

A comparison of weighed and recalled intakes for schoolchildren and mothers in rural Kenya†

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Abstract

Objective: To evaluate the relative validity of recalled intake among schoolchildren and mothers in rural Kenya.

Design: Cross-sectional study. Mothers' recall of both the schoolchild's intake and her own intake on the previous day were compared with intakes that were weighed by an interviewer in the home.

Setting: Karurumo location of Embu District in Kenya.

Subjects: A total of forty-two sets of grade 1 students and their mothers.

Results: Between 08.00 and 17.00 hours, when foods were both weighed and recalled, approximately 70% of weighed food items were recalled the next day. Under-reporting of food amounts was seen across most food categories for the recall, with added sugars, sweets and fats being most affected. The recall underestimated energy intakes by approximately 6–9% during this period. Correlation coefficients between nutrient intakes ranged from 0.43 to 0.65, while weighted κ values ranged from 0.22 to 0.50. Higher levels of agreement were noted for nutrient densities, with correlation coefficients between 0.46 and 0.82 and weighted κ values between 0.30 and 0.73.

Conclusions: Although the recall method provides an acceptable alternative to the more labour-intensive and expensive food-weighing method, there is need to further improve its performance in this population through more accurate recall of single foods, especially fruits, as well as added sugars, fats, dairy products and meats, which are often added in small amounts to mixed dishes.

Keywords
Dietary assessment
Schoolchildren
Mothers
Kenya

Because of its relative ease of administration and relatively lower respondent burden, the 24 h recall remains the method of choice for assessing dietary intake among different populations around the world, and still is the most commonly used instrument in dietary assessment in Africa. However, a literature review revealed that few 24 h recall validation studies have been reported for this region. Those that have been carried out have compared nutrient intakes among pre-school children, women, the elderly and households^(1–5). Studies assessing how well the current tools capture schoolchildren's intake are still missing. In the present study, we evaluated the relative validity of recalled intake among schoolchildren and their mothers in rural Kenya by assessing how well mothers can recall the child's

as well as their own dietary intake compared with intake measured by weighing the foods that were consumed.

Methods

The study was conducted in Karurumo location of Embu District in Kenya, where one local elementary school was conveniently selected to participate in the study. All grade 1 students at the school participated in the study. The children were aged 6 to 8 years. Through the students, their households were identified and mothers were requested to participate in the study. Mothers were the main respondents in the study. Informed verbal consent was obtained from the study participants. The study was approved by the UCLA Institutional Review Board and the University of Nairobi Human Subjects Review Board. A total of forty-four schoolchildren and their mothers were selected to participate.

† Preliminary results of these analyses were presented in a paper entitled 'Diet assessment of school children in rural Kenya: a comparison of weighed observations versus a 24-hour recall' at the Federation of American Societies for Experimental Biology (FASEB) conference, Washington, DC, USA, 17–21 April 2004.

Estimation guides such as food models, measuring cylinders (250 ml and 1000 ml) and local household measures were utilized during recall interviews. Food models were two-dimensional line drawings of locally available foods in different sizes, while the household measures consisted of spoon, cup and tin measures. A pre-test of the dietary assessments was conducted in non-study households and adjustments made where necessary. Data collection commenced on 4 May 1998 and lasted one month. For each participating household, dietary assessment was conducted on two consecutive weekdays. Participation rates were high with 95% (forty-two out of forty-four) of households having dietary records for both mother and schoolchild on both days of assessment. Fifty-five per cent of the schoolchildren were male.

Weighed/recalled diet records

Women with previous experience in dietary data collection methods⁽⁶⁾ were recruited as interviewers and re-trained. On the first day of data collection, interviewers arrived at the participating households at 08.00 hours and stayed in the assigned households until 17.00 hours. The interviewers weighed all foods consumed by the schoolchildren and the mothers during that period. They also weighed the ingredients and resulting mixed dish yields produced in the household. All leftover foods were also weighed and recorded. Food servings and ingredients were weighed using the 2 kg Chatillon scale, which measured to the nearest 5 g, while total mixed recipe yields were weighed using the 25 kg Salter hanging scale, which measured to the nearest 50 g. The weighing scales were calibrated daily. Before leaving the households at 17.00 hours the interviewers asked the mothers to prepare ingredients that would be used to make the evening meal. These ingredients were then weighed and recorded by the enumerator.

In addition to weighing foods consumed while present in the home, the interviewers also utilized recall methodologies to collect dietary intake information for the remainder of the 24 h period, following a weighed/recalled diet records (WRDR) protocol that has been previously used in other validation studies in Africa⁽⁷⁾. On arrival at the household, the interviewer used a recall interview to estimate the foods that had been consumed by the identified schoolchild and the mother from the time they woke up in the morning to 08.00 hours. Recall interviews were also utilized to assess any foods that were consumed outside the home by both mothers and schoolchildren. The school day ended at midday for grade 1 students and they returned home for lunch. Schoolchildren were instructed to report any foods that were consumed outside the home through their mothers. They were also asked to describe the primary ingredients in any mixed dishes consumed or brought from outside the home. The same interviewer returned to the household the following day at 08.00 hours to conduct a recall

interview to estimate all the foods that had been prepared and consumed after she had left the household (17.00 hours to bedtime) on the previous day.

24 h recall interviews

A second interviewer visited the household, on the day following the food weighing, to conduct a 24 h recall of all foods consumed the previous day. Although the WRDR and 24 h recall interviewers visited the household on this same day, they chose different times to visit with the WRDR interviewer arriving before the 24 h recall interviewer. Two 24 h recall interviews were conducted with the mother; the first one focused on the schoolchild's intake and the second one focused on her own intake. A standard 24 h recall interview was used. The procedures used have been previously described⁽⁸⁾. Respondents were asked to estimate amounts of foods consumed, mixed dishes prepared in the home and ingredients used in mixed dishes prepared in the home. Common portion measures were used, and later converted to gram weights using a database developed for this purpose. Mothers were probed about food items consumed and asked to describe the primary ingredients for any mixed dishes brought from outside the home. Standard recipes were used to help determine the nutrient intake from these mixed dishes, for both the recalls and the WRDR. The standard recipes were based on dietary information that had been collected in a previous study in the same study area⁽⁶⁾ and were updated to reflect any recipe-content changes that had taken place over the years. Other studies in Africa have reported utilizing standard recipes when actual recipes are not available⁽¹⁾.

Identifying foods that were recalled on the 24 h recall

Food items on the WRDR were compared with the 24 h recall, and were considered matched if the food appeared in both the WRDR and the 24 h recall interview at approximately the same time of day. Those that appeared only on the WRDR and not on the 24 h recall were identified as having been omitted by the 24 h recall. If a food item in the WRDR was similar to another on the 24 h recall and appeared at approximately the same time, we did not count it as a missed food but it was identified as a mis-specified food item. Mis-specified foods had the same main ingredients with slight changes in the other remaining ingredients. Foods that appeared only on the 24 h recall and not on the WRDR were identified as intrusions by the 24 h recall.

Food group and nutrient intake calculations

Food groups of interest included starchy foods, vegetables, fruits, dairy, meats (includes wild and domesticated animals, insects and fish), legumes and/or nuts, high-fat foods/fats/oils, added sugars and sweets, and beverages (tea, cocoa and coffee). Single food items and ingredients in food mixtures were each assigned to the appropriate food group. Nutrient intakes were calculated for each individual using

an international food composition table which was adapted for use in the present study (University of California, Berkeley, CA, USA)⁽⁹⁾. This table contains complete nutrient values for the common foods consumed in rural Kenya. The nutrient contents of less common foods were estimated from similar foods. Nutrients of interest included energy, carbohydrates, total protein, fat, total Fe, total Zn, Ca, vitamins A and C, and riboflavin.

Statistical analysis

Data analysis was carried out using the SAS statistical software program version 8.2 (SAS Institute, Cary, NC, USA). A total of forty-two households with a WRDR and 24 h recall for both the mother and schoolchild were included in the analysis. Ratios between recalled and weighed amounts (in grams) were calculated for the different food groups for those who mentioned the food group in either one of the methods. Differences between nutrient amounts from the two methods were assessed using Wilcoxon's signed-rank test. Spearman's ranked correlations were utilized to assess association between the two methods. Agreement between methods was further assessed by Bland–Altman and limits of agreement (LOA) plots⁽¹⁰⁾. Differences between the two methods were plotted against the average of the two methods to examine the spread of the differences between the two methods and the presence of any systemic bias. The joint classification of nutrients by the two methods was assessed using quartiles of intake for each nutrient for each of the methods. The proportion correctly classified reflects how often an individual's intake fell in the same quartile using both methods. The proportion that was grossly misclassified reflects how often one method assigned an individual's intake into the lowest quartile and the other method assigned it into the highest quartile. Weighted κ values are presented. An α level of 0.05 was used to indicate statistical significance.

Results

Nutrient intake estimates

Energy intake for the full 24 h period was 7883 (SD 3414) kJ (1884 (SD 816) kcal) for the WRDR method and 7301 (SD 3464) kJ (1745 (SD 828) kcal) for the recall method for the schoolchildren. For the mothers, intake was 9037 (SD 4159) kJ (2160 (SD 994) kcal) for the WRDR method compared with 8594 (SD 3410) kJ (2054 (SD 815) kcal) for the recall method. Thus, the recall method underestimated total intake by approximately 7% for the children and by 5% for the mothers. To further investigate this difference, we examined foods consumed at different time periods: before 08.00 hours, between 08.00 and 17.00 hours, and after 17.00 hours. Small differences were seen between the foods consumed before 08.00 hours and those consumed after 17.00 hours, when intakes were recalled by both the WRDR and 24 h recall methods.

These differences may be attributed to slightly different methodologies for the recalls during these time periods. However, for the remaining analyses, we focus on foods consumed between 08.00 and 17.00 hours, when most of the foods were weighed within the WRDR. Thus, the WRDR during this time period should be an accurate standard to which the recalled intakes can be compared.

Approximately 50% of the schoolchildren's and mothers' food items were consumed between 08.00 hours and 17.00 hours and provided approximately 50% of the daily energy intake. The number of food items consumed during this time period ranged from one to seven for the schoolchildren and mothers. Most of the food items (84% and 94% of schoolchildren's and mothers' food items, respectively) consumed during this period were estimated by direct weighing on the first day of data collection. Sixty-eight per cent of the schoolchildren's and 73% of the mothers' intakes were matched in the 24 h recall. Overall, 21% and 15% were omitted from the 24 h recall, and approximately 10% were mis-specified within the 24 h recall. Of the food items on the recalls, 8% and 10% were not present in the weighed records.

Nutrient intake estimates from the 24 h recall for this time period were generally lower compared with those from the weighed records (Table 1). Energy intakes were lower by 9% for the children and by 6% for the mothers. Significant between-method differences in absolute intakes were noted only for vitamins A and C for both the schoolchildren and the mothers. When intakes were expressed as nutrient densities (nutrient per 4184 kJ), vitamin A intake remained significantly different for the schoolchildren while Ca intake became significantly different for the mothers. Correlation coefficients between absolute nutrient intakes for the two methods ranged from 0.49 to 0.64 for the schoolchildren's intakes and from 0.43 to 0.65 for mothers' intakes (Table 2). Correlation coefficients were higher for most of the nutrient densities, with only one value being below 0.50 and a majority of them being above 0.60. All correlation coefficient values for both the absolute nutrients and nutrient densities were statistically significant.

For the schoolchildren, Bland–Altman plots show that the mean differences between 24 h recall and weighed records for the nutrient densities were small, with the individual differences generally clustered around zero (Fig. 1). Most of the nutrient density values fell within the 95% LOA for all nutrients. Between-method differences appeared to increase as mean intake increased for fat and vitamins A and C, suggesting a violation of the assumption of constant variance. Such violations were not noted for the other nutrients. The plots for the mothers' intakes were similar (data not shown).

The joint classification ranged from fair to moderate⁽¹¹⁾, with weighted κ values ranging from 0.30 to 0.50 for the schoolchildren's absolute nutrient intakes and from 0.22

Table 1 Nutrient intake estimated from weighed foods and from a 24 h recall between 08:00 hours and 17:00 hours: schoolchildren (*n* 42) and their mothers (*n* 42), Kenya, May 1998

	Schoolchildren				Mothers			
	Weighed		Recalled		Weighed		Recalled	
	Mean	SD	Mean	SD	Mean	SD	Mean	SD
Absolute nutrient intakes								
Energy (kJ)	4042	1895	3682	2218	4606	2238	4322	1904
Energy (kcal)	966	453	880	530	1101	535	1033	455
Carbohydrate (g)	188	87	171	108	217	110	205	91
Fat (g)	16	11	14	9	17	12	15	9
Protein (g)	28	18	26	19	31	18	31	16
Ca (mg)	234	223	172	108	275	248	189	138
Fe (mg)	11.0	7.5	10.0	7.4	12.2	8.2	11.8	6
Zn (mg)	4.8	2.9	4.6	3.2	5.3	3.0	5.4	2.8
Riboflavin (mg)	0.66	0.38	0.59	0.32	0.73	0.40	0.67	0.30
Vitamin A (RAE)	320	436	174*	148	307	341	188*	191
Vitamin C (mg)	103	119	73*	98	67	62	45*	45
Nutrient intakes per 4184 kJ								
Carbohydrate (g)	197	21	195	23	197	22	199	19
Fat (g)	16	9	17	11	16	9	14	8
Protein (g)	29	7	29	7	29	8	29	7
Ca (mg)	231	126	205	115	246	160	190*	134
Fe (mg)	11	3	11	3	11	3	11	3
Zn (mg)	4.8	1	5	1	5	1	5	1
Riboflavin (mg)	0.70	0.10	0.70	0.17	0.70	0.18	0.66	0.19
Vitamin A (RAE)	319	362	214*	178	272	246	206	203
Vitamin C (mg)	131	220	103	145	77	120	52	51

RAE, retinol activity equivalent.

Mean values from the recall were significantly different from those weighed using Wilcoxon's signed-rank test: **P* < 0.05.**Table 2** Spearman correlation, nutrient classification and weighted κ values between weighed and recalled nutrient intakes between 08:00 hours and 17:00 hours: schoolchildren (*n* 42) and their mothers (*n* 42), Kenya, May 1998

	Schoolchildren				Mothers			
	ρ	Classification			ρ	Classification		
		%Cc	Wk	95% CI of Wk		%Cc	Wk	95% CI of Wk
Absolute nutrient intakes								
Energy (kJ)	0.51	29	0.30	0.12, 0.48	0.44	36	0.26	0.06, 0.47
Carbohydrate (g)	0.49	40	0.34	0.14, 0.53	0.45	31	0.22	0.02, 0.43
Fat (g)	0.56	45	0.42	0.22, 0.61	0.65	45	0.46	0.26, 0.65
Protein (g)	0.63	48	0.50	0.31, 0.68	0.56	33	0.34	0.15, 0.53
Ca (mg)	0.60	38	0.42	0.23, 0.60	0.50	43	0.38	0.17, 0.59
Fe (mg)	0.64	45	0.46	0.26, 0.65	0.56	36	0.26	0.06, 0.47
Zn (mg)	0.62	48	0.50	0.31, 0.68	0.55	38	0.30	0.09, 0.51
Riboflavin (mg)	0.56	31	0.34	0.17, 0.51	0.48	36	0.34	0.15, 0.53
Vitamin A (RAE)	0.62	48	0.46	0.27, 0.65	0.49	41	0.38	0.17, 0.58
Vitamin C (mg)	0.50	38	0.30	0.08, 0.52	0.43	48	0.34	0.10, 0.58
Nutrient intakes per 4184 kJ								
Carbohydrate (g)	0.68	55	0.53	0.35, 0.72	0.75	52	0.50	0.30, 0.69
Fat (g)	0.72	60	0.57	0.38, 0.76	0.82	67	0.69	0.53, 0.84
Protein (g)	0.74	50	0.53	0.36, 0.71	0.85	71	0.73	0.58, 0.87
Ca (mg)	0.51	48	0.38	0.16, 0.60	0.46	41	0.34	0.13, 0.55
Fe (mg)	0.75	52	0.50	0.30, 0.69	0.73	64	0.61	0.42, 0.80
Zn (mg)	0.76	48	0.53	0.37, 0.70	0.82	60	0.61	0.45, 0.77
Riboflavin (mg)	0.53	57	0.50	0.29, 0.70	0.60	65	0.42	0.22, 0.61
Vitamin A (RAE)	0.58	50	0.46	0.25, 0.66	0.57	50	0.46	0.25, 0.66
Vitamin C (mg)	0.62	48	0.46	0.26, 0.65	0.66	52	0.50	0.30, 0.69

 ρ , Spearman's correlation; %Cc, percentage correctly classified; Wk, weighted κ ; RAE, retinol activity equivalent.All correlation coefficients and weighted κ values were statistically significant: *P* < 0.05.

to 0.46 for the mothers' intakes (Table 2). The proportion of absolute nutrient intakes that were classified within the same quartile ranged from 29% to 48% for schoolchildren and from 31% to 48% for mothers. Gross misclassification

was very low, ranging from 0 to 5%. Joint classification and weighted κ values were considerably higher for most of the nutrient densities, showing fair to moderate agreement for schoolchildren and fair to good agreement for mothers⁽¹¹⁾.

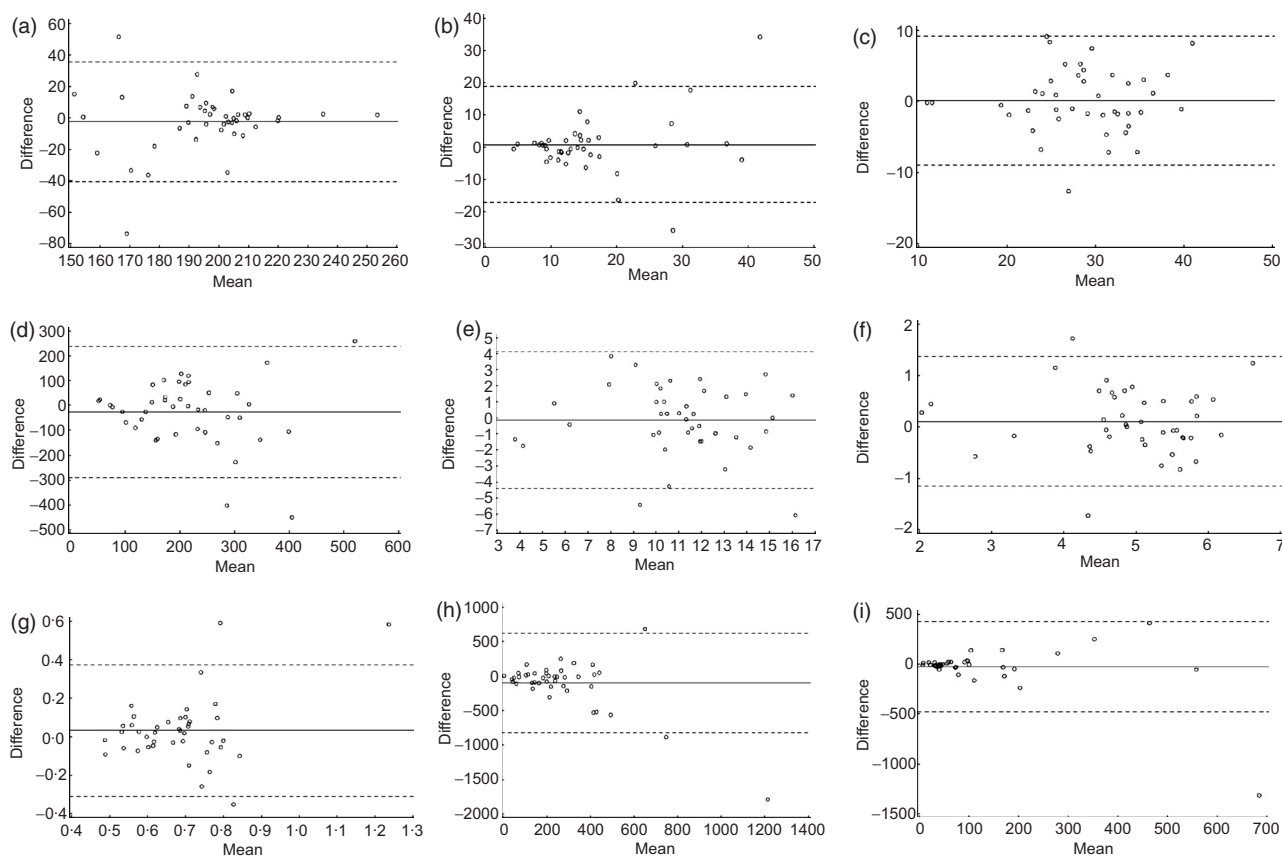


Fig. 1 Bland–Altman plots of the difference between recalled and weighed nutrient intakes, expressed per 4184 kJ, for Kenyan schoolchildren (n 42): (a) carbohydrate (g); (b) fat (g); (c) protein (g); (d) calcium (mg); (e) iron (mg); (f) zinc (mg); (g) riboflavin (mg); (h) vitamin A (μ g retinol activity equivalent); and (i) vitamin C (mg). - - -, upper and lower limits of agreement; —, mean difference

Under-reporting of food amounts was seen across most food categories for the recall (Table 3). Added sugars and sweets and fats were particularly low relative to the weighed intakes, for both schoolchildren and mothers. Reported levels of dairy and beverages were also notably lower than the weighed amounts in the mother’s diets, while meats were lower in the diets of the four children who reported meat intake. Beverage intake was over-reported, on average, for the children.

Omissions and intrusions on the recall

Fifty per cent (21 out of 42) of the schoolchildren and 31% (13 out of 42) of the mothers had at least one weighed food item that was omitted from the recalled intake. A majority of the omitted foods were single food items, especially fruits.

Foods that appeared only on the 24 h recall and not on the WRDR were identified as intrusions on the 24 h recall. Twenty-one per cent (9 out of 42) of the schoolchildren and 19% (8 out of 42) of the mothers had at least one food item identified as an intrusion on the 24 h recall. A majority of the intrusions were fruits, and most of these fruits appeared to replace those that had been omitted. Thus, these fruit items might be considered possible

Table 3 Ratio (expressed as a percentage) of recalled to weighed mean intakes (g/d) between 08:00 hours and 17:00 hours for the main food groups consumed (only for subjects who reported the food group for at least one of the methods): schoolchildren (n 42) and their mothers (n 42), Kenya, May 1998

Food type	Schoolchildren		Mothers	
	<i>n</i>	Ratio (%)	<i>n</i>	Ratio (%)
Starchy foods	42	81	41	93
Vegetables	40	73	38	73
Fruits	24	100	19	75
Dairy	14	99	22	64
Meats	4	63	4	84
Legumes/nuts	35	100	35	113
Fats	31	61	29	55
Added sugars & sweets	23	43	27	31
Beverages	12	127	19	70

indications of mis-specifications instead of being complete omissions. As shown in Table 3, 100% of the amount of the fruit food group on the weighed intakes was present on the recalled intakes for the children, although the specific type fruit may have been different. The percentage of the fruit food group that was recalled was lower for the mothers’ recalls (75%), probably indicating that more true omissions occurred.

Discussion

In the current study, most of the weighed food items (approximately 70%) were reported on the 24 h recall. Approximately 10% of the food items were mis-specified, and 15–20% were omitted on the recall. There were very low levels of food intrusions. Most of the foods that were omitted or added were single food item snacks, with fruits making up the highest proportion of these foods. Further examination of the foods added to the recalls indicated possible mis-specification of these fruits.

The results showed lower mean intake levels for both foods and nutrients from the recall method than from the weighed method, with differences being more evident with the schoolchildren's intakes. Under-reporting on dietary recalls has been reported in many countries^(12–14), although the magnitude of the bias is often greater than what was seen in this study in Kenya. Underestimation was seen across most of the food categories (Table 3). The energy-related errors may have resulted from underestimation of fats and added sugars. Errors in estimating energy intake, arising mostly from estimation of snack foods and main staples, have been previously noted among Ghanian children and rural Malawian women, respectively^(1,15).

The correlation coefficients were relatively high for most of the nutrients of interest in our study (at least 0.47 for the schoolchildren and at least 0.43 for the mothers), indicating that the intakes were ranked similarly by the two methods. Weighted κ values and proportions of nutrients classified into the same categories were fair to moderate, while gross misclassification was low for all nutrients, illustrating that the recall ranks schoolchildren's and mothers' intakes moderately well.

The nutrient densities showed smaller mean differences and higher correlation coefficients, as well as higher weighted κ values and percentages of nutrients classified into the same quartile, indicating that adjustment for differences in energy intake removes much of the difference and increases the level of agreement in nutrient intake. This also suggests that between-method differences may have resulted more from differences in the quantity of foods, rather than differences in the type or mix of foods consumed. Heteroscedasticity was noted for fat and vitamins A and C, which may lead to an underestimation of the standard errors but should not bias the correlation coefficients⁽¹⁶⁾.

Comparison with other validation studies

Dop *et al.* concluded that two 24 h recalls were as precise as a single day's weighed records among weanlings in Senegal⁽⁵⁾. Kigutha found no significant differences in mean nutrient intakes when comparing repeated 24 h recalls and weighed food records among pre-school children and elderly people in rural Kenya. However, she also noted that the 24 h recall had a tendency to overestimate Fe, Ca and vitamin A and C intakes among pre-school children and to underestimate energy, protein, fat,

thiamin and niacin among the elderly⁽²⁾. However, the study did not go beyond the use of correlations to compare the relative ranking of intakes with the two methods. Ferguson *et al.* noted that the 24 h recall performed poorly in estimating portion sizes of main staples, soups and fruits among Ghanian pre-school children and main staples and relishes among women in rural Malawi^(1,15). The authors concluded that the 24 h recall performed better in categorizing individuals according to their dietary patterns rather than in estimating absolute daily intakes^(1,15).

There are several strengths of the present analysis, including the availability of weighed intakes during the period of 08.00 to 17.00 hours. Also, our extensive food composition table allowed us to compare intakes of a variety of nutrients and food groups. Although our sample size is relatively small (n 42 for both schoolchildren and mothers), there was sufficient power to detect intake differences as small as 10% for most nutrients. Although a multi-day study, with a full 24 h period of weighed intakes, would have been desirable⁽¹⁷⁾, funds were not available to extend the intensive weighed intake method.

The presence of an enumerator in the household for the WRDR may have influenced household dietary behaviour. However, the schoolchildren's mean daily energy intake in the present study was close to that reported using 24 h recalls in a larger study of other children in the same area⁽¹⁸⁾. No recent data are available for adult females in this area of Kenya, but the mothers' mean energy intake levels were higher than those reported in the study population in 1984–5⁽⁶⁾. However, this increase may reflect dietary changes that have taken place over the years, because similar increases have been noted in the schoolchildren's nutrient intake⁽¹⁸⁾.

A large nutrition study was conducted in this district of Kenya almost 15 years ago, but due to the long time lag, it is unlikely that the young mothers who participated in the current study would have an increased awareness of their food intake as a result. Thus, these findings should be applicable to similar populations in other communities.

Strengths and weaknesses of the two methods

The weighed method should provide an accurate estimate of food intake during the time that the enumerator is in the home. However, this method has several limitations for large nutrition studies. It is a relatively expensive method because it requires the presence of an enumerator for all or part of the period being studied. It is seldom feasible for the enumerator to remain in the household overnight, so foods prepared and consumed between 17:00 hours and 08:00 hours the next morning must be estimated via a recall method. If it is necessary to collect multiple days of dietary data, as is usually the case, then the cost of data collection can be prohibitive. Finally, the presence of the enumerator in the home may alter food habits, so that the resulting intake data do not reflect usual intakes. In some cultures, the presence of a stranger in the home may be considered unacceptable.

The recall method of collecting intakes is often chosen to reduce the expense of collecting intake data. However, as we found in our study, intakes are likely to be underestimated due to errors of omission, mis-specification of foods and incorrect estimates of amounts consumed. Furthermore, accurate intake estimates require a convenient method of converting common portion sizes into gram weights. For our study, we developed a database of almost 2500 entries that allowed us to automatically convert common measures, such as a piece of fruit or a cup of milk, into the corresponding gram weight. However, once developed, such a database can be used for further studies in the same area or, with some modifications, for studies in similar populations elsewhere.

Neither method can accurately capture foods consumed outside the home, which may be particularly important for schoolchildren⁽¹⁹⁾. Because the interviewer is present in the home during the day for the weighing method, it is possible to ask both adults and children about foods consumed elsewhere at the time they return from school or other activities outside the home. However, we found similar reporting of mixed dishes consumed outside the home with the two methods: 5% and 3% of the mixed dishes for the mothers' and schoolchildren's intakes from the weighed records, and 4% for both groups from the recalls. As children grow older and spend more time outside the home, the number of foods consumed outside the home is expected to rise, thus making it more difficult for mothers to recall the children's intakes. Thus, the study's results should be generalized primarily to mothers' abilities to recall their own intake and that of their young children who still rely heavily on them for their daily intake. As children grow older, they should be encouraged to participate more prominently in reporting their own dietary intake.

In conclusion, in our study the recall method underestimated energy intake by about 6% for the mothers and 9% for the children compared with weighed records between 8.00 and 17.00 hours. Furthermore, recalls provided a relatively good method of ranking and classifying nutrient intake. The recall method performed better when comparing nutrient densities, suggesting that estimates of absolute nutrient values and of energy intake should be interpreted with more caution. To further improve 24 h recall performance in this population, there is a need to improve the recall of single foods, especially fruits, and that of added sugars, fats, dairy products and meats, which are often added in small amounts to mixed dishes. Because weighing all foods consumed is not feasible in most large studies, the 24 h recall method provides an acceptable alternative.

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