had significantly higher birth weights. One of the possible explanations might be related to increased levels of prenatal growth hormones, including insulin and insulin-like growth factors I and II, which can increase the risk of obesity later in life⁽⁴⁸⁾.

A recent study has demonstrated that low levels of maternal nutritional knowledge are associated with an elevated risk of childhood obesity⁽⁴⁹⁾. Although we hypothesized that children of mothers with high nutritional knowledge would have a lower prevalence of ECIAF, the study indicates no significant association between obesity and mothers' knowledge of basic nutrition and their food preferences, despite their awareness of the relationship between diet and health, in line with Forh et al. (2022) study⁽⁵⁰⁾. This circumstance justifies the fact that knowledge does not always translate into practice and that a number of potential confounding variables, including availability, accessibility, utilization, and stability of food, may influence the situation.

To highlight the strengths of this study, a high number of children were reached, as well as the simultaneous collection of data from the mothers of these children enabled the findings to be obtained through realistic observations of measures. Moreover, we had the chance to obtain anthropometric data objectively via calibrated devices by trained researchers. Despite these strengths, the cross-sectional design of the study limits the determination of the causal relationship between variables. Also, another handicap is that data on birth-related factors of children are collected based on the mother's memory. Moreover, although ECIAF evaluates all aspects of the child's anthropometry, insufficient reflection of micronutrient deficiencies precludes the evaluation of the triple burden of malnutrition.

5. Conclusion

This is the first study in Türkiye that used the ECIAF to determine the concurrent phenomenon of malnutrition, including single and multiple anthropometric failures among children under and over 60 months of age, and investigated the possible contributors associated with child-specific and maternal factors.

According to the findings of this study, we can conclude that weight excess was the main contributor to total anthropometric failures, most of which were observed to be slightly higher in boys except stunting only and co-occurrence of stunting and underweight. Moreover, the

simultaneous existence of stunting and overweight, the observation that one in every three stunted children has a condition related to weight excess, and the elevated central obesity risk can be counted as concerning public health issues. This study also contributes an important aspect by investigating the role of maternal and birth-related factors in controlling childhood malnutrition. Accordingly, among the mother-related factors, including higher BMI and WC, low maternal age at delivery, low number of children in the household, and being a single parent may be considered predisposing factors to any phenomenon of childhood malnutrition. Among child-related factors, birth weight being 3500 grams and over, had a higher risk for ECIAF failure, and children aged 60 months or older were more likely to experience stunting and underweight, while those under 60 months had a higher prevalence of weight excess.

In conclusion, it is possible to assess all aspects of childhood malnutrition better using ECIAF rather than conventional measures, making it a promising tool for both future research and clinical practice.

Ethical Standards Disclosure: This study was conducted according to the guidelines laid down in the Declaration of Helsinki and all procedures involving research study participants were approved by the Marmara University Faculty of Health Sciences Non-Interventional Clinical Research Ethics Committee with the date and number 27.02.2020/35. Written informed consent was obtained from all participants.

References

- Svedberg P (2000) *Poverty and Undernutrition: Theory, Measurement, and Policy*. (Oxford, 2000; online edn, Oxford Academic), Oxford University Press.
- 2. Nandy S, Irving M, Gordon D *et al.* (2005) Poverty, child undernutrition and morbidity: New evidence from India. *Bull World Health Organ* **83**(3):210–6.
- Bejarano IF, Oyhenart EE, Torres MF et al. (2019) Extended composite index of anthropometric failure in Argentinean preschool and school children. Public Health Nutr 22(18):3327–35.
- 4. Kuwornu JP, Amoyaw J, Manyanga T *et al.* (2022) Measuring the Overall Burden of Early Childhood Malnutrition in Ghana: A Comparison of Estimates From Multiple Data Sources.

Int J Heal Policy Manag **11**(7):1035–46.

- 5. Doak CM, Adair LS, Bentley M *et al.* (2005) The dual burden household and the nutrition transition paradox. *Int J Obes* **29**(1):129–36.
- 6. Wells JC, Sawaya AL, Wibaek R *et al.* (2020)The double burden of malnutrition: aetiological pathways and consequences for health. *Lancet* **395**(10217):75–88.
- World Health Organisation, The double burden of malnutrition (2017) https://iris.who.int/bitstream/handle/10665/255413/WHO-NMH-NHD-17.3-eng.pdf?ua=1 (accessed 15 December 2023).
- 8. Alem AZ, Yeshaw Y, Liyew AM *et al.* (2023) Double burden of malnutrition and its associated factors among women in low and middle income countries: findings from 52 nationally representative data. *BMC Public Health* **23**(1):1–16.
- 9. Popkin BM, Corvalan C & Grummer-Strawn LM (2020) Dynamics of the double burden of malnutrition and the changing nutrition reality. *Lancet* **395**(10217):65-74.
- Comba A, Demir E & Barış Eren N (2019) Nutritional status and related factors of schoolchildren in Çorum, Turkey. *Public Health Nutr* 22(1):122–31.
- 11. Görçin Karaketir Ş, Lüleci NE, Eryurt MA *et al.* (2023) Overweight and obesity in preschool children in Turkey: A multilevel analysis. *J Biosoc Sci* **55**(2):344–66.
- Amaha ND & Woldeamanuel BT (2021) Maternal factors associated with moderate and severe stunting in Ethiopian children: analysis of some environmental factors based on 2016 demographic health survey. *Nutr J* 20(1):1–9.
- Mamidi RS, Banjara SK, Manchala S *et al.* (2022) Maternal Nutrition, Body Composition and Gestational Weight Gain on Low Birth Weight and Small for Gestational Age—A Cohort Study in an Indian Urban Slum. *Children* 9(10):1460.
- Katoch OR (2022) Determinants of malnutrition among children: A systematic review. *Nutrition* 96(April):111565.
- 15. Hsieh SD, Yoshinaga H & Muto T (2003) Waist-to-height ratio, a simple and practical index for assessing central fat distribution and metabolic risk in Japanese men and women. *Int J*

Obes **27**(5):610–6.

- 16. Batmaz H (2018) Yetişkinler İçin Beslenme Bilgi Düzeyi Ölçeği Geliştirilmesi ve Geçerlik-Güvenirlik Çalışması. Graduate Thesis, Marmara University, İstanbul. https://toad.halileksi.net/wp-content/uploads/2022/07/yetiskinler-icin-beslenme-bilgi-duzeyiolcegi-gelistirilmesi-ve-gecerlik-guvenirlik-calismasi-toad.pdf (accessed 15 December 2023).
- Laopaiboon M, Lumbiganon P, Rattanakanokchai S *et al.* (2019) An outcome-based definition of low birthweight for births in low- and middle-income countries: A secondary analysis of the WHO global survey on maternal and perinatal health. *BMC Pediatr* 19(1):1–9.
- Shi J, Guo Q, Fang H *et al.* (2024)The Relationship between Birth Weight and the Risk of Overweight and Obesity among Chinese Children and Adolescents Aged 7–17 Years. *Nutrients* 16(5).
- 19. United Nations, Report of the Open Working Group of the General Assembly on Sustainable Development Goals (2014) http://www.un.org/ga/search/view_doc.asp?symbol=A/68/970&Lang=E (accessed 15 December 2023).
- Alper Z, Ercan İ & Uncu Y (2018) A meta-analysis and an evaluation of trends in obesity prevalence among children and adolescents in Turkey: 1990 through 2015. *JCRPE J Clin Res Pediatr Endocrinol* 10(1):59–67.
- World Health Organisation (2022) European Regional Obesity Report 2022. Copenhagen: World Health Organisation.
- Pedrero-Tomé R, López-Ejeda N, Sánchez Alvarez M *et al.* (2022) Household food insecurity and nutritional status of schoolchildren in rural regions of Bajo Lampa, El Salvador (2018– 2019). *Ecol Food Nutr* 61(2):128–43.
- 23. Berhanu A, Garoma S, Arero G *et al.* (2022) Stunting and associated factors among school-age children (5–14 years) in Mulo district, Oromia region, Ethiopia. SAGE Open Med 10(October):1–10.
- 24. Thurstans S, Opondo C, Seal A *et al.* (2020) Boys are more likely to be undernourished than girls: A systematic review and meta-analysis of sex differences in undernutrition. *BMJ Glob*

Heal **5**(12):e004030.

- 25. Díez Navarro A, Marrodán Serrano MD, Gómez De Arriba A *et al.*(2017) Ecoestabilidad femenina y malnutrición severa infantil: Evidencia a partir de intervenciones de ayuda humanitaria de Acción Contra el Hambre en países africanos, Asiáticos y Latinoamericanos. *Nutr Clin y Diet Hosp* **37**(4):127–34.
- Stini WA (1975) Adaptive strategies of human populations under nutritional stress. In: Watts ES, Johnston FE, Lasker GW, editors. Biosocial interrelations in population adaptation. The Hague; Mouton.
- Ma J, Qiao Y, Zhao P *et al.*(2020) Breastfeeding and childhood obesity: A 12-country study. *Matern Child Nutr* 16(3):1–9.
- 28. Vehapoglu A, Goknar N, Turel O *et al.* (2017) Risk factors for childhood obesity: Do the birth weight, type of delivery, and mother's overweight have an implication on current weight status? *World J Pediatr* **13**(5)457-464
- 29. Anderson PM & Butcher KF (2006) Childhood obesity: Trends and potential causes. *Futur Child* **16**(1):19–45.
- 30. Fuemmeler BF, Lovelady CA & Zucker NL (2013) Parental obesity moderates the relationship between childhood appetitive traits and weight. *Obes (Silver Spring)* 21(4):815–23.
- Duriancik DM & Goff CR (2019) Children of single-parent households are at a higher risk of obesity: A systematic review. *J Child Heal Care* 23(3):358–69.
- 32. Huffman FG, Kanikireddy S & Patel M (2010) Parenthood-a contributing factor to childhood obesity. *Int J Environ Res Public Health* **7**(7):2800–10.
- 33. Hsu PC, Hwang FM, Chien MI *et al.* (2022) The impact of maternal influences on childhood obesity. *Sci Rep* **12**(1):1–6.
- Liu S, Lei J, Ma J *et al.* (2020) Interaction between delivery mode and maternal age in predicting overweight and obesity in 1,123 Chinese preschool children. *Ann Transl Med* 8(7):474.

- 35. Hu S, Wacharasin C & Sangin S (2023) Factors associated with feeding behaviors among mothers of obese infants: a cross-sectional study. *Transl Pediatr* **12**(5):1004–16.
- 36. Woronko C, Merry L, Uckun S *et al.* (2023) Prevalence and determinants of overweight and obesity among preschool-aged children from migrant and socioeconomically disadvantaged contexts in Montreal, Canada. *Prev Med Reports* 36(September):102397.
- 37. Gaillard R, Rurangirwa AA, Williams MA *et al.* (2014) Maternal Parity, fetal and childhood growth, and cardiometabolic risk factors. *Hypertension* **64**(2):266–74.
- Park SH & Cormier E (2018) Influence of Siblings on Child Health Behaviors and Obesity: A Systematic Review. *J Child Fam Stud* [Internet] 27(7):2069–81.
- 39. Qiao J, Dai LJ, Zhang Q *et al.* (2020) A meta-analysis of the association between breastfeeding and early childhood obesity. *J Pediatr Nurs* **53**(July):57–66.
- 40. Grube MM, Von Der Lippe E, Schlaud M *et al.* (2015) Does breastfeeding help to reduce the risk of childhood overweight and obesity? A propensity score analysis of data from the KiGGS study. *PLoS One* 10(3):e0122534.
- 41. Lefebvre CM & John RM (2014) The effect of breastfeeding on childhood overweight and obesity: A systematic review of the literature. *J Am Assoc Nurse Pract* **26**(7) 386-401.
- 42. World Health Organisation (2015) WHO Statement on Caesarean Section Rates. Geneva: World Health Organisation https://iris.who.int/bitstream/handle/10665/161442/WHO_RHR_15.02_eng.pdf?sequence=1(a ccessed 15 December 2023).
- 43. Betran AP, Ye J, Moller AB *et al.* (2021)Trends and projections of caesarean section rates: Global and regional estimates. *BMJ Glob Heal* 6(6):e005671.
- 44. Li HT, Zhou YB & Liu JM (2013) The impact of cesarean section on offspring overweight and obesity: A systematic review and meta-analysis. *Int J Obes* **37**(7):893–9.
- Zhang S, Qin X, Li P *et al.* (2022) Effect of Elective Cesarean Section on Children's Obesity From Birth to Adolescence: A Systematic Review and Meta-Analysis. *Front Pediatr* 9(January):793400.

- Schellong K, Schulz S, Harder T *et al.* (2012) Birth Weight and Long-Term Overweight Risk:
 Systematic Review and a Meta-Analysis Including 643,902 Persons from 66 Studies and 26
 Countries Globally. *PLoS One* 7(10):e47776.
- Qiao Y, Ma J, Wang Y *et al.* (2015) Birth weight and childhood obesity: a 12-country study. *Int J Obes* 5(S2):S74–9.
- 48. Yu ZB, Han SP, Zhu GZ *et al.* (2011) Birth weight and subsequent risk of obesity: A systematic review and meta-analysis. *Obes Rev* **12**(7):525–42.
- Xu Z, Zhao Y, Sun J *et al.* (2022) Association between dietary knowledge and overweight and obesity in Chinese children and adolescents: Evidence from the China Health and Nutrition Survey in 2004–2015. *PLoS One* 17(12 December):1–14.
- 50. Forh G, Apprey C & Frimpomaa Agyapong NA (2022) Nutritional knowledge and practices of mothers/caregivers and its impact on the nutritional status of children 6–59 months in Sefwi Wiawso Municipality, Western-North Region, Ghana. *Heliyon* 8(12):e12330.

		0		1292)	< 60 m	ontl	ns	≥60 m	ontł	15	
		Overall (<i>n</i> 1283)		(n 822,	(<i>n</i> 822, 64·1 %)			(<i>n</i> 461, 35·9 %)			
		n		%	n		%	n		%	
Corr	Boys	664		51.8	414		50.4	250		54.2	
Sex	Girls	619		48.2	408		49.6	211		45.8	
		Mean	±	Min	Mean	±	Min	Mean	±	Min.	•
		SD		Max.	SD		Max.	SD		Max.	
Age (me	onths)	57·9 6·5	±	41 - 84	53.9 3.5	±	41 - 59	65·1 4·2	±	60 - 84	
Weight	(kg)	19·2 3·5	±	10·8 - 39·2	18·7 3·2	±	10·8 - 39·2	20·2 3·7	±	13·1 37·2	-
Height	(cm)	109·3 5·7	±	89 - 128	107·5 5·2	±	89·0 - 126·0	112·3 5·5	±	95·0 128·0	-
MUAC	(cm)	17·5 1·9	±	11 - 27	17·5 1·9	±	11 - 27	17·6 2·03	±	12 - 24	
WC (cn	n)	54·4 5·1	±	40 - 104	54·2 5·0	Ŧ	40 - 104	54·8 5·2	±	43·0 85·5	-
											P ^a
WHtR		0·5 0·04	±	0.4 - 1.03	0·50 0·04	±	0.4 - 1.03	0·49 0·04	±	0.4 - 0.8	< 0.001**

Table 1 Demographic and anthropometric characteristics of children

MUAC, mid-upper arm circumference; WC, Waist circumference; WHtR, waist-to-height ratio; SD, standard deviation; Min, minimum; Max. maximum; ^a independent sample *t*-test; ^{*} P < 0.05; ^{**} P < 0.01;

	Overall	< 60	≥ 60		
	(<i>n</i> 1283)	months	months		
	% (<i>n</i>)	% (n)	% (n)	Р	OR [95% CI]
CIAF (B+C+D+E+F+Y)	3.7 (47)	2.6 (21)	5.6 (26)	$0.005^{a^{**}}$	2.28 [1.27 -
ECIAF	31.3 (402)	32.5	29.3	0.236^{a}	1.16 [0.91 -
(A) No failure for CIAF	96.3	97.4	94.4	-	-
(A) No failure for ECIAF	68.7	67.5	70.7	-	-
(B) Wasting Only	1.3 (17)	1.1 (9)	1.7 (8)	0.336^{a}	1.6 [0.6 -
(C) Wasting and Underweight	0.3 (4)	0.2 (2)	0.4 (2)	0.622^{b}	1.8 [0.25 -
(D) Stunting, Wasting and	0.2 (2)	0.2 (2)	0	NA	NA
(E) Stunting and Underweight	0.7 (9)	0.5 (4)	1.1 (5)	0.296^{b}	2.24 [0.60 -
(F) Stunting only	2.3 (30)	1.7 (14)	3.5 (16)	0.044^{a^*}	2.10 [1.0 -
(Y) Underweight only	1.0 (13)	0.5 (4)	2.0 (9)	0.018^{b^*}	4.10 [1.25 -
(G) ^c Weight excess	28.4 (365)	30.9	24.1	0.009^{a^*}	1.41 [1.1 -
(H) ^c Stunting and weight	0.8 (10)	1.0 (8)	0.4 (2)	0.510^{b}	2.3 [0.5 -

Table 2 Prevalence of CIAF and ECIAF with different forms of anthropometric failures

CIAF, Composite Index of Anhtropometric Failure; ECIAF, Extended Composite Index of Anhtropometric Failure; OR, Odds Ratio; CI, Confidence Interval; NA, Not applicable; ^a Pearson Chi-square test; ^b Fisher's exact test; ^c parameters only included for ECIAF; ^{*} P < 0.05; ^{**} P < 0.01

	ECIAF		
	Failure (<i>n</i> 402)	No failure (<i>n</i> 881)	
	Mean ± SD (min max.)	Mean ± SD (minmax.)	Р
WC (cm)	57·8 ± 6·2 (43 - 104)	$52.9 \pm 3.4 \ (40.0 - 85.5)$	$< 0.001^{a^{**}}$
WHtR	$0.53 \pm 0.05 (0.38$ -	$0.49 \pm 0.03 \ (0.4 - 0.8)$	$< 0.001^{a^{**}}$
MUAC (cm)	18.8 ± 2.1 (13 - 26)	16·9 ± 1·5 (11 - 24)	$< 0.001^{b^{**}}$

Table 3 Anthropometric measurements regarding the ECIAF status of the children

ECIAF, Extended Composite Index of Anthropometric Failure; SD, standard deviation; Min, minimum; Max. maximum; WC, Waist circumference; WHtR, waist-to-height ratio; MUAC, mid-upper arm circumference; ^a independent sample t-test; ^b Mann Whitney U test; ^c parameters only included for ECIAF; ^{**} P < 0.01

		Overall (<i>n</i>	ECIAF Failure (<i>n</i>	No failure (n	
		Mean ± SD	Mean ± SD	Mean ± SD	P ^a	
Age (years)		$(minmax.) 34 \pm 5.6 (21 - 56)$	(minmax.) 33.5 ± 5.5 (21 - 47)	(minmax.) $34 \cdot 2 \pm 5 \cdot 6$ (21 - 56)	0.052	
BMI (kg/m ²) (n 1038)		(21 - 50) $28 \cdot 0 \pm 5 \cdot 6$ $(15 \cdot 9 - 60 \cdot 2)$	29.2 ± 5.7	27.5 ± 5.5	< 0.001**	
WC (cm) (<i>n</i>	1038)	$87.7 \pm 12.9^{\circ}$	$(17 \cdot 2 - 51 \cdot 8)$ $90 \cdot 1 \pm 13 \cdot 0$	(15.9 - 60.2) 86.7 ± 12.7	< 0.001**	
Number of	children	(53 - 136) $2 \cdot 3 \pm 0 \cdot 9$	(57 - 127) $2 \cdot 2 \pm 0 \cdot 9$	(53 - 136) $2 \cdot 3 \pm 1 \cdot 0$	0.021*	
	ge at delivery	(1 - 7) 28.4 ± 5.7	(1 - 6) 27.8 ± 5.6	(1 - 7) 28.7 ± 5.8	0.040^{*}	
(years)		(16 - 52) 44·6 ± 10·9	(17 - 41) 43.9 ± 11.2	(16 - 52) 44.8 ± 10.9	0.235	
Basic nutrition score Food preference score		(17 - 71) 30.3 ± 12.9	(20 - 71) 29.4 ± 13.0	(17 - 70) 31.0 ± 12.8		
		(0 - 48)	(0 - 48)	(0 - 48)	0.147	
					-h OR	2
		% (<i>n</i>)	% (<i>n</i>)	% (n)	P^b [95	
Marital	Currently	% (<i>n</i>) 98·3 (1026)	% (<i>n</i>) 28·8 (295)	% (<i>n</i>) 71·2 (731)	$\frac{P^{\nu}}{0.003^{*}} = \frac{95}{3.8}$	9 9
Marital status	Currently married Single				<i>P²</i> [95	9 5-
	married	98.3 (1026)	28.8 (295)	71.2 (731)	$ \begin{array}{c} P^{\nu} \\ $	9 5-
	married Single Illiterate Literate	98·3 (1026) 1·7 (18)	28·8 (295) 61·1 (11)	71·2 (731) 38·9 (7)	$ \begin{array}{c} P^{\nu} \\ $	9 5-
status Education	married Single Illiterate Literate Primary school	98·3 (1026) 1·7 (18) 1·3 (14)	28.8 (295) 61.1 (11) 28.6 (4)	71·2 (731) 38·9 (7) 71·4 (10)	$ \begin{array}{c} P^{\nu} \\ $	9 5- 1]
status	married Single Illiterate Literate Primary	98.3 (1026) 1.7 (18) 1.3 (14) 1.6 (17)	$28 \cdot 8 (295) \\61 \cdot 1 (11) \\28 \cdot 6 (4) \\23 \cdot 5 (4)$	71.2 (731) 38.9 (7) 71.4 (10) 76.5 (13)	P [≠] [95] 0.003* 3.8 * [1.4] 10.	9 5- 1]
status Education	married Single Illiterate Literate Primary school Middle	98.3 (1026) 1.7 (18) 1.3 (14) 1.6 (17) 22.7 (237)	28.8 (295) 61.1 (11) 28.6 (4) 23.5 (4) 30 (71)	71.2 (731) 38.9 (7) 71.4 (10) 76.5 (13) 70 (166)	P [≠] [95 0.003* 3.8 * [1.4] 10.	9 5- 1]
status Education	married Single Illiterate Literate Primary school Middle school	98.3 (1026) 1.7 (18) 1.3 (14) 1.6 (17) 22.7 (237) 20.7 (216)	$28 \cdot 8 (295)$ $61 \cdot 1 (11)$ $28 \cdot 6 (4)$ $23 \cdot 5 (4)$ $30 (71)$ $29 \cdot 6 (64)$	71.2 (731) 38.9 (7) 71.4 (10) 76.5 (13) 70 (166) 70.4 (152)	P [≠] [95 0.003* 3.8 * [1.4] 10.	9 5- 1]
status Education	married Single Illiterate Literate Primary school Middle school High school	98.3 (1026) 1.7 (18) 1.3 (14) 1.6 (17) 22.7 (237) 20.7 (216) 35.2 (368)	$28 \cdot 8 (295)$ $61 \cdot 1 (11)$ $28 \cdot 6 (4)$ $23 \cdot 5 (4)$ $30 (71)$ $29 \cdot 6 (64)$ $28 \cdot 3 (104)$	71.2 (731) 38.9 (7) 71.4 (10) 76.5 (13) 70 (166) 70.4 (152) 71.7 (264)	P [≠] [95 0.003* 3.8 * [1.4] 10.	9 5- 1]

Table 4 Demographic, anthropometric, and scores of nutrition knowledge level of mothers

ECIAF, Extended Composite Index of Anthropometric Failure; SD, standard deviation; Min, minimum; Max. maximum; WC, Waist circumference; BMI, Body Mass Index; ^a independent sample *t*-test; ^b Pearson Chi-square test; OR, Odds Ratio; CI, Confidence Interval; NA, Not applicable; *P < 0.05; **P < 0.01

			ECIAF		
		Overall (n	Failure	No failure	
		% (<i>n</i>)	% (<i>n</i>)	% (<i>n</i>)	P ^a OR
		70 (<i>n</i>)	70 (<i>IL</i>)	70 (<i>n</i>)	[95%
Breastfed	Breastfed	97.2 (1015)	29.4 (298)	70.6 (717)	1.1
status				50 4 (04)	0.836 [0.5-
status	Not Brestfed	2.8 (29)	27.6 (8)	72.4 (21)	2.5]
Formula	Formula	43.2 (451)	28.8 (130)	71.2 (321)	0.96
use	usod	56.0 (502)	0 = (1 = c)	70.2 (417)	0.764 [0.7-
use	No Formula	56.8 (593)	29.7 (176)	70.3 (417)	1.3]
Mode of	Vaginal	44.9 (469)	28.6 (134)	71.4 (335)	1.1
delivery					0.636 [0.8-
uenvery	C-section	55.1 (575)	29.9 (172)	70.1 (403)	1.4]
Birth	< 3500 g	62.5 (653)	26.3 (172)	73.7 (481)	1.5
weight					0.006 [1.1-
categories	≥3500 g)	37.5 (391)	34.3 (134)	65.7 (257)	1.9]
		Mean ± SD	Mean ± SD	Mean ± SD	P ^b
		(minmax.)	(min	(min	-
Duration o	f breastmilk	$19{\cdot}3\pm9{\cdot}5$	18.9 ± 9.7	$19{\cdot}5\pm9{\cdot}4$	0.386
(months)		(0.5 - 60)	(0.5 - 60)	(1 - 60)	
Birth weight	(g)	3229·4 ±	$3288.4 \pm$	$3204.9 \pm$	0.049^{*}
Dirtin Weight	18/	624.5	629.2	621.4	

 Table 5 Birth and feeding characteristics of children

ECIAF, Extended Composite Index of Anthropometric Failure; SD, standard deviation; Min, minimum; Max. maximum; OR, Odds Ratio; CI, Confidence Interval; ^a stands for Pearson Chi-square test; ^b stands for independent sample *t*-test; ^{*} P < 0.05

_	В	SE	Wald	df	Р	OR	95% CI
Maternal BMI	-0.47	0.021	4.825	1	0.028	0.955	0.916 - 0.995
Maternal WC	-0.008	0.009	0.737	1	0.391	0.992	0.974 - 1.010
Number of children	0.170	0.083	4.181	1	0.041	1.186	1.007 - 1.396
Maternal age at delivery	0.023	0.013	2.940	1	0.086	1.023	0.997 - 1.051
Marital status							
Married	-	-	-	-	-	-	-
Single	-1.239	0.513	5.841	1	0.016	0.290	0.106 - 0.791
Test	χ2	df	Р				
Hosmer and Lemeshow	11.098	8	0.196				
Omnibus likelihood ratio	38.355	5	0.000				

Table 6 Contributions of some maternal factors to the ECIAF status of the children

B, coefficient for the constant; SE, Standard Error; df, degree of freedom; OR, Odds Ratio; CI, Confidence Interval; BMI, Body Mass Index; WC, Waist circumference; * P < 0.05, ** P < 0.01

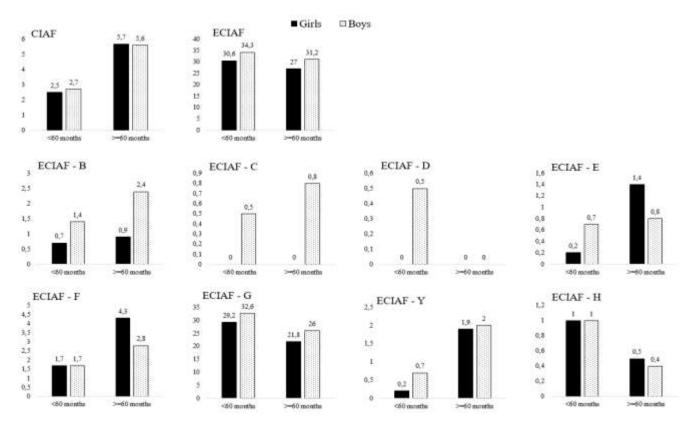


Fig 1 CIAF and ECIAF failure for age groups by sex

CIAF, Composite Index of Anthropometric Failure; ECIAF, Extended Composite Index of Anthropometric Failure; CIAF - A, No failure for CIAF; ECIAF - A, B, Wasting Only; C, Wasting and Underweight; D, Stunting, Wasting and Underweight; E, Stunting and Underweight; F, Stunting only; Y, Underweight only; G, Weight excess (Overweight or obese) only; H, Stunting and weight excess